$f_{X}(x) = c_{k} \cdot \sin^{k}(x), x \in (0,\pi), k = 1,2,3...\# \text{ columns} - 1, \text{ and } c_{k} = \frac{\Gamma(k/2+1)}{\sqrt{\pi}\Gamma(k/2+1/2)}$ $F_{X}(x;k) \sim \frac{1}{2} - \left(\frac{1}{2}\right) \cdot F_{Beta}\left[\cos^{2}(x); \frac{1}{2}, \frac{1+k}{2}\right] \text{ for } x < \frac{\pi}{2}, \qquad \sim \frac{1}{2} + \left(\frac{1}{2}\right) \cdot F_{Beta}\left[\cos^{2}(x); \frac{1}{2}, \frac{1+k}{2}\right] \text{ for } x \ge \frac{\pi}{2}$

Beating the Correlation Breakdown, for Pearson's and Beyond:

Robust Inference and Flexible Scenarios and Stress Testing for Financial Portfolios

JD Opdyke

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$$= \pi - \arccos\left(\sqrt{F_{Beta}^{-1}\left(1 - 2\left[1 - p\right]; \frac{1}{2}, \frac{1 + k}{2}\right)}\right) \text{ for } p \ge 0.5$$

 $F^{-1}(p;k) = \arccos\left(\sqrt{F_{Beta}^{-1}\left(1-2p;\frac{1}{2},\frac{1+k}{2}\right)}\right)$ for p < 0.5;



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JD Opdyke, the sole author of the work contained herein, presents this work in the public domain with the permission of, and gratitude to, Mark Prindiville, PhD, CRO – Allstate. The author also would like to thank Dan Kirsner, PhD, and Ming Dai, PhD for reviewing early drafts of this presentation. Any errors are my own.

This is a presentation format of a chapter in my forthcoming book, "Beating the Correlation Breakdown, for Pearson's and Beyond: Robust Inference and Flexible Scenarios and Stress Testing for Financial Portfolios," Elements in Quantitative Finance series, Cambridge University Press, eds. Ricardo Rebonato, PhD.



Abstract

• We live in a multivariate world, and effective modeling of financial portfolios, including their construction, allocation, forecasting, and risk analysis, simply is not possible without explicitly modeling the dependence structure of their assets. Correlation/concordance matrices rightly play a central, ubiquitous, and foundational role here. They often stand as the most impactful parameter in portfolio models, and yet both the literature and practitioners typically fail to treat them with the same level of quantitative and probabilistic rigor as the other estimated parameters in these models. This is especially troubling given the widely documented 'correlation breakdowns' that occur during times of extreme market stress, which is when risk analytics, and the consequences of the methods chosen for portfolio construction and allocation, matter the most.

• This work builds on prior research covering geometric frameworks to derive the finite-sample distributions of a very broad class of the most widely used correlation matrices and dependence measures – including Pearson's product moment, Spearman's Rho, Kendall's Tau, Szekely's (generalized) distance correlation, the Tail Dependence Matrix (see Embrechts et al, 2016, and Shyamalkumar & Tao, 2020), and more. These distributions maintain validity under the most general conditions possible, requiring only the positive definiteness of the matrix. The proposed Nonparametric Angles-based Correlation method (NAbC) unifies estimation of the confidence intervals and p-values associated with each and every pairwise correlation cell, with those associated with the entire matrix, determining both simultaneously, consistently, and quantitatively. It also provides the quantile function for both the individual cells and the entire matrix: when given a matrix of cumulative distribution function values, it provides the unique, corresponding correlation matrix.

• All results obtained under challenging, real-world data conditions (e.g. varying degrees of tail heaviness, asymmetry, non-stationarity, and serial correlation in the margins, as well as under complex copula functions and near-singular matrices) are consistent with those well established in the Random Matrix Theory literature. But they are more robust and accurate under many conditions than more complex and limited spectral methods. Finally, NAbC provides something no other approach does: flexible scenario definition and stress testing, allowing for selective perturbation of chosen cells in the matrix, while holding the remaining non-chosen cells constant. NAbC's range of application is as broad as the widespread and necessary use of these correlation/dependence matrices themselves, making it a potent tool for proactively flagging, probabilistically monitoring, and potentially mitigating and avoiding the worst consequences of correlation breakdowns.





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XII. APPENDICES 1, 2, 3: Empirical Results for NAbC–Full Matrix, Targeted Scenarios,

and Kendall's and Spearman's Matrices

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- Responsible use of any (portfolio) model that incorporates any type of correlation matrix* requires knowledge of its sampling distribution. To date, an easily implemented, robust, and flexible solution valid under general conditions has remained elusive, even as this is often the most impactful parameter in such models. Knowledge of its finite sample density is especially important for portfolio models used in stress testing, or requiring the specification of particular scenarios with particular correlation values (e.g. views-based portfolio analyses a la Black-Litterman (1991) and its variants).
- <u>Objective</u>: Develop a method that provides the finite sample density of the correlation matrix, and its inverse ('quantile function'), <u>under the most general conditions possible</u>, requiring only positive definiteness. Key characteristics:
 - Flexible: make this density

i. a direct function of the densities associated with each of the correlation matrix cells, and ii. valid and applicable to **subsets of cells in the matrix** while holding the rest constant

Satisfying i. and ii. allows for flexible stress testing and scenario specification, with the added benefit that specific scenarios for the correlation matrix now can be defined probabilistically.

- **<u>Robust inference</u>**, even when the matrix approaches singularity (i.e. non-positive definiteness).
- <u>Accurate</u>: does not rely on approximations that can be inaccurate under conditions that are not uncommon in financial portfolios (e.g. Fisher Z transformation under near singularity).
- <u>Scalable</u>: reasonably fast for reasonably high dimensions (e.g. 100x100), with an implementation that does not change or become unwieldy in higher dimensions.
- Requirements: The above is conditional on 1. a specified or well estimated correlation matrix (this is a very rich literature) and 2. a specified or well estimated data generating mechanism.
 - * Initial focus is on Pearson's product moment correlation (Pearson, 1895); Spearman's Rho (Spearman, 1904) & Kendall's Tau (Kendall, 1938) are treated explicitly later herein.



Here I restate the **Objective** in the form of <u>research questions that</u>, taken together, only this method (Nonparametric Angles-based Correlations, "<u>NAbC</u>") <u>answers under general conditions</u>:

- Given a specified or well-estimated correlation matrix [A], and its specified or well-estimated data generating mechanism...
- General:
 - 1. Confidence Intervals: What are the two correlation matrices that correspond to the lower- and upper-bounds of the 95% (or any) confidence interval for [A]? What are, simultaneously, the individual confidence intervals for each and every cell of [A]?
 - **2.** Quantile Function: What is the unique correlation matrix associated with [B], a matrix of cumulative distribution function values associated with the corresponding cells of [A]?
 - **3. p-values**: Under the null hypothesis that observed correlation matrix [C] was sampled from the data generating mechanism of [A], what is the p-value associated with [C]? And simultaneously, what are the individual p-values associated with each and every cell of [C]?

[A]				
1				
0.2	1			
-0.1	0.3	1		
0.3	-0.3	-0.1	1	
0.6	0.4	0.0	0.1	1

_					
	0.8				
	0.7	0.8			
	0.8	0.7	0.7		
	0.7	0.8	0.8	0.7	

[D]

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[C]				
1				
0.40	1			
0.20	0.10	1		
0.03	-0.07	-0.20	1	
0.33	0.60	0.25	-0.23	1



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Under a specific scenario only selected pairwise correlation cells of [A] will vary, while the rest are held constant, unaffected by the scenario (e.g. COVID). This is matrix [D].

Scenario Specific:

- 4. Confidence Intervals: What are the two correlation matrices that correspond to the lower- and upper-bounds of the 95% (or any) confidence interval for [D] (holding constant the non-selected red cells)? What are, simultaneously, the individual confidence intervals for only those cells of [D] that are relevant to the scenario (green)?
- 5. Quantile Function: What is the unique correlation matrix associated with [E], a matrix of cumulative distribution function values associated with the corresponding cells of [D]?
- 6. p-values: Under the null hypothesis that observed correlation matrix [F] was sampled from the data generating mechanism of [D], what is the p-value associated with [F] (with red cells held constant)? And simultaneously, what are the individual p-values associated with every (non-constant, green) cell of [F]?





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Only NAbC can simultaneously answer 1.-6. above under general conditions – that is, with no restrictions other than those required for the existence of a correlation matrix (i.e. positive definiteness / second-order marginals*). NAbC does this robustly, even for near-singular matrices, and at scale, relying only on straightforward, long-established, verifiable geometric structures. Answers Below**:

Q1	Q2	Q3	Q4	Q5	Q6
	1	p-value=0.1473		1	p-value=0.0526
	0.631 0.488 0.116 0.183 1	0.0033		0.5988 0.4304 0.0521 0.1312 1	0.04032
1	1	0.0227 0.0297 0.0079	1	1	0.00008 0.00184 0.01088
-0.017 1	0.406 1	0.0401 0.0021 0.0101 0.0049	0.1996 1	0.1996 1	
-0.316 0.117 1	0.130 0.517 1		-0.0995 0.0827 1	-0.0995 0.5166 1	
0.089 -0.558 -0.214 1	0.486 0.056 0.190 1		0.2996 -0.2991 -0.0998 1	0.2996 -0.2991 -0.0998 1	
0.439 0.126 -0.345 -0.136 1	0.727 0.631 0.368 0.443 1		0.5988 0.3110 -0.1545 -0.0654 1	0.5988 0.4633 0.1605 0.2680 1	
1	1		1	1	
0.049 1	0.347 1		0.1996 1	0.1996 1	
-0.250 0.165 1	0.060 0.452 1		-0.0995 0.1432 1	-0.0995 0.4537 1	
0.154 -0.497 -0.203 1	0.435 -0.056 0.091 1		0.2996 -0.2991 -0.0998 1	0.2996 -0.2991 -0.0998 1	
0.491 0.212 -0.253 -0.089 1	0.693 0.569 0.265 0.341 1]	0.5988 0.3404 -0.1122 -0.0169 1	0.5988 0.4492 0.1158 0.2181 1	

* Note this is a Hilbert (and a Banach) space. ** Solely for ease of replication, the data generating mechanism for these examples is simply multivariate standard normal, with the given correlations. N sims=25k, n obs=160.



II. Definitions

• Pearson's product moment correlation is defined as ρ ; its sample analog *r* uses sample moments:

$$\rho = \frac{\sum_{i=1}^{N} (X_i - E(X)) (Y_i - E(Y))}{\sqrt{\sum_{i=1}^{N} (X_i - E(X))^2} \sqrt{\sum_{i=1}^{N} (X_i - E(X))^2}} = \frac{Cov(X,Y)}{\sigma_X \sigma_Y} \qquad r = \frac{\sum_{i=1}^{n} (X_i - \frac{1}{n} \sum_{i=1}^{n} X) (Y_i - \frac{1}{n} \sum_{i=1}^{n} Y)}{\sqrt{\sum_{i=1}^{n} (X_i - \frac{1}{n} \sum_{i=1}^{n} X)^2} \sqrt{\sum_{i=1}^{n} (X_i - \frac{1}{n} \sum_{i=1}^{n} Y)^2}} = \frac{Cov(X,Y)}{s_X s_Y}$$

- The corresponding correlation matrix R is the matrix of all pairwise correlations, with the following characteristics:
 - $\begin{array}{c|c} \text{Symmetry:} & r_{i,j} = r_{j,i} \\ \text{Unit diagonal entries:} & r_{i=j} = 1 \\ \text{*Bounded non-diagonal entries:} & -1 \leq r_{i,j} \leq 1 \\ \text{The matrix is positive definite, i.e. all eigenvalues } \lambda_i > 0 \end{array} \qquad R = \begin{bmatrix} 1 & r_{1,2} & r_{1,3} & r_{1,4} \\ r_{2,1} & 1 & r_{2,3} & r_{2,4} \\ r_{3,1} & r_{3,2} & 1 & r_{3,4} \\ r_{4,1} & r_{4,2} & r_{4,3} & 1 \end{bmatrix}$
 - For completeness, we define eigenvalues below: If there exists a nonzero vector *v* such that *Rv* = λ*v* then λ is an eigenvalue of *R* and *v* is its corresponding eigenvector. λ and *v* can be obtained by solving det(λ*I* - *R*) = 0, then det(λ*I* - *R*)*v* = 0, where *I* is the identity matrix and det is the determinant The eigenvalue can be thought of as the magnitude of the (portfolio) variance in the direction of the eigenvector.

* Of course, these can be tighter under specific circumstances, such as for equicorrelation matrices where $-1/(n-1) \le \rho \le 1$, $n = \dim(\rho)$.

• For the identity matrix (correlations all equal zero) under Gaussian data and d≥2, Gupta & Nagar (2000) provide the pdf

$$f(R) = \frac{\left[\Gamma\left(\left[n-1\right]/2\right)\right]^2 |R|^{(n-p-2)/2}}{\Gamma_p\left(\left[n-1\right]/2\right)}, \quad \text{where } || \text{ is the determinant function,} \\ -1 \le r_{i,j} = r_{j,i} \le 1, \ r_{i,i} = 1, \ 1 \le i, j \le p \text{ and } \Gamma_p\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]} \prod_{i=1}^p \Gamma\left(\left[n-i\right]/2\right)$$

where *R* is the empirical correlation matrix, *p* is its dimension (#rows/cols), *n* is the sample size, and Γ is the gamma function. However, we focus below on methods for efficiently sampling this density.

- The c-vines and onion methods (Lewandowski et al., 2009)
- The restricted Wishart distribution approach of Wang et al. (2018).
- The direct formulation method of Madar (2015)
- The Cholesky-Metropolis method of Cordoba et al. (2018)
- The polar angles distribution of Makalic and Schmidt (2018) combined with the polar representation of Pinheiro & Bates (1996), Rebonato & Jaeckel (2000), & Rapisarda et al. (2007).
- The geometric interpretation provided by the polar angles approach is based on COSINE SIMILARITY: the cosine of the angle between two mean-centered vectors X, Y is equal to Pearson's correlation:

$$\cos(\theta) = \frac{\text{inner product}}{\text{product of norms}} = \frac{\langle \mathbf{X}, \mathbf{Y} \rangle}{\|\mathbf{X}\| \|\mathbf{Y}\|} = \frac{\sum_{i=1}^{N} (X_i - E(X))(Y_i - E(Y))}{\sqrt{\sum_{i=1}^{N} (X_i - E(X))^2} \sqrt{\sum_{i=1}^{N} (X_i - E(X))^2}} = \frac{Cov(X, Y)}{\sigma_X \sigma_Y} = \rho, \text{ with } 0 \le \theta \le \pi$$

• The framework of our proposed method relies on this geometric interpretation, so it is worth examining one sampling implementation based on it here.



- Of the sampling methods listed above, Makalic and Schmidt (2018) arguably is the most simple and among the fastest (see Cordoba et al., 2019).
- The polar angles approach to sampling the (Gaussian data) identity matrix is an efficient recognition
 of the fact that merely perturbing non-diagonal correlation values uniformly, between -1 and 1, will
 generate mostly non-positive definite matrices. In fact, Bohn & Hornik (2014) and Pourahmadi &
 Wang (2015) show that the ratio of valid correlation matrices to all matrices generated this way that
 LOOK like correlation matrices is

$$\Pr\left(rand "R" \sim PosDef\right) = X = \frac{\prod_{j=1}^{p-1} \left[\sqrt{\pi} \Gamma\left(\frac{j+1}{2}\right)\right]^{j}}{2^{p(p-1)/2}} < \prod_{j=1}^{p-1} \left[\frac{\sqrt{\pi}}{2}\right]^{j} = \left[\frac{\sqrt{\pi}}{2}\right]^{p(p-1)/2}; \lim_{p \to \infty} \left[X\right] = 0$$

• For even relatively small matrices of dimension p=25, the odds of successfully randomly generating a single valid correlation matrix are less than 2 in ten quadrillion: hence, for the sake of computational efficiency, we need to ONLY generate valid, positive definite matrices by constraining our sampling to those matrices on the hyper hemisphere, as described below.



•

- The Cholesky factorization of a correlation matrix (see below) has rows whose squares sum to 1.0, so
 it is commonly used as a convenient way to ensure that samples remain on the unit hypersphere (or
 technically in this case, the unit hyper-hemisphere of dimension p).
- Pourahmadhi & Wang (2015) and others show that the uniform distribution of positive definite matrices on the p-dimensional hemisphere is proportional to the determinant of the Jacobian, which is defined in terms of the Cholesky factorization as shown below (see also Cordoba et al., 2018)

det $[J(U)] = 2^p \prod_{i=1}^{p-1} u_{ii}^i$ where U is the Cholesky factorization of correlation matrix $R = UU^t$

So Makalic and Schmidt (2018) and others (see Pourahmadi & Wang, 2015) recognized that sampling polar angles based on pdf

$$f_X(x) = c_k \cdot \sin^k(x), x \in (0,\pi), k = 1, 2, 3, ..., (\# \text{columns} - 1), \text{ and } c_k = \frac{\Gamma(k/2 + 1)}{\sqrt{\pi}\Gamma(k/2 + 1/2)}$$

satisfies this constraint. Although not mentioned in Makalic and Schmidt (2018), importantly note that k = #columns - column# (so for the first column of a p=10x10 matrix, k=9; for the second column, k=8, etc.). So the spread of the angles distributions is a function of the column number (as shown in graphical results in the Appendices below).



- For completeness, we include below the definition of the Cholesky factor and corresponding formulae:
- A correlation matrix *R* will be real, symmetric positive-definite, so the unique matrix *B* that satisfies

 $R = BB^T$

where *B* is a lower triangular matrix (with real and positive diagonal entries), and B^{T} is its transpose, is the Cholesky factorization of *R*. Formulaically, *B*'s entries are as follows:

$$B_{j,j} = (\pm) \sqrt{R_{j,j} - \sum_{k=1}^{j-1} B_{j,k}^2} \qquad B_{i,j} = \frac{1}{B_{j,j}} \left(R_{i,j} - \sum_{k=1}^{j-1} B_{i,k} B_{j,k} \right) \text{ for } i > j$$

The Cholesky factorization can be thought of as the matrix analog to the square root of a scalar.



• For sampling from parametric distributions, ideally the cdf can be inverted analytically, and then inverse probability sampling can be used. But in this case, Maklic and Schmidt (2018) state:

"Generating random numbers from this distribution is not straightforward as the corresponding cumulative density [sic] function, although available in closed form, is defined recursively and requires O(k) operations to evaluate. The nature of the cumulative density [sic] function makes any procedure based on inverse transform sampling computationally inefficient, especially for large k."

- However, we shall see below that, in fact, Opdyke (2020, 2022, 2023) derives the analytical cdf of this distribution, as well as its analytical inverse, thus enabling the use of the inverse probability transform for more efficient sampling here, especially for large k (i.e. for large dimensional matrices).
- Maklic & Schmidt's (2018) approach was rejection sampling with a scaled beta distribution envelop:

1. Generate X ~ Beta
$$(k+1,k+1)$$
 ~ $\frac{x^k(\pi-x)^k}{B(k+1,k+1)\pi^{(2k+1)}}$, $x \in (0,\pi)$, $k \ge 1$, $B(q,r) = \frac{\Gamma(q)\Gamma(r)}{\Gamma(q+r)}$ = Euler's beta function

2. Generate U ~ Uniform (0,1)

3. Accept X if
$$\frac{\ln(U)}{k} \le \ln\left(\frac{\pi^2 \sin(X)}{4X(\pi - X)}\right)$$
 4. Otherwise go to 1.

• The algorithm's maximum expected iterations per sample is $\pi/(2\sqrt{2})=1.11$, making its theoretical time complexity more than 10% less efficient than an analytical solution. In practice, the speed of customized analytical functions make it over 30% less efficient (see Roman (2023) below).



- The more general and important point here is that we have a direct relationship between the angles between two data vectors, and their corresponding correlations. Translation between the two is straightforward and non-directional. This will be explored in depth later, but in summary:
 - 1. estimate the correlation matrix
 - 2. obtain the Cholesky factorization of the correlation matrix
 - 3. Use inverse trigonometric functions on 2. to obtain corresponding spherical angles

And in reverse:

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- 3. Start with a matrix of spherical angles
- 2. apply trigonometric functions to obtain the Cholesky factorization
- 1. multiply 2. by its transpose to obtain the corresponding correlation matrix

see Rebonato & Jaeckel (2000) and Rapisarda et al. (2007) (note a typo in the formula in Pourahmadi & Wang (2015) for the first 3 steps)

- Note that this relationship is true generally, not only for the identity matrix under Gaussian data.
- Note also the inverse relationship between angles and correlations: correlations decrease monotonically in their corresponding angles, i.e. correlations increase as angles decrease to zero, and decrease as angles increase to π (see Zhang et al. (2015) and Lu et al. (2019)). The range from 0 to π rather than 0 to 2π is why this is the p-dimensional hyper hemisphere rather than the hypersphere.



So for *R*, a *p* x *p* correlation matrix:

$$R = \begin{bmatrix} 1 & r_{1,2} & r_{1,3} & \cdots & r_{1,p} \\ r_{2,1} & 1 & r_{2,3} & \cdots & r_{2,p} \\ r_{3,1} & r_{3,2} & 1 & \cdots & r_{3,p} \\ r_{4,1} & r_{4,2} & r_{4,3} & \cdots & r_{4,p} \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ r_{p,1} & r_{p,2} & r_{p,3} & \cdots & 1 \end{bmatrix} \quad R = BB' \text{ where } B \text{ is the Cholesky factor of } R \text{ and}$$

$$B = \begin{bmatrix} 1 & 0 & 0 & \cdots & 0 \\ \cos(\theta_{2,1}) & \sin(\theta_{2,1}) & 0 & \cdots & 0 \\ \cos(\theta_{3,2})\sin(\theta_{3,1}) & \sin(\theta_{3,2})\sin(\theta_{3,1}) & \cdots & 0 \\ \cos(\theta_{4,1}) & \cos(\theta_{4,2})\sin(\theta_{4,1}) & \cos(\theta_{4,2})\sin(\theta_{4,1}) & \cdots & 0 \\ \vdots & \vdots & \vdots & \cdots & \vdots \\ \cos(\theta_{p,1}) & \cos(\theta_{p,2})\sin(\theta_{p,1}) & \cos(\theta_{p,2})\sin(\theta_{p,2})\sin(\theta_{p,1}) & \cdots & \prod_{k=1}^{n-1}\sin(\theta_{p,k}) \end{bmatrix}$$
for $i > j$ angles $\theta_{i,j} \in (0,\pi)$.

So from $R = BB^{t}$, for an individual pairwise correlation we have

 $r_{i,j} = \cos(\theta_{i,1})\cos(\theta_{j,1}) + \prod_{k=2}^{i-1}\cos(\theta_{i,k})\cos(\theta_{j,k})\prod_{l=1}^{k-1}\sin(\theta_{i,l})\sin(\theta_{j,l}) + \cos(\theta_{j,i})\prod_{l=1}^{i-1}\sin(\theta_{i,l})\sin(\theta_{j,l}) \text{ for } 1 \le i < j \le n$

And recursively, in the other direction, for an individual angle $\theta_{i,j}$ we have:

For i > 1: $\theta_{i,1} = \arccos(b_{i,1})$ for j=1; and $\theta_{i,j} = \arccos\left(\frac{b_{i,j}}{\prod_{k=1}^{j-1}}\sin(\theta_{i,k})\right)$ for j > 1

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III.1 Literature: Identity Matrix & Gaussian Data

Correlations to Angles	Angles to Correlations
* INPUT rand_R is a valid correlation matrix;	* INPUT rand_angles is a valid matrix of correlation angles;
cholfact = T(root(rand_R, "NoError"));	Bs=J(nrows, nrows, 0); do i=1 to nrows:
rand corr angles = $J(nrows nrows 0)$.	do i=i to nrows:
do i=1 to nrows:	if i>1 then do:
do i=i to nrows:	if i>i then do:
if i=i then rand corr angles[i,i]=.:	sinprod=1:
else do:	do $ag=1$ to (i-1):
cumprod $sin = 1;$	sinprod = sinprod*sin(rand_angles[i,gg]);
if j=1 then rand_corr_angles[i,j]=arcos(cholfact[i,j]);	end;
else do;	Bs[i,j]=cos(rand_angles[i,j])*sinprod;
do kk=1 to (j-1);	end;
cumprod_sin = cumprod_sin*sin(rand_corr_angles[i,kk]);	else do;
end;	sinprod=1;
rand_corr_angles[i,j]=arcos(cholfact[i,j]/cumprod_sin);	do gg=1 to (i-1);
end;	sinprod = sinprod*sin(rand_angles[i,gg]);
end;	end;
end;	Bs[i,j]=sinprod;
end;	end;
	end;
* OUTPUT rand_corr_angles is the corresponding matrix of angles;	else do;
	if i>1 then Bs[i,j]=cos(rand_angles[i,j]);

SAS/IML code (v9.4)

* OUTPUT rand_R is the corresponding correlation matrix;

else Bs[i,j]=1;

rand_R = $Bs^{T}(Bs)$;

end; end; end;



- These geometric relationships all are well established in the statistical sampling literature, and are key to the method presented later herein.
- <u>Limitations</u>: Of course, all the methods listed above for sampling the correlation matrix, including that of Makalic and Schmidt (2018), have two major limitations: they only are valid when the true correlation matrix is the identity matrix, and they only are valid when the underlying data is Gaussian.
- While the identity matrix is fundamental and useful for testing one specific hypothesis, it arguably is the least interesting or useful for general inference since for any given instance in practice, the occurrence of zero's in all non-diagonal cells is highly improbable. Also, the requirement that data is strictly multivariate Gaussian is quite restrictive.



Case of d=2:

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Taraldsen (2021) derived the exact sampling/empirical distribution of Pearson's correlation under Gaussian data for the pairwise case, i.e. for d=2, as below:

$$\pi(\rho | r) = \frac{\left(1 - r^2\right)^{(\nu-1)/2} \cdot \left(1 - \rho^2\right)^{(\nu-2)/2} \cdot \left(1 - r\rho\right)^{(1-2\nu)/2}}{\sqrt{2}B(\nu + 1/2, 1/2)} \cdot {}_2F_1\left(\frac{3}{2}, -\frac{1}{2}; \nu + \frac{1}{2}; \frac{1 + r\rho}{2}\right) \text{ where }$$

 $B(X,Y) = \left[\Gamma(X) \Gamma(Y) \right] / \Gamma(X+Y)$ the Beta function, v = n - 1 > 1, and F is the Gaussian hypergeometric function where ${}_{2}F_{1}[a,b;c;z] = \sum_{n=1}^{\infty} \frac{(a)_{n}(b)_{n}}{(c)} \cdot \frac{z^{n}}{n!} \text{ where } (h)_{n} = h(h+1)(h+2)\cdots(h+n-1), n \ge 1, (h)_{0} = 1$

- Note that Taraldsen (2021) shows that the approximate density of the pairwise correlation using **Fisher's Z-transformation loses accuracy** as $\rho \rightarrow 1$, especially for smaller samples.
- Case of d>2: Pham-Gia & Choulakian (2014) provide

 $f\left(R\right) = \frac{\left[\Gamma\left(\left[n-1\right]/2\right)\right]^{2} \exp\left\{-\sum_{i < j} \frac{\lambda_{i,j} s_{i,j}}{\sqrt{\sigma_{i,i} \sigma_{j,j}}}\right\}}{\Gamma_{p}\left(\left[n-1\right]/2\right)\left[\left|\Lambda\right|\prod_{i=1}^{p} \lambda_{i,i}\right]^{\left(\left[n-1\right]/p\right)}} \left|R\right|^{(n-p-2)/2}}$ with sample covariance $\{\sigma_{i,i}\}, 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-i\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right), 1 \le i < j \le p, \Gamma_{p}\left(\left[n-1\right]/2\right) = \pi^{\left[p(p-1)/4\right]}\prod_{i=1}^{p} \Gamma\left(\left[n-1\right]/2\right) = \pi^{\left$ and $\lambda_{i,i}$ the diagonals of Λ^{-1}

Limitations: Application of Pham-Gia & Choulakian (2014) above and their equivalent expressions requires a priori knowledge of true (not estimated) variances. It also arguably is quite cumbersome to implement. Finally, it remains valid only under the fairly restrictive case of multivariate Gaussian data.



- Archakov & Hansen (2021) introduce a parameterization of the correlation matrix that maps uniquely, one-to-one, to the positive definite space, thus providing a density for inference. It is based on the Fisher Z transformation, remains invariant to reorderings of the variables, and is accompanied by an algorithm that provides the inverse mapping from the parameterization to the correlation matrix.
- <u>Limitations</u>: The authors state, "This makes the transformation potentially useful for ... inference. These attributes tend to deteriorate as C approaches singularity. This is not unexpected, because it is also true for the Fisher transformation when the correlation is close to ±1."
- As previously noted, Taraldsen (2021) shows that the approximate density of the pairwise correlation using Fisher's Z-transformation loses accuracy as $\rho \rightarrow 1$, especially for smaller samples. This is consistent with the authors' comments here, but they state this may only be material under extreme conditions. All else equal, having a method that avoided this non-robustness issue altogether would be preferrable.
- The method only provides the density of the entire correlation matrix: it does not appear to be able to modify correlation matrices, cell-by-cell, <u>probabilistically</u>, based on their individual densities. Again, this may not be an objective of the method, but all else equal, it is a very useful feature for stress testing and scenario analysis. The method we present herein NAbC has this capability.
- Finally, it is unclear whether the method can be applied successfully to submatrices of the correlation matrix (while holding the rest of the matrix constant). This is relevant as many reasonable scenario specifications would require this. The NAbC method we present herein has this capability also.



- Lan et al. (2020) take a fully Bayesian approach to this problem for both covariance and correlation matrices. Similar to the NAbC method we present herein, they use the Cholesky factorization to maintain positive definiteness, and by defining distributions on spheres as we do, utilize a large class of flexible prior distributions.
- <u>Limitations</u>: This approach is very comprehensive, involved, and includes estimation, which our method does not. However, it appears that for modeling correlation matrices specifically, their approach has some limitations. The authors state: "The priors for correlation matrix specified through the sphere-product representation are in general dependent among component variables. For example, the method we use to induce uncorrelated prior between y_i and y_j (i < j) by setting $l_{jk} \approx 0$ for $k \leq i$ has a direct consequence that $Cor(y'_i, y_j) \approx 0$ for $i' \leq i$. In another word, more informative priors (part of the components are correlated) may require careful ordering in $\{y_i\}$. To avoid this issue, one might consider the inverse of covariance (precision) matrices instead. This leads to modeling the *conditional* dependence, or *Markov network* ... Our proposed methodology applies directly to (dynamic) precision matrices/processes, which will be our future direction."
- The method we develop herein (NAbC) for correlation matrix inference and stress testing/scenario specification can be applied successfully to many submatrices, while holding the rest of the matrix constant, without unintended 'dependencies' and all while automatically enforcing positive definiteness. In fact, our approach largely fixes the 'unintended dependencies' problem that other researchers note, yet fail to control (this is discussed later). This control at the correlation cell level, allowing selective perturbation of many different combinations thereof, is very powerful, useful, and important as flexible scenario specification requires nothing short of this.



- Ghosh et al. (2020) also take a fully Bayesian approach to this problem, and just like our approach, they reparameterize Cholesky factors in terms of hyperspherical coordinates where the angles vary freely in the range [0, π). Their focus is on estimation, although as a Bayesian approach it is comprehensive.
- <u>Limitations</u>: This approach is involved, and includes estimation, which our method does not. However, its use is restricted to parametric priors, which may limit its implementation under complex real world conditions (some of which are empirically implemented herein). In contrast, NAbC makes use of flexible nonparametric kernels that fit ANY angles distribution resulting from ANY data generating mechanism (with finite first and second moments). Also, Ghosh's et al. (2020) approach does not appear to have the capability of modeling submatrices while leaving select cells of the correlation matrix 'untouched.' This control at the correlation cell level, allowing selective perturbation of many different combinations thereof, is very powerful, useful, and important as flexible scenario specification requires nothing short of this. This is one of the advantages of the NAbC method we develop herein and present below.



- Papenbrock et al. (2021) develop a novel approach to simulating correlation matrices for financial markets using evolutionary algorithms. It allows for the flexible yet robust incorporation of many observed features of real world financial correlation matrices. The algorithm scales well and can be used for backtesting, pricing, and hedging correlation-dependent investment strategies and financial products.
- <u>Limitations</u>: This approach does not appear to automatically enforce positive definiteness in its simulations. Enforcing positive definiteness ex post is less efficient, and always introduces systematic changes in the resulting matrices; here it could very well materially alter the very characteristics the algorithm is attempting to embed in the synthetic data. Because the method we propose below (NAbC) automatically enforces positive definiteness while also providing control at the correlation cell level to specify many different combinations of cells for perturbation, it may well be the perfect partner for the evolutionary approach. It could be applied as a very flexible filter for invalidly generated correlation matrices, probabilistically ensuring that the correlation density generated is not systematically biased or skewed in unintended ways. It also could assess the degree to which synthetically generated 'clusters' are outlying, probabilistically.



Opdyke (2020, 2022, 2023) derived the analytic cdf for the pdf used by Makalic and Schmidt (2018): $f_{X}(x) = c_{k} \cdot \sin^{k}(x), x \in (0,\pi), k = 1,2,3... \# \text{columns} -1, \text{ and } c_{k} = \frac{\Gamma(k/2+1)}{\sqrt{\pi}\Gamma(k/2+1/2)} \text{ as...}$ $F_{X}(x;k) \sim \frac{1}{2} - c_{k} \cdot \cos(x) \cdot {}_{2}F_{1}\left[\frac{1}{2}, \frac{1-k}{2}; \frac{3}{2}; \cos^{2}(x)\right] \text{ for } x < \frac{\pi}{2},$ $\sim \frac{1}{2} + c_{k} \cdot \cos(x) \cdot {}_{2}F_{1}\left[\frac{1}{2}, \frac{1-k}{2}; \frac{3}{2}; \cos^{2}(x)\right] \text{ for } x < \frac{\pi}{2},$ where the Gaussian hypergeometric function ${}_{2}F_{1}[a, b; c; r] = \sum_{n}^{\infty} \frac{(a)_{n}(b)_{n}}{(c)_{n}} \cdot \frac{r^{n}}{n!}$ where $(h)_{n} = h(h+1)(h+2)\cdots(h+n-1), n \ge 1, (h)_{0} = 1, \text{ and } |r| < 1, c \ne 0, -1, -2, ...$

But this can be simplified further using two established (if not obscure) identities. simultaneously, which holds in this setting, we have ${}_{2}F_{1}[a,b](a,b](a,b)(a)/r^{a})$

where $B(r; a, b) = \int_{0}^{b} u^{a-1} (1-u)^{b-1} du$ = the incomplete beta function

In addition, we have

$$F_{Beta}(r;a,b) = B(r;a,b)/B(a,b) \text{ where } B(a,b) = \frac{\Gamma(a)\Gamma(b)}{\Gamma(a+b)} = \text{the complete beta function, so}$$
$$B(r;a,b) = F_{Beta}(r;a,b) \cdot B(a,b)$$

* Note that the (Gaussian) hypergeometric function is not uncommon in this setting, making an appearance in derivations of the distribution of individual correlations (see Muirhead, 1982, and Taraldsen, 2021), moments of the spectral distribution under some conditions (see Adams et al. 2018, and https://reference.wolfram.com/language/ref/MarchenkoPasturDistribution.html), and the definition of positive definite functions (Franca & Menegatto, 2022).

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- Taken together we have: $F_{X}(x;k) \sim \frac{1}{2} - c_{k} \cdot \cos(x) \cdot F_{Beta}\left[\cos^{2}(x);\frac{1}{2},\frac{1+k}{2}\right] \cdot \frac{\Gamma\left(\frac{1}{2}\right)\Gamma\left(\frac{1+k}{2}\right)}{\Gamma\left(\frac{2+k}{2}\right)} \cdot \left(\left[1/2\right]/\sqrt{\cos^{2}(x)}\right) \text{ for } x < \frac{\pi}{2},$ $F_{X}(x;k) \sim \frac{1}{2} + c_{k} \cdot \cos(x) \cdot F_{Beta}\left[\cos^{2}(x);\frac{1}{2},\frac{1+k}{2}\right] \cdot \frac{\Gamma\left(\frac{1}{2}\right)\Gamma\left(\frac{1+k}{2}\right)}{\Gamma\left(\frac{2+k}{2}\right)} \cdot \left(\left[1/2\right]/\sqrt{\cos^{2}(x)}\right) \text{ for } x \ge \frac{\pi}{2}$
- Recognizing that the complete Beta function is the inverse of the normalization factor of c(k) for these values, their product equals 1 and cancels, as do the two cosine terms, and we have:

$$F_{X}(x;k) \sim \frac{1}{2} - \left(\frac{1}{2}\right) \cdot F_{Beta}\left[\cos^{2}\left(x\right);\frac{1}{2},\frac{1+k}{2}\right] \text{ for } x < \frac{\pi}{2},$$
$$\sim \frac{1}{2} + \left(\frac{1}{2}\right) \cdot F_{Beta}\left[\cos^{2}\left(x\right);\frac{1}{2},\frac{1+k}{2}\right] \text{ for } x \ge \frac{\pi}{2}$$

• The cdf of the well-known Beta distribution is so straightforward that it readily can be used in a spreadsheet. And now, we can even obtain an analytic* quantile function of the angle distribution:

* Note that we use the term 'analytic' as opposed to 'closed-form' because we are unaware of a closed-form algorithm for the inverse cdf of the beta distribution (see Sharma and Chakrabarty, 2017, and Askitis, 2017). However, for all practical purposes this is essentially a semantic distinction since this quantile function is hard-coded into all major statistical / econometric / mathematical programming languages.





Let
$$p = \Pr(x \ge X)$$
. Then for $x < \frac{\pi}{2}$,
 $p = \frac{1}{2} - \left(\frac{1}{2}\right) \cdot F_{Beta} \left[\cos^{2}(x); \frac{1}{2}, \frac{1+k}{2}\right]$
 $-2p = -1 + F_{Beta} \left[\cos^{2}(x); \frac{1}{2}, \frac{1+k}{2}\right]$
 $1 - 2p = F_{Beta} \left[\cos^{2}(x); \frac{1}{2}, \frac{1+k}{2}\right]$
 $F_{Beta}^{-1} \left(1 - 2p; \frac{1}{2}, \frac{1+k}{2}\right) = \cos^{2}(x)$
 $\sqrt{F_{Beta}^{-1} \left(1 - 2p; \frac{1}{2}, \frac{1+k}{2}\right)} = \cos(x)$
 $\operatorname{arcos} \left(\sqrt{F_{Beta}^{-1} \left(1 - 2p; \frac{1}{2}, \frac{1+k}{2}\right)}\right) = x$

We must reflect the symmetric angle density for $p \ge 0.5$, so we have

$$\begin{aligned} x &= \arccos\left(\sqrt{F_{Beta}^{-1}\left(1 - 2p; \frac{1}{2}, \frac{1 + k}{2}\right)}\right) \text{ for } p < 0.5, \\ &= \pi - \arccos\left(\sqrt{F_{Beta}^{-1}\left(1 - 2\left[1 - p\right]; \frac{1}{2}, \frac{1 + k}{2}\right)}\right) \text{ for } p \ge 0.5 \end{aligned}$$

Note that arcos is arc-cosine, the inverse of the cosine function.



So now we have for the angles distribution, under the Gaussian identity matrix, for the first time together, the pdf, cdf, and (analytic) quantile function:

$$f_{X}(x) = c_{k} \cdot \sin^{k}(x), x \in (0,\pi), k = 1,2,3... \text{#columns} - 1, \text{ and } c_{k} = \frac{\Gamma(k/2+1)}{\sqrt{\pi}\Gamma(k/2+1/2)}$$

 π

$$F_{X}(x;k) \sim \frac{1}{2} - \left(\frac{1}{2}\right) \cdot F_{Beta}\left[\cos^{2}(x);\frac{1}{2},\frac{1+k}{2}\right] \text{ for } x < \frac{\pi}{2},$$
$$\sim \frac{1}{2} + \left(\frac{1}{2}\right) \cdot F_{Beta}\left[\cos^{2}(x);\frac{1}{2},\frac{1+k}{2}\right] \text{ for } x \ge \frac{\pi}{2}$$

$$F^{-1}(p;k) = \arccos\left(\sqrt{F_{Beta}^{-1}\left(1-2p;\frac{1}{2},\frac{1+k}{2}\right)}\right) \text{ for } p < 0.5;$$

 $= \pi - \arccos \left[\sqrt{F_{Beta}^{-1} \left(1 - 2\left[1 - p\right]; \frac{1}{2}, \frac{1 + \kappa}{2} \right)} \right] \text{ for } p \ge 0.5$

Importantly, although often ignored in the sampling literature, note that properly adjusting for sample size, *n*, and degrees of freedom gives

$$k \leftarrow k + n - \# cols - 2$$

Apparently the first (and only other) presentation of this quantile function result comes from an anonymous blog post in March, 2018, although it was obtained via a different derivation, which serves to further validate the result.



^{*} See Xi'an, March, 2018 (https://stats.stackexchange.com/questions/331253/draw-n-dimensional-uniform-sample-from-a-unit-n-1-sphere-defined-by-n-1dime/331850#331850 and https://xianblog.wordpress.com/2018/03/08/uniform-on-the-sphere-or-not/).

In the interest of proper attribution, a reference on the website to the book "The Bayesian Choice" hints that the Xi'an pseudonym is Christian Robert, a professor of Statistics at Université Paris Dauphine (PSL), Paris, France, since 2000 (https://stats.stackexchange.com/users/7224/xian).

- So contrary to the assertions of Makalic and Schmidt (2018), the analytic quantile function now does, in fact, allow for the straightforward use of the inverse transform method for sampling, using a uniform random variate in place of "p" in its formula above, obviating the need for less efficient rejection sampling a la Makalic and Schmidt (2018), or any of the other much more complicated sampling schemes listed above.
- Not surprisingly, code implementing this inverse transform solution, without the need to reject samples, is about 10% faster than that of Makalic and Schmidt (2018), all else equal, as expected. However, based on this presentation, Roman (2023) has implemented the formulae presented herein using specialized Python functions to increase speed to over 30% faster than Makalic & Schmidt's (2018) downloadable code.
- For shorthand, we refer to this density (collectively, the pdf+cdf+analytic quantile function) as "C3" "C"osine, plus the subscripts on the Gaussian hypergeometric function 2+1=3 for "C3."
- Of course, as stated previously, the identity matrix, while fundamental and useful for testing one specific hypothesis, arguably is the least interesting for general inference, as empirical correlation matrices in practice rarely if ever are characterized by non-diagonals that all are zero. Also, the presumption of multivariate Gaussian data is quite restrictive.
- Below we present a method to obtain the finite sample density of the correlation matrix, as well as its inverse ('quantile function'), under general conditions, requiring only the existence of the first two moments (for Pearson's and Spearman's) and positive definiteness (generally).
- Note again that this is conditional on 1. a specified or well estimated correlation matrix and 2. a specified or well estimated data generating mechanism.

IV.2 New Results: General Conditions

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- Estimation is beyond scope herein: our starting point is with a presumably well estimated or specified/known correlation matrix. The estimation literature is rich and includes a very wide range of approaches, often overlapping and combined, including a variety of shrinkage approaches (Ledoit and Wolf, 2003, 2022, Lolas and Ying, 2021, Huang and Fryzlewicz, 2019); regularization (Friedman et al., 2008, Chen et al., 2019, Skachkov et al., 2021, Cui et al., 2016, Lam, 2020); hypersphere decomposition (Li, 2018); robust alternative Cholesky decompositions & ensembles thereof (Lu et al., 2020, Kang et al., 2019); thresholding (Goes et al., 2018, Tanioka et al., 2021) and other robust methods (Nanda et al., 2019; Serra et al., 2018, Yang and Cui, 2019; Auguin et al., 2018; Lounici & Pacreau, 2023); functional shrinkage estimators (Leng et al., 2024); shrinkage with reinforcement learning (Mattera & Mattera, 2023); hierarchical clustering estimators (Garcia-Medina et al., 2023); statistically robust information filtering network estimation (Wang et al., 2023); factor-based regularization (Zhang & Guo, 2024); Cholesky GAS models (Zheng & Ye, 2024); semiparametric under ellipticity (Pelagatti & Sbrana, 2024); dimension-free structured estimation (Puchkin & Rakhuba, 2024); extensions of DCC (Bauwens & Otranto, 2020); total positivity (Agrawal et al., 2020); sparse graphical modeling (Riccobello et al., 2022); CNN (Fang et al., 2021); factorized kernel (Zhang and Li, 2021); generalized autoregressive score (Stollenwerk, 2022); clustering (Begušić and Kostanjčar, 2019); weighted least squares (Li et al., 2022); reinforcement learning (Lu and Simaan, 2022); regular vines (Zhu and Welsch, 2018); coupled regularized, linear pooling, and spatial sign methods (Raninen, 2022); bias-corrected finite-sample observed covariance matrix (Li, 2023); Bayesian & spherical approaches (Lan et al., 2020, Ghosh et al., 2021); combined multiple iterated EWMAs (Johansson et al., 2023); random matrix theory (RMT) (Leidoit & Péché, 2011; Potters & Bouchaud, 2020), and more specifically, rotationally invariant estimators (RIE) (Bun et al., Bun, 2018, Bun et al., 2016) under non-stationarity (Bongiorno et al., 2021).
- In addition to a well-estimated correlation matrix, we presume a well-estimated, or specified/known, data generating mechanism. All responsibly and scientifically defined empirical methods require defining an appropriate range of application based on data structure, so this requirement is part and parcel of the standard scientific approach.
- Herein we continue with the geometric framework described above and established in Pinheiro & Bates (1996), Rebonato & Jaeckel (2000), Rapisarda et al. (2007), and Pourahmadi & Wang (2015).



• As we saw above in the case of the Gaussian identity matrix, for the geometric framework all boils down to the distribution of the angle associated with each pairwise correlation.

Inference:

A crucially important characteristic of these angles distributions is that, unlike the distributions of the correlations themselves, they vary freely, remaining independent of each other (see Pourahmadi & Wang, 2015, Tsay & Pourahmadi, 2017, and Ghosh et al., 2020).

- This means that the multivariate density of angles is simply a function of the products of their individual densities.
- Consequently, not only do the angles define, deterministically, the correlation values themselves via $r_{i,j} = \cos(\theta_{i,1})\cos(\theta_{j,1}) + \prod_{k=2}^{i-1}\cos(\theta_{i,k})\cos(\theta_{j,k})\prod_{l=1}^{k-1}\sin(\theta_{i,l})\sin(\theta_{j,l}) + \cos(\theta_{j,i})\prod_{l=1}^{i-1}\sin(\theta_{i,l})\sin(\theta_{j,l}) \text{ for } 1 \le i < j \le n$

and thus each of the marginal distributions of the pairwise correlations (see Pourahmadi & Wang, 2015 and Ghosh et al., 2020), but also determine in a very straightforward way the probability of observing an entire correlation matrix (this probability is equivalent to a p-value, and described below).

- All we need now is the distribution of the angles under general conditions, just as we derived them above for the (Gaussian) identity matrix.
- This remains an open, and apparently non-trivial problem, at least analytically. The RMT literature shows that spectral distributions change notably and nontrivially based on the characteristics of the underlying data ensembles (e.g. heavy-tailedness (see Burda et al., 2004, Burda et al., 2006, Akemann et al., 2009; Abul-Magd et al., 2009, Bouchaud & Potters, 2015, Martin & Mahoney, 2018), and serial correlation (see Burda et al., 2004, 2011)), indicating that deriving a spectral distribution valid for all cases would be challenging, if at all possible. A similar derivation for angles distributions likely would be as challenging, and is not found in the extant literature.



- However, while an "all-cases" analytic derivation for the angles distribution would be very useful, not having it is not a show stopper: we can proceed with nonparametric kernel density estimates of these distributions. The process has five steps:
- 1. Simulate samples (say, N=10k) based on the specified/known or well estimated data generating mechanism.
- 2. Calculate the corresponding N correlation matrices and their Cholesky factorizations, and transform each of these into a lower triangle matrix of angles (as described above on pp15-17).
- 3. Fit kernel densities to the p(p-1)/2 empirical angle distributions, each having N observations.
- 4. Generating samples from these densities in 3. is identical to perturbing the actual angle datapoints from 2. based on the fits in 3.
- 5. All samples from 4. are converted back to a re-parameterized Cholesky factorization (per pp.15-17), and then multiplied by its transpose to obtain a set of N validly sampled correlation matrices. Positive definiteness is enforced automatically as the Cholesky factor places us on the <u>unit</u> hyper-hemisphere.
- Throughout the rest of this paper, we exhaustively compare various metrics based on 3. vs. those based on 4., as 3. is the empirical 'truth' against which we are testing the validity of the samples generated in 4. (technically, 4. and 5.). Once 3. is generated, further sampling from 4. is many orders of magnitude faster than turning to 3. again, although the two are equivalent empirically. However, 4. allows us to isolate the distribution of the correlation matrix itself, separate from any other characteristics of the data, which is the point of the methodology.
- We dub the above sampling method via 4. "NAbC" Nonparametric Angles-based Correlations.



• All samples generated from Step 4. (i.e. NAbC) will yield only positive definite correlation matrices.

Correlation matrix sampling distribution + quantile function via NAbC:

- A. Using NAbC, we can now specify any (positive definite) correlation matrix of the same dimension p and use the densities from 4. (or 3.) empirically to obtain the unique cdf's associated* with each of its correlation cells: simply translate the correlation matrix to its matrix of angles, match each angle to the closest angle** in the corresponding empirical distribution of 4., return the associated empirical cdf.
- B. Conversely, if we specify cdf's for each of the correlation cells, we can obtain the inverse ('quantile function') i.e. the unique correlation matrix associated with those cdf's. Simply 'lookup' the angles in 4. (or 3.) associated with each specified cdf, translate the resulting matrix of angles to its reparameterized Choleksky factor, and then multiply by its transpose to obtain the corresponding correlation matrix.
- The simple rule is this: <u>any sampling or perturbation must be done after translating correlations or</u> <u>cdf's to angles to enforce positive definiteness – sampling/perturbation can never be performed on the</u> <u>correlations themselves</u> as the resulting matrix almost certainly will not be positive definite.
- In either A. or B. above, the independence of the individual angles distributions leads to both a novel matrix-level 'distance metric' and a straightforward matrix-level p-value, conditional on the observed/estimated matrix, as described below.

^{**} Technically this is a match with the closest angle with a cdf at least as large as the specified cdf, but empirically, for N samples ≥ 10k, these are the same.



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^{*} Note that because this 'look up' is based on an empirically generated density, near-singular extreme examples sometimes may not have been simulated, thus preventing a successful 'look up.' For these more extreme cases, the n=simulations must be increased to contain additional, more 'extreme' cases. But this is very rarely required beyond n=10k due in part to the bounded range (from 0 to π) of the angles as opposed to, say, the unbounded range of the corresponding spectral distribution.

IV.2 New Results: General Conditions

- To define the matrix-level p-value, we first need to calculate the p-values for each of the cells. This requires calculating the empirical mean correlation matrix from the simulations in 3., and translating this to a matrix of angles. Then we obtain the empirical cdf's of these "mean-angles" with a 'look up' to the entire angles distributions from 3. Often these will be close to 0.5, but asymmetry can be notable in some cases (see APPENDICES) and must be properly accounted for as below.
- The cell-level p-values (2-sided) are based on the distance between the cdf's of the angles of the specified correlation matrix and those of the "mean-angles," and the sum of the (usually two) tails BEYOND this distance (under notable asymmetry, this is sometimes just one tail – see Graph 1 below). Below are the tail probabilities for two cells with the same "mean-angles" cdf's and distributions, but different specified correlation values/cdf's:







 Note that while a cdf=0.1 is hardly more 'extreme' than a cdf=0.85 in absolute terms, relative to the mean cdf=0.6, it is twice as 'extreme,' i.e. twice as far from mean cdf=0.6, and associated with only 1/5 the probability of being observed.



IV.2 New Results: General Conditions

- Now, all the cell-level p-values define the matrix (2-sided) $pvalue = \left[1 \prod_{i=1}^{p(p-1)/2} (1 pvalue_i)\right]$ which is one minus the probability of no false positives, thus controlling the entire matrix's family-wise error rate (note that due to angle independence, there exists no dependence between the cell-level p-values).
- Both the cell-level and matrix-level p-values are very useful. The latter can be used for hypothesis testing and monitoring correlation matrices in statistical process control frameworks (WITHOUT a computationally expensive Phase 1 burn-in period), while the former can be used as a distance metric which has some advantages over widely used norms (e.g. Frobenius/Euclidean, Chebyshev, and taxi norms). The 'distances' of the cell-level p-values are probabilities, not absolute distances. Norms do not recognize that an absolute distance means different things depending on the value of the particular correlation cell. A shift of 0.01 from an original correlation value of 0.5 means something very different than the same shift from 0.98, and the cell-level p-value recognizes this, giving the 'distance' of a specified correlation matrix from an estimated one probabilistic meaning.
- We can take the product of the cell-level p-values as an overall measure of 'distance,' but this value, even in relatively small matrices (say, p=10x10), becomes very small very quickly (e.g. p(p - 1)/2 =45, and taking all cells as having 0.25 tail probabilities as an example, p-value = 0.25^45 = 8.07794E-28). So for practical purposes we take the log of this product of p-values by simply summing the individual logs (which for this example, "LNP" = -62.38).

$$LNP'' = \ln\left(\prod_{i=1}^{q} pvalue_i\right) = \sum_{i=1}^{q} \ln\left[pvalue_i\right] \text{ where } q = p(p-1)/2$$

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- LNP has some interesting properties worth further investigation.
- When the correlation matrix is the identity matrix, LNP aligns extremely well with the entropy of the matrix.
- As defined in Felippe, et al. (2021 and 2023), entropy simply is the Shannon entropy using the eigenvalues of the correlation matrix after it has been scaled by its dimension size, p.

Entropy = $Ent(R/p) = -\sum_{j=1}^{p} \lambda_j \ln(\lambda_j)$ where *R* is the sample correlation matrix and λ_j are the p eigenvalues of the correlation matrix scaled by its dimension, *R/p*.

- Note that, like NAbC, this result from Felippe et al. (2021 and 2023) is valid for ANY positive definite measure of dependence.
- Below I show the relationship between LNP and entropy via 10,000 simulations under the gaussian identity matrix (dim=p=5). The resulting Pearson's correlation is just shy of 0.99.
- This relationship is a very intriguing result. It allows for the interpretation of LNP as a type of "generalized entropy" in the sense that it is not restricted to the baseline of the identity matrix, which serves as the maximum entropy upper bound with a value of ln(p) (while perfect dependence, ie a matrix of 1.0's, serves as the minimum entropy lower bound of zero).
- In contrast, LNP can be calculated using ANY correlation matrix as its reference baseline, not just the identity matrix, so it is a more general measure of matrix distance/dispersion.
- Yet the correspondence of LNP to entropy under the identity matrix speaks to its natural interpretation as a meaningful measure of deviation/distance/dispersion, and one that also is more flexible and granular than entropy as it is measured cell-by-cell, p(p-1)/2 times, as opposed to only p times for p eigenvalues.

IV.2 New Results: General Conditions

Identity Matrix: LNP v Entropy



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- Alternately, the sum or average of the cell-level p-values can be used as a matrix-level distance metric, but all results herein are based on LNP.
- Because LNP is a monotonic transformation, the relative rankings of different correlation matrices compared to the estimated/observed one will be preserved. And unlike norms, its scale is not dependent on sample size, although it is dependent on the dimension of the matrix, p.
- Comparisons of the empirical distribution of LNP to those of common norms are made below in the Appendices.
- Appendices. • For completeness, a norm is a distance metric defined as: $\|x\| = \left(\sum_{i=1}^{d} |x_i|^m\right)^{1/m}$...Taxi, Frobenius/ Euclidean, & Chebyshev norms correspond to m=1, 2, and ∞ , respectively (with the latter \rightarrow max[x]).
- The above treats p-values, both at the matrix level and at the cell level. Regarding confidence intervals, at the cell-level take $\alpha/2$ and $1-\alpha/2$ and simply use "look ups" on the empirical cdf for each cell to obtain the corresponding correlation values for the $(1-\alpha)$ confidence interval.
- At the matrix level, the calculation of **simultaneous confidence intervals** that is, two correlation matrices, one "low" and one "high", representing bounds of, say, $(1-\alpha) = 0.95$ probability that all cell values are contained between the two matrices is greatly simplified by the fact that all the angles distributions are independent (just as with the p-value): $\alpha_{crit-simult-LOW} = \left(1 - \left[1 - \alpha/2\right]^{(1/[p(p-1)/2])}\right)$ • The two new α_{crit} values are then used in "look ups" $\alpha_{crit-simult-HIGH} = 1 - \alpha_{crit-simult-LOW}$
- The two new α_{crit} values are then used in "look ups" $\alpha_{crit-simult-HIGH} = 1 \alpha_{crit-simult-LOW}$ on the empirical cdf for each cell to obtain the two corresponding correlation matrices. Note that this corresponds with the p-value calculation (although α is rightly divided by two (to make it one-tailed), and the #cells root is used, rather than its power, to obtain the critical value rather than the p-value).



- Note that as the density of the angle distribution is bounded θ∈ (0,π), the kernel must be appropriately reflected at the boundary via: (see Silverman, 1986) if θ < 0 then θ ← -θ; if θ > π then θ ← (2π θ) ... note that for dependence measures whose domain is 0 to 1.0 instead of -1.0 to 1.0, eg the tail dependence matrix, the corresponding range of angle values is 0 to π/2, and reflection at the boundary is adjusted accordingly
- Both the Gaussian kernel and the Epanechnikov kernel have been tested extensively, along with three different bandwidth estimators, *h*, from Silverman (1986):

$$f_{h}(\theta) = \frac{1}{N} \sum_{i=1}^{N} K_{h}(\theta - \theta_{i}) = \frac{1}{hN} \sum_{i=1}^{N} K_{h}([\theta - \theta_{i}]/h)$$

$$h = 1.06 \cdot \hat{\sigma} \cdot N^{-1/5}$$

$$h = 0.79 \cdot IQR \cdot N^{-1/5}$$

$$h = 0.9 \cdot \min(IQR/1.34, \hat{\sigma}) \cdot N^{-1/5}$$

$$h = 0.9 \cdot \min(IQR/1.34, \hat{\sigma}) \cdot N^{-1/5}$$
where $\hat{\sigma}$ = sample standard deviation and IQR = sample interquartile range

- The Epanechnikov kernel with $h = 0.9 \cdot \min(IQR/1.34, \hat{\sigma}) \cdot N^{-1/5}$ shows a slight performance edge, when comparing metrics of 3. vs. 4., across a wide range of settings, all else equal. This is likely due to the fact that in many cases, its bounded sample space will more often avoid the angle boundaries.
- Bandwidths *h* typically are the most important issue when using kernels. Here, for larger matrices (e.g. p=100x100), bandwidths need to be tightened by multiplying by a factor of 0.15 to avoid a slight drift in the density based on the large number of density estimations performed (i.e. p(p 1)/2 = 4,950). Multiplying by this factor for smaller matrices does not adversely affect the density estimation in any way, so this factor always is used. For matrices much larger than p=100, a further tightening of this factor may be required.



- <u>Limitations</u>: NAbC is relatively **computationally intensive**, **but not prohibitively so**. On a commodity laptop with RAM=32GB but with no multi-threading, NAbC code processes a 100x100 matrix, with N samples = 10,000, in about 2.4 hours. However, in a multithreaded environment, let alone one with more memory, NAbC could be applied on similarly non-small matrices in minutes.
- Applying inverse probability transform sampling using the "C3" analytic quantile function derived herein for the specific case of the Gaussian identity matrix (no Step 3. data generated), a 100x100 matrix with N samples = 10,000 takes under 25 minutes to run on the same laptop.
- Notably, however, Roman (2023), following this analytic solution for the Gaussian identity matrix, has implemented it over 30% faster than Makalic & Schmidt (2018)).
- Also note that once the angles distributions are generated only once in Step 3., using Step 4. to obtain the correlation matrices for specified cdf's ('quantile function'), or obtaining cdf matrices for specified correlation matrices, or obtaining both matrix-level and cell-level p-values and confidence intervals, is extremely fast. For a 100x100 matrix, generating 100 of each requires only about 2 minutes in total.
- So we would argue that any disadvantage associated with NAbC not being purely 'real time' on nonsmall matrices is outweighed by two factors: i. its generality – it can be applied under the most general conditions possible (ie ANY data conditions, as long as the matrix is positive definite) – and ii. its unmatched flexibility – it can be applied to numerous submatrices while holding all other correlation cells constant (this is discussed in the following sections). No other method can make these claims.



- Another arguable limitation is that NAbC does not include estimation, only inference and scenarios and stress testing.
- However, the extant literature on estimation already is very extensive, so NAbC fills a hole in the literature where there is one to fill. Specifically, it unifies inference across previously very disparate methods, as NAbC can be applied in exactly the same way to ANY positive definite measure of dependence/concordance, including Kendall's tau, Spearman's Rho, and others. No other method can make this claim. This is discussed, and empirical results presented, further below.
- We now examine NAbC results from toy cases (p=5x5) as well as p=100x100 from some of the below-listed data generating mechanisms (DGMs). We focus on comparing results based on Step 3. vs. those based on Step 4. Note that the angles graphs are meant to be toggled to more accurately see differences and similarities. Empirical results for Kendall's and Spearman's follow at the end.



- We exhaustively test to ensure that Steps 3. and 4. are, in fact, the same distributions* by comparing:
 - the angles distributions of 3. vs. 4.
 - the spectral distributions of 3. vs. 4.
 - the distributions of Euclidian/Frobenius (and other) norm(s) of 3. vs. 4.
 - the distributions of LNP = ln(p-value) of 3. vs. 4.
 - the risk (VaR-based aggregated loss / economic capital) distributions of 3. vs. 4.
 - the mean empirical correlation matrices of 3. vs. 4.
 - a scatterplot and Pearson's correlation between LNP and the Euclidian/Frobenius norm (4.)
- Under the following data generating mechanisms (DGMs):
 - MVU Multivariate Uniform
 - MVG Multivariate Gaussian (Gaussian marginals)
 - MVGM Multivariate Gaussian (varying marginals)
 - MVGTS Serially correlated approximate Multivariate Gaussian
 - MVT Multivariate student's t
 - MVTM Multivariate student's t copula with varying marginals
 - MVTV Multivariate Asymmetric student's t copula with varying degrees of freedom, varying marginals and varying asymmetry
 - MVTVTS Serially correlated approximate MVTV
 - MVTVNS Non-stationary approximate MVTVTS
 - Archimedean Copulas (Gumbel, Frank, Clayton) with Lognormal(0,1) marginals.
- See Appendix 1 for graphical and tabular results. A summary follows below.

* Note that showing 3. and 4. are distributionally identical validates this entire approach. But in practice, we always would want to use 4. because sampling the empirical angles densities already generated in 3. once, by using 4., is orders of magnitude faster than re-simulating them via 3. And of course, 4. isolates the density of the correlation matrix without touching any other aspect of the data.



- Empirical Results from NAbC Notes (see APPENDIX 1 for Graphs/Tables):
- NOTE: All results show a comparison of Step 3. vs. Step 4. as validation of the latter EXCEPT Case A using the C3 quantile function (because no data is being simulated the method is analytical).
- Where noted results are for p=100x100, but otherwise, for illustrative purposes all are for p=5x5 (sample size n=p+1, and n=126 for 6 months of trading days; for p=100, n=p+1 and n=252 for a year of trading days). Kernel bandwidth $h = 0.9 \cdot \min(IQR/1.34, \hat{\sigma}) \cdot N^{-1/5}$ version of Silverman (1986) x 0.15. Results include distributions of the angles, eigenvalues, Euclidian/Frobenius norm, LNP, square-root-rule elliptical aggregated 'capital,'* and the differences between all the cells in the two empirical mean correlation matrices. A scatterplot and Pearson's correlation between LNP and the Euclidian/Frobenius norm also is presented. All tables of matrices are presented in lower triangle column format. The matrices include conversions from i. specified correlation matrices to cdf matrices**; ii. specified cdf matrices to unique correlation matrices; and iii. specified %CDF-shifts from the empirical mean matrix to correlation matrices (wherein each cell's %shift is the percentage of the cumulative density above or below the cdf of the mean angle). Finally, angle number is determined based on the following fill-order (the motivation for this is discussed in a later section):

Rightmost Triangle Fill Order

2						
	11					
	12	7				
	13	8	4			
	14	9	5	2		
	15	10	6	3	1	

** Note that the inverse relationship between angles cdf's and correlation values applies here: for the %CDF-shift tables, this has been reversed so that positive/negative CDF shifts correspond with increases/decreases in correlation values.

* Under multivariate ellipticity, extreme portfolio quantiles (aka Value-at-Risk, "VaR", aka "Capital") can be calculated asymptotically with the 'square root rule,' where portfolio VaR=Capital = $\sqrt{VRV^t}$ where V = vector of p VaRs (quantiles) and R = correlation matrix of dimension p (see Tao et al., 2019, and Frachot et al., 2001). While this does not hold under other distributions, its distribution under Step 3. still should be identical to that of Step 4., so they are presented for comparison purposes. All 99VaR's are set to \$50m herein: when different marginals are used, they are scaled so that 99VaR = \$50m.

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- Empirical Results from NAbC Example Cases (see APPENDIX 1 for Graphs/Tables):
- A. C3 v MVG Gaussian Identity Matrix: side-by-side comparison of use of fully analytic C3 quantile function (no data generation) vs. a multivariate gaussian data generating mechanism (n=6, n=126)
- B. MVG with all ρ = 0.2 (n=6, n=126)
- C. MVG with ρ = 0.4 for one cell (#5) v. ρ = 0.4 for one factor (row 4, cells 2,5,9) (n=6)
- D. MVT (df=3) with all ρ = 0.2 (n=6, n=126)
- E. MVTM MVT (df=3) with different marginals (2 LN(0,1), 3 Gamma(1,1)), all ρ = 0.2 (n=6, n=126)
- F. MVTVNS (ρ = block structure, df=3,4,5,6,7, skewness=1,0.6,0,-0.6,-1, autocorrelation= -0.25,0,0.25,0.5,0.75, nonstationarity=3 σ , σ /3, σ , n/3 each) (n=6, n=126) **1** -0.1
- G. MVTVNS (all *ρ* = 0.2, **p=100x100**, df=3 to 27.5 by 0.5, skewness=1 to -0.96 by 0.04, autocorrelation=0.68 to -0.3 by 0.02, nonstationarity=3σ, σ/3, σ, n/3 each) (n=101, n=252)

26)	1	-0.1	-0.1	0.2	0.2
,	-0.1	1	-0.1	0.2	0.2
Case F.: <i>R</i> =	-0.1	-0.1	1	0.2	0.2
	0.2	0.2	0.2	1	0.5
	0.2	0.2	0.2	0.5	1

- H. Gumbel Copula (ϕ = 1.25), as upper tail dependence is relevant in this setting (n=6, n=126)
- NOTE: Cases of n ≤ p are not within the scope of this work, as NAbC requires positive definite, and thus, full rank matrices.



- <u>Empirical Results from NAbC Summary (see APPENDIX 1 for Graphs/Tables)</u>:
- All results are distributionally identical for Steps 3. v 4., indicating validity in NAbC's methodology.
- All spectral distributions for both 3. and 4. match those defined in the RMT literature (e.g. Marchenko-Pastur distribution (Marchenko & Pastur, 1967) under the identity matrix, and otherwise where spectral distributions are derived analytically (e.g. tests not included herein were performed using Lillo & Mantegna, 2005, and Livan et al., 2011 as examples). Also the effects of heavy-tailed distributions and serial correlation appear consistent with those of the RMT literature generally (see for the former Burda et al., 2004, Burda et al., 2006, Akemann et al., 2009; Abul-Magd et al., 2009, Bouchaud & Potters, 2015, Martin & Mahoney, 2018; and for the latter, see Burda et al., 2004, 2011).
- The relationship between LNP and the Euclidian/Frobenius norm is consistent with the latter's inability to distinguish between the RELATIVE distances from the estimated/given correlation matrix: they are very similar under the Gaussian identity matrix (especially under larger sample sizes), where relative and absolute distances are more similar, but when the original correlation matrix is asymmetric or contains more extreme values, the strength of their association diminishes somewhat, as expected.
- Note that the scale of the LNP distribution is not dependent on sample size, whereas that of the Euclidean/Frobenius and other norms do change with sample size. The scales of both are functions of the dimension of the matrix.
- On tests not presented herein, LNP remains numerically robust even under p=100x100, and readily provides rankings of 'distance' for all specified correlation and cdf matrices.



• For completeness, we include the pdf of the Marchenko-Pastur (1967) distribution below:

 $f(x) = \frac{1}{2\pi\sigma^2} \frac{\sqrt{(\lambda_+ - x)(\lambda_- - x)}}{\lambda x} \text{ for } \lambda_- \le x \le \lambda_+, \text{ and zero otherwise, when } 0 \le \lambda \le 1,$ and when $\lambda > 1$ we have an additional mass point of $\left(1 - \frac{1}{\lambda}\right)$ at x = 0

where
$$\lambda_{\pm} = \sigma^2 \left(1 \pm \sqrt{\lambda} \right), \ \lambda = \frac{p}{n}, \ p = \# \text{factors/rows/columns}, \ n = \text{sample size},$$

and $\sigma = 1$ for correlation matrices

- This analytic result overlays the empirical spectral distributions presented in the appendices.
- Note that exceptions to convergence to this celebrated distribution do exist (see Li and Yao (2018), Hisakado and Kaneko (2023), and Maltsev and Malysheva (2024) for examples).



- Empirical Results from NAbC Summary (see APPENDIX 1 for Graphs/Tables):
- <u>Case A</u>: shows distributionally identical results when comparing those generated from 3. vs. those based on 4. Also, NAbC applied to multivariate gaussian data yields distributionally identical results vs. probability inverse transform sampling using the analytic C3 quantile function derived herein, thus validating the latter empirically.
- <u>Case B</u>: shows identical results distributionally for 3. v 4., with the expected effects for sample size and for non-zero correlations on both the spectral and angles distributions.
- <u>Case C</u>: demonstrates how eigen-structure/decomposition approaches are too blunt a tool for analyzing correlation matrices. We start with the identity matrix and examine the spectral distributions of two cases: one where an entire factor/row (cells 2, 5, and 9) is assigned $\rho = 0.4$ vs. one where only a single correlation cell (#5) gets $\rho = 0.4$. Under small sample sizes (n=p+1) they look virtually the same. The angles distributions, however, show the story very clearly, for each of the right cells, and the untouched zero cells show exact matches with both the defined pdf (sin^k) and the C3 analytic derivation. As sample size increases, as expected, the spectral distributions can more readily be distinguished, but still do not provide the unambiguous story provided by the angles distributions. Note also how cell #1 (in row 5) is (unintentionally) affected by the changes of the factor corresponding to row 4 due to the rightmost cell-change rule described on pp. 47-49. NAbC is the first method able to explicitly control that kind of unintended correlation 'contamination.'
- <u>Case D</u>: shows the incremental effects of heavy tails, all else equal, compared to Case B. NAbC still matches empirical 'truth' perfectly, by every distributional criteria.
- <u>Case E</u>: shows the incremental effects of different marginals, all else equal, compared to Case D.
 NAbC still matches empirical 'truth' perfectly, by every distributional criteria.
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- Empirical Results from NAbC Summary (see APPENDIX 1 for Graphs/Tables):
- <u>Case F</u>: shows a more 'real world' example (with a block correlation matrix) including heavy tails (varying by factor), skewness (varying by factor), autocorrelation (varying by factor), and nonstationarity. We see, again, identical results distributionally for 3. v 4. across all comparison criteria. Examination of both the spectral and angles distributions in this case shows complex structure, providing further evidence that deriving an 'all-cases' analytic density for either appears to be nontrivial.
- <u>Case G</u>: shows Case F. but for a non-toy matrix of p=100x100 and a constant correlation matrix with $\rho = 0.2$. Again, NAbC still matches empirical 'truth' perfectly, by every distributional criteria.
- <u>Case H</u>: shows a copula with upper tail dependence (Gumbel) and we see, again, identical results distributionally for 3. v 4. across all comparison criteria.
- These extensive results confirm empirically the assertions of Fernandez-Duran & Gregorio-Dominguez (2023), and Zhang & Songshan (2023): that the angles distributions preserve the entire dependence structure represented by the correlation/dependence measures themselves, thus making them the key to a single, unified approach to all inference and scenario/stress testing of such measures (as well as COMPARISONS across such measures).



<u>Empirical Results from NAbC – Code</u>:

- NOTE: Complete SAS/IML (v9.4) code that generates the above results and many possible specifications for a wide range of input parameters, including input .csv files (e.g. the specified cdf matrices, correlation matrices, %CDF-shift matrices, the baseline correlation matrices, eigenvalues for defining baseline correlation matrices, etc.) will be made publicly available for download on GitHub and in the forthcoming book, "Beating the Correlation Breakdown, for Pearson's and Beyond: Robust Inference and Flexible Stress Testing and Scenarios for Financial Portfolios," Elements in Quantitative Finance series, Cambridge University Press, eds. Ricardo Rebonato, PhD.
- The SAS/IML code has extensive functionality. In addition to the optionality mentioned above, correlation matrices can be specified either as correlation matrices, or defined by eigenvalues (which are converted via Givens rotations a la Davies and Higham, 2000), matrix re-orderings for targeted submatrices can be specified as an input parameter (either as a list or in a .csv file), scales of axes on graphs and kernel bandwidths on the kernel-based graphs are input parameters, the number of angles graphs generated is a parameter, etc.



- NAbC is more robust than its competitors in several ways. Why is this the case?
- Almost all of NAbC's competitors, whether designed for stress testing or making inferences about an observed / estimated correlation matrix, focus on its eigen structure: either its spectral distribution (i.e. the distribution of its eigen values), or they explicitly manipulate its eigen decomposition (Archakov & Hansen (2021) is a notable exception).
- Unfortunately, in practice, the larger the matrix, the more likely it is to approach singularity, i.e. non-positive definiteness (NPD).
- So these 'eigen' approaches become only as good as the matrix is 'far' away from being NPD, which for most practical purposes of financial portfolio analysis, often is not very 'far.'
- For example, Hardin et al. (2013) (responsibly) acknowledge that, in their eigen value perturbative approach, "The amount of noise that can be added to the original matrix is determined by its smallest eigenvalue. ... We provide the user with ... a general algorithm to apply to any correlation matrix for which the smallest eigenvalue can be reasonably estimated." (emphases added).
- This constraint is true of these types of approaches generally, but it is seldom acknowledged.
- Unfortunately, this eliminates some of the most widely observed correlation matrices in finance those close to a 'spiked' covariance matrix (see Johnstone, 2001) where one or few eigenvalues dominate and the majority of eigenvalues are close to zero, i.e. not reliably estimated.



- Also, eigen approaches remain at the level of the factors in the portfolio (i.e. only p factors), NOT all the pairwise associations between these factors (i.e. p(p-1)/2 associations). This is the wrong level of aggregation for flexibly analyzing the correlation matrix, cell-by-cell. Using only p eigenvalues/vectors/factors simply is too blunt a tool for inference regarding all of their p(p-1)/2 pairwise associations, at the level of the correlation cell.
- So while eigen decompositions can be indispensable for things like identifying non-random effects/factors in a portfolio, they are much less so when examining the correlation matrix per se.



NAbC for Robust Inference:

- <u>Structurally</u>: Level of aggregation becomes important here. There are many more angles distributions than there are spectral distributions (i.e. p(p-1)/2 vs p, a factor of (p-1)/2 more). As a matrix approaches NPD, a much smaller <u>proportion</u> of angles distributions will approach degeneracy than is true for eigenvalue distributions. Consequently, the overall construction of the correlation matrix via $R = BB^T$ will remain much more stable than one based on an eigendecomposition of $R = V\Lambda V^{-1}$ where V is a matrix with column eigenvectors and Λ is a diagonal matrix of the corresponding eigenvalues.
- <u>Empirically</u>: Even if an angle distribution approaches degeneracy, all its values will simply approach 0 or π . But the trigonometric functions of these values are stable, and will simply approach 0 or 1. This makes $R = BB^T$ a stable calculation empirically, even if it produces an R that is approaching NPD. Eigenvalue/vector estimations are more numerically involved and comparatively less stable as matrices approach NPD.
- <u>Distributionally</u>: The distributions of angles are well behaved: they are continuously differentiable (smooth), unimodal, and clearly bounded on $\theta \in (0, \pi)$. Spectral distributions are unbounded (in the general case) and thus characterized by larger variances and less tail accuracy. They also are more complex in the general case.
- All of this adds up to a more robust basis for inference when relying on the geometric framework of angles distributions as opposed to spectral distributions. And for examination of the correlation matrix per se, spectral distributions simply are not at the right level of aggregation: they are indispensable for factor analysis, but utilizing the 'eigen' information on only p factors to analyze p(p-1)/2 pairwise associations of those factors simply is too blunt a tool.



Flexible Stress Testing and Scenarios:

- Many approaches to stress testing perturb the underlying distributions to the correlation matrix (e.g. Kupiec, 1998; Ng et al., 2013; Zhang et al., 2015; Packham & Woebbeking, 2019), but this is not sufficient, just as it would not be sufficient for understanding the density of any other parameter in a portfolio model.
- It is important to reemphasize here that stressing the correlation matrix directly, as a model parameter, is not inconsistent in any way with concurrently stressing the other parameters in the model and/or the underlying marginal distributions. The Bank for International Settlements (BIS) recommends doing both concurrently.
- <u>BIS-BCBS, 2011</u>, pp. 28-29: "However, in order to calculate stressed VaR accurately it is also necessary to stress the correlation matrix used in all VaR methodologies. ... In general, most correlations tend to increase during market crises, asymptotically approaching 1.0 during periods of complete meltdown, such as occurred in 1987, 1998 and 2008. ...<u>Certain methods that could be</u> <u>meaningful in this context can be identified in the earlier literature on stress testing. Employing fattailed distributions for the risk factors and replacing the standard correlation matrix with a stressed one are two examples." (emphasis added).
 </u>
- Unfortunately, when it comes to stressing the correlation matrix directly, the literature proposes many methods that exhibit demonstrable inaccuracies and deficiencies, and unlike NAbC, none are based on probabilistic control over the individual correlation cells or the entire matrix, let alone both simultaneously.

• In a widely circulated paper, Galeeva et al. (2007) propose several approaches, one of which is perturbing the polar angles directly via:

$$\hat{\theta}_{i,j} = \arctan\left(\tan\left(\theta_{i,j} + \frac{\pi}{2}\right)\left(1 + \sigma z_{i,j}\right)\right) + \frac{\pi}{2} \text{ where } z_{i,j} \sim N(0,1)$$

"with the goal of generating random angles around the base angles with some distribution which is symmetric and centered around the base-correlation $\theta_{i,j}^0$ [angle] for every i.j"

- This approach has several problems. First, the choice of a Gaussian distribution here is arbitrary, and it is not proportional to the determinant of the Jacobian as described above (see Pourahmadi & Wang, 2015). We can immediately see that the proposed distribution is not a function of the angle's column position in the matrix, which is the structure imposed by the determinant of the Jacobian. Also, in the general case the distribution is, of course, not necessarily symmetric, and is often skewed.
- Another approach proposed by Galeeva et al. (2007) is to perturb the eigenvalues by the exponentiated Gaussian distribution (i.e. Lognormal), using the historical variances as parameters in the distribution. Here the choice of the Lognormal is arbitrary, and simply does not work when matrices approach NPD: with eigenvalues ≈ 0, there is nothing to perturb.
- Other approaches proposed by Galeeva et al. (2007) (e.g. bootstrapping) simply do not preserve positive definiteness. Neither does the approach proposed in So et al. (2013).
- Ho (2015) provides a approach using empirical likelihood to modify the probability weights of sample observations to construct a stress correlation matrix. While this has advantages in its statistical interpretation (in the K-L divergence sense), non-parametric estimation, and straightforward computation, it does not provide the perfect control at the correlation cell level that NAbC provides. Neither does Loland et al., (2013).



- Hardin et al. (2013) utilize a normalized vector of independent gaussian random variables to perturb the observed correlation matrix. While relatively straightforward to implement, a nontrivial limitation is acknowledged by the authors: "The amount of noise that can be added to the original matrix is determined by its smallest eigenvalue. ... We provide the user with ... a <u>general algorithm to apply to</u> <u>any correlation matrix for which the smallest eigenvalue can be reasonably estimated</u>." (emphases added)
- Unfortunately, this eliminates some of the most widely observed correlation matrices in finance those close to a 'spiked' covariance matrix (see Johnstone, 2001) where one or few eigenvalues dominate and the majority of eigenvalues are close to zero, i.e. not reliably estimated.
- The correlation parameterization of Packham and Woebbeking (2021) does not automatically enforce positive definiteness, and 'nearest correlation' enforcement adjustments (e.g. Higham, 2002) do systematically alter both the spectral and angles distributions. It also relies on Fisher's ztransformation, which Taraldsen (2021) has shown to sometimes be inaccurate under extreme values.
- The market stress approach of Chmielowski (2014) is a notable exception to many of the limitations listed above as it remains is explicitly invariant under a change of basis of risk factors. Parlatore and Phillippon's (2022) Kalman Filter approach in the same setting systematically and quantitatively incorporates correlation priors. But both stop short of the granular, probabilistic control needed for flexible stress testing/scenarios.
- So none of these methods provide the highly granular, probabilistic control over the correlation matrix directly that is required for flexible stress and scenarios testing. NAbC provides control at the correlation cell level, with the ability to specify scenarios for many submatrices of the correlation matrix while holding the rest of the values constant (e.g. for flexible use in models like Black-Litterman (1991) and its many variants), by combining two crucial findings (I. and II.) below.



I. Pourahmadi and Wang (2015) and others show that both the marginal distributions of individual correlations in a correlation matrix, as well as the overall distribution of the entire correlation matrix, are invariant to the ordering of the rows and columns of the matrix. Respectively,

$$r_{i,j} \propto \left(1 - r_{i,j}^{2}\right)^{[k+p/2-1]}, \ i > j = Beta\left(k + \frac{p}{2}, k + \frac{p}{2}\right) \text{ distribution on } (-1,1), \text{ and}$$

$$f(r) = c_{p}\left(k\right)\left(\prod_{j=1}^{p}\prod_{l=1}^{j-1}\left[\sin\left(\theta_{j,l}\right)\right]^{2}\right)^{k} = c_{p}\left(k\right)\left[\det(R)\right]^{k}, \ j = 1, \dots, p-1, \ i > j \quad \text{ where } c_{p}\left(k\right) = \prod_{j=1}^{n-1}\left[\frac{\Gamma\left(\frac{2k+j}{2}+1\right)}{\sqrt{\pi}\Gamma\left(\frac{2k+j+1}{2}\right)}\right]^{j}$$

II. Additionally, focusing on the lower triangle, we observe that, based on $R = BB^{T}$ and equivalently,

$$r_{i,j} = \cos\left(\theta_{i,1}\right)\cos\left(\theta_{j,1}\right) + \prod_{k=2}^{i-1}\cos\left(\theta_{i,k}\right)\cos\left(\theta_{j,k}\right)\prod_{l=1}^{k-1}\sin\left(\theta_{i,l}\right)\sin\left(\theta_{j,l}\right) + \cos\left(\theta_{j,l}\right)\prod_{l=1}^{i-1}\sin\left(\theta_{i,l}\right)\sin\left(\theta_{j,l}\right) \text{ for } 1 \le i < j \le n$$

changing an arbitrary angle in *B* only will change any correlations that are to its right in the same row, and under the diagonal in the corresponding column.* Several examples are below.



* Note that some of the orange cells, and some of the rightmost triangle cells can remain unchanged if some of the corresponding correlation values are zero.



- Taking I. and II. together, we see that if we first reorder the correlation matrix so that only targeted cells are in the rightmost triangle of *B*, filling it in the numbered order of the matrix below, then no other cells in the matrix will be affected by changes in their values.
 Rightmost Triangle Fill Order
- This means that we can specify MANY combinations of cells in a matrix and reorder the rows and columns accordingly so that <u>only those correlations</u> <u>will be changed!</u> Note that there are only p! possible ways to sort the matrix, but [p(p-1)/2]! ways to order all p(p-1)/2 cells, so not all groupings/reorderings
 - into the lower rightmost triangle are possible. The further apart the cells (e.g. cell 1 vs cells 11, 12, and 7) the less likely it is that they can be grouped together into the lower right triangle; this ultimately depends on how many cells are being placed in the lower right triangle, as well as their positioning. For example, one rule governing this is as follows: if a cell is in the rightmost lower triangle position, to be continguous with it other cells must share at least one of its row or column indices; so cells 11, 12, 7, 13, 8, and 4 can never be contiguous with cell 1 if cell 1 is in the rightmost lower triangle position.
- However, even if one or two extraneous correlation cells must be placed with the 'chosen' ones for perturbation in the lower rightmost triangle, this is a vast improvement over no control, and still provides tremendous flexibility for scenario specification where before none existed.
- To take one example, reordering the correlation matrix so that rows 1-6 are now 6-1 and columns 1-6 are now 6-1, means that the original cells 1,2 and 1,3 and 2,3 and 4,3 are now in the rightmost triangle of the lower triangular matrix, in the fill order shown above, and changes to the corresponding cells in the angles matrix B will only change these same cells, after $R = BB^{T}$, in the resulting correlation matrix (see below).



Determine Targeted Change Cells



Reorder Rows/Cols to Fill Rightmost Triangle with Targets According to Fill Order

11					
12	7				
13	8	4			
14	9	5	2		
15	10	6	3	1	

Changes in Corresponding Angles Cells ONLY change Same in Resorted Matrix

11					
12	7				
13	8	4,3			
14	9	5	2,3		
15	10	6	1,3	1,2	

- Note that the targeted cells (green) do not have to be contiguous. Once they are re-ordered into the lower right triangle, <u>changes in these (orange) cells in the angles matrix will ONLY change these</u> <u>corresponding (orange) cells in the correlation matrix!</u>
- Changes to targeted cells WILL change other cells in the targeted submatrix (due to $R = BB^T$). HOWEVER, the ordering of the submatrix matters here, and can be exploited. For example, if we want to change the 4 cells above, but subsequently want to perform 'what if' analyses on only one of those cells (e.g. cell 1,2) without changing the other three, we reorder the original correlation matrix, if possible, to place that cell as the 'first' in the lower triangle of the B matrix, as shown. Then, subsequent changes to it will not affect the other (orange) cells. This 'rightmost' change rule is nested / hierarchical.
- This gives us dramatically increased control and flexibility in specifying scenarios: many
 combinations of correlation cells can be selected and defined as relevant to the scenario, while
 those unrelated to the scenario will remain completely 'untouched' by NAbC's perturbation, all
 while automatically preserving positive definiteness.



- Yet one question remains: when implementing NAbC for targeted scenarios, what do we do with the correlation cells that we do not want to change? To keep these cells 'untouched,' we must obtain the mean empirical correlation matrix from the simulations in Step 3., translate these to a matrix of angles, and then when sampling angles via nonparametric kernel in Step 4., simply insert these mean angles as constant values in each of the appropriate cells, across all N samples.
- This way, when we convert sampled angles matrices back to correlation matrices in step 5., the correlation values in the 'untouched' cells will be their expected values,* and remain 'untouched.'
- Using the mean angles in this way not only preserves the mean correlation values,* but also still enforces positive definiteness, since we used the ANGLES from the mean empirical correlation matrix, and not the mean empirical correlation matrix itself; thus, we remain on the unit hyperhemisphere, and remain positive definite.
- Note that other researchers have struggled with the issue of separating and isolating the effects of individual correlations without affecting others in the matrix. Ng et al. (2013) and Yu et al. (2014) identify this as changes in "peripheral" correlations due to changes in "core" correlations, and are not able to control it.
- However, the simple method described above reorder the correlation matrix as needed, and then
 only modify the corresponding angles in the lower triangular matrix while filling the 'untouched' cells
 with the mean-angle value constant allows for perfect control over which correlations are perturbed,
 and which remain completely untouched, thus solving the problem.
- The spectral distributions of these submatrix scenarios appropriately reflect their selectively limited, yet still positive definite, perturbation, as shown below in empirical results.

^{*} Strictly from the perspective of the sampling distribution, the empirical mean correlation matrix is the right source to obtain the constant-valued angles. Note, however, that the empirical mean matrix will not match the one specified in the model under non-elliptical data. This non-preservation under nonlinear transformations is discussed below.





- Empirical Results from NAbC Example Cases, Targeted Scenario versions (APPENDIX 2):
- A. TS: MVG Gaussian Identity Matrix: Targeted Scenario with only first 5 (of 10) cells targeted (n=6, n=126)
- F. TS: MVTVNS: Targeted Scenario with only first 5 (of 10) cells targeted(ρ = block structure, df=3,4,5,6,7, skewness=1,0.6,0,-0.6,-1, autocorrelation=-0.25,0,0.25,0.5,0.75,nonstationarity=3 σ , $\sigma/3$, σ , n/3 each) (n=6, n=126)1 -0.1 -0.1 0.2 0.2

	1	-0.1	-0.1	0.2	0.2
	-0.1	1	-0.1	0.2	0.2
Case F.TS: <i>R</i> =	-0.1	-0.1	1	0.2	0.2
	0.2	0.2	0.2	1	0.5
	0.2	0.2	0.2	0.5	1

- <u>Empirical Results from NAbC Summary (see APPENDIX 2 for Graphs/Tables)</u>:
- <u>Case A.TS</u>: shows the effects of imposing targeted scenarios wherein only specified cells (cells 1-5) are allowed to vary. Note that the tabular results show that the values of targeted cells are essentially identical to those of Case A., as they should be, while the non-targeted cells (cells 6-10) all remain untouched (see the spikes in these angles' distributions, which are a constant value): they are the cdf's/correlation values associated with the mean empirical correlation matrix (which will be the same as the specified one under elliptical data); yet all the simulated correlation matrices that result from converting the angles matrices to Cholesky factorizations to correlation matrices all are positive definite, as required, because we remain on the hyper hemisphere by sampling angles.
- <u>Case F.TS</u>: shows the same results as Case A.TS under more real world data conditions. The relative effects of imposing a targeted scenario compared to Case F., the unconstrained case, are the same as those when Case A.TS is compared to Case A.

- We have seen above that the angles associated with the cells of Pearson's product moment correlation matrix fully capture the dependence structure represented by its values; in other words, the distributions of the angles reproduce the distributions of the correlation cells. This is demonstrated with NAbC, both analytically (in the case of the Gaussian identity matrix) and empirically (under general conditions). This allows the estimation of p-values, confidence intervals, and quantile function (at both the matrix and cell levels, simultaneously), as well as flexible scenarios (perturbation of only selected cells), all by isolating the effects of dependence structure without altering any other aspect of the data generating mechanism.
- To again spell out why the **angles fully capture dependence structure** here:
- **First**, any correlation matrix (based on real/observed/actual data) must be positive definite.
- Second, the Cholesky factor, which is how we obtain the angles between data vectors and which only and always exists under positive definiteness, places us on the <u>unit</u> hyper-(hemi)sphere, where scale does not matter: as is widely known, $\rho(X,Y) = \rho(aX,bY)$; a,b > 0
- <u>Proposition</u>: given a correlation matrix that is positive definite and scale invariant (as would be expected of most, if not all types of widely used correlation), the angles associated with its values fully capture the dependence structure represented by its values – the relationship is direct & bi-directional.
- Both Spearman's Rho and Kendall's Tau are positive definite (see Sabado et al., 2007), as well as scale invariant (see Schreyer et al., 2017): $\rho_{SP}(X,Y) = \rho_{SP}(aX,bY)$; $\tau(X,Y) = \tau(aX,bY)$; a,b > 0
- Therefore, we propose that the distributions of the angles corresponding to each of their matrix cells fully represent the distributions of those cells, just as with Pearson's.
- This would enable the use of NAbC with Spearman's Rho and Kendall's Tau as well.



VII. Beyond Pearson's: Application to Spearman's Rho & Kendall's Tau

$$\rho_{SP} = \frac{\sum_{i=1}^{N} (R_{X_i} - E(R_X)) (R_{Y_i} - E(R_Y))}{\sqrt{\sum_{i=1}^{N} (R_{X_i} - E(R_X))^2} \sqrt{\sum_{i=1}^{N} (R_{Y_i} - E(R_Y))^2}}$$
$$\rho_{SP} = \frac{\sum_{i=1}^{n} (R_{X_i} - E(R_X)) (R_{Y_i} - E(R_Y))}{\sqrt{\sum_{i=1}^{n} (R_{X_i} - E(R_X))^2} \sqrt{\sum_{i=1}^{n} (R_{Y_i} - E(R_Y))^2}}$$

where R_X, R_Y are ranks of X and Y

$$\tau(X,Y) = \frac{2}{n(n-1)} \sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \operatorname{sgn}(x_i - x_j) \operatorname{sgn}(y_i - y_j)$$

where $\operatorname{sgn}(z) = 1$ if $z > 0$, $\operatorname{sgn}(z) = -1$ if $z < 0$,

 $\operatorname{sgn}(z) = 0$ if z = 0, for both N and n

A commonly used variant ("Tau-b") when tie values exist is:

$$\tau_{b}(X,Y) = \frac{\left[\sum_{i=1}^{n-1} \sum_{j=i+1}^{n} \operatorname{sgn}(x_{i} - x_{j}) \operatorname{sgn}(y_{i} - y_{j})\right]}{\sqrt{(n_{0} - n_{1})(n_{0} - n_{2})}}$$

where

$$n_0 = n(n-1)/2; n_1 = \sum_{i=1}^{i \text{ grps ties}} t_i(t_i - 1)/2; n_2 = \sum_{j=1}^{j \text{ grps ties}} u_j(u_j - 1)/2$$

 $t_i = \#$ ties in i-th group of x; $u_j = \#$ ties in j-th group of y

- <u>Proposition</u>: the distributions of the angles corresponding to each of the matrix cells, for both Spearman's Rho and Kendall's Tau, fully represent the distributions of those cells, just as with Pearson's (see Fernandez-Duran & Gregorio-Dominguez, 2023; Zhang & Songshan, 2023).
- We demonstrate this empirically below, using the same real-world data simulations as completed above with Pearson's correlation matrix (Case F.).



- Note that while the data generating mechanism for Cases F.KT and F.SP is a copula that utilizes a Pearson's correlation matrix (defined in Case F above), this is only for computational convenience to generate a real-world multivariate dataset whose marginal distributions exhibit different degrees of serial correlation, asymmetry, tail heaviness, and non-stationarity. Generating data directly using Kendall's and/or Spearman's typically is more cumbersome than using Pearson's, with a few notable exceptions (eg using Kendall's for multivariate t and Archimedian copulas: see McNeil et al, 2005; Embrechts et al., 2003; McNeil & Neslehova, 2009; and McNeil et al., 2022).
- However, the data generation mechanism is unrelated to the application of NAbC to Kendall's Tau
 and Spearman's Rho here, because the point of the exercise is to verify that the angles corresponding
 to Kendall's Tau and Spearman's Rho form the basis of their distributions; in other words, that their
 angles can be perturbed systematically to generate, via NAbC, the distributions of their correlation
 cells and hence, their matrices, just as with Pearson's.
- The empirical results below show identical distributions of eigenvalues, angles, capital (VaR), commonly used norms (only Frobenius/Euclidean is shown), and correlation cells when all are based directly on the data simulations, compared to when all are based on a kernel-perturbation of the angles via NAbC. These results do, in fact, mirror those of Pearson's, as proposed above, thus confirming empirically the assertions of Fernandez-Duran & Gregorio-Dominguez (2023), and Zhang & Songshan (2023).
- This speaks to a very wide range of application for NAbC to many additional measures of dependence / concordance beyond Pearson's, Spearman's, and Kendall's (eg distance correlation (see Szekely et al., 2007, and Sejdinovic et al., 2013), the Tail Dependence Matrix (see Embrechts et al, 2016, and Shyamalkumar & Tao, 2020), and others), because positive definiteness and scale invariance are unrestrictive and typical, if not required characteristics of most such measures.

- Empirical Results from NAbC Example Cases (APPENDIX 3):
- **F. KT and F.SP**: MVTVNS (*ρ* = block structure, df=3,4,5,6,7, skewness=1,0.6,0,-0.6,-1, autocorrelation= -0.25,0,0.25,0.5,0.75, nonstationarity=3σ, σ/3, σ, n/3 each) (**n=12**, n=126)

Cases F.KT and F.SP: *R* (*Pearson's*) =

1	-0.1	-0.1	0.2	0.2
-0.1	1	-0.1	0.2	0.2
-0.1	-0.1	1	0.2	0.2
0.2	0.2	0.2	1	0.5
0.2	0.2	0.2	0.5	1

- Empirical Results from NAbC Summary (see APPENDIX 3 for Graphs/Tables):
- <u>Cases F.KT and F.SP</u>: These results are based on a 'real world' example (Case F.) where marginal distributions vary by tail heaviness, skewness, autocorrelation, and degree of nonstationarity. As with Pearson's matrix in Case F., we again see identical results distributionally for 3. (direct data simulations) v 4. (kernel-based angle perturbation via NAbC) across all comparison criteria: eigenvalue distributions, correlation cell distributions, angle distributions, capital (VaR) distributions, and norm distributions (only Frobenius/Euclidean is shown, but Taxicab & Chebyshev also are equal).
- Note the increased precision in the angles distributions of Kendall's over Spearman's, and Spearman's over Pearson's (for the latter, Case F.PR is presented to include n = 12 for consistency).
- Also note that numerical considerations become relevant for small samples (eg matrix dimension p = n+1) when calculating Kendall's and Spearman's matrices: to avoid a non-trivial percentage of numerically non-positive definite results (out of N=10k simulations), when p=5, n must equal 12 rather than 6 (or p+1), as in Case F. for Pearson's. So for consistency, Case F.PR is presented for n = 12.



• Even as its long history (see Pearson, 1895) and widespread usage and accessibility inevitably has lead to its misuse and misunderstanding in some settings, Pearson's product moment correlation, as the scaled variance-covariance matrix, remains fundamental and foundational, and it's relevance and usage will remain ubiquitous. However, misuses notwithstanding, we must always consider, as with any method, its proper range of application and potential limitations under various conditions to assess whether it is fit-for-purpose for answering particular research questions in particular settings.

Linearity:

The widely held view that Pearson's product moment correlation is only appropriate for measuring linear relationships is called a myth in recent research in an American Statistical Association journal. Van den Heuval & Zhan (2022) show, analytically, a number of cases where Pearson's outperforms its rank-based counterparts (e.g. Spearman's Rho and Kendall's Tau) in terms of power for detecting nonlinear monotonic associations. Conversely, they also show cases where the opposite is true for linear relationships. They conclude that defining the conditions under which one measure of association is better than another is more complex than the literature would lead us to believe, and that rejecting Pearson's correlation a priori for assessing anything but linear relationships would be illadvised: "Pearson's correlation coefficient should not be ruled out a priori for measuring nonlinear monotonic associations....Our examples show that existing views on linear and monotonic associations are myths." Note also that even in methodological settings characterized as highly nonlinear, Pearson's often can play a central role (e.g. "Pearson Correlation Coefficient-based performance enhancement of Vanilla Neural Network for Stock Trend Prediction," Thakkar et al., 2021), as it often does in investment strategies (see Zhang et al., 2022) where it can dominate numerous other measures of association.



- <u>Non-preservation under nonlinear transformations</u>: Unlike its rank-based counterparts, Pearson's correlation is not preserved under nonlinear transformations. For example, it is preserved in elliptical copulas (e.g. Gaussian and student's t copulas, among others), but generally not non-elliptical copulas. While it is sometimes possible to recover approximations of the original correlation values based on the pairwise empirical correlation values (see Channouf and L'Ecuyer, 2012), or generalize the two-sample problem with copulas with many parameters or using vine copulas (see Barbiero, 2019), the latter is not straightforward and remains the topic of continuing research.
- However, in many situations, it may not matter. For example, if we specify a particular Pearson's correlation for use in a non-elliptical copula (say, the asymmetric student's t-copula with varying df of Church, 2012), and its <u>estimation</u> fits the data very well, then for practical usage, it may not matter that the actual Pearson's correlation embedded in the model is not that specified as an input parameter. Conversely, if we were using this model to <u>simulate</u> data, the Pearson's correlation matrix that we would estimate from the simulated data would be different from that specified as an input, but as long as it produced the required multivariate density, it would not adversely affect the analysis.
 - Still, lacking an analytical translation between true and empirical correlation matrices in the face of nonlinear transformations remains an important gap in the literature and a challenge for usage under many conditions. In the interim, to address this the SAS/IML code accompanying this monograph and to be posted on GitHub automatically generates two sets of results: one where the 'center' matrix in the sampling density is the original, specified correlation matrix, and one where this 'center' is the empirical mean correlation matrix. Under multivariate ellipticity, these two results are the same: deviations from ellipticity cause deviations in the two sets of results, and actually can be taken as a measure of departures from ellipticity. The empirical correlation matrix forms the basis of the sampling distribution under general conditions, and these are the results provided in this presentation.



- <u>Tail Risk</u>: For assessing risk in financial portfolios, much focus is rightly placed on tail densities / dependence / co-movement, and less on the co-movement of the means (i.e. Pearson's) of the marginal distributions. It turns out, however, that under many conditions Pearson's correlation plays a direct role here as well. For elliptically distributed data, Hult and Lindskog (2002) show that the tail dependence coefficients are fully determined by only two things: the tail index, and Pearson's correlation. Similarly, Lauria et al. (2021) analytically define some tail coefficients exclusively in terms of Pearson's correlation for some semi-heavy tailed distributions (e.g. elliptical Generalized Hyperbolic). Similar to an argument one might make for using Pearson's <u>alongside</u> its rank-based counterparts in light of Van den Heuval & Zhan's (2022) findings, these works support the use of Pearson's as an important component of portfolio risk analysis even when the focus is on an area, i.e tail risk, that many would (wrongly) assume to be unrelated to Pearson's.
- Any quantitative method has limitations and finite ranges of appropriate application. Cavalierly using the most readily available methods, based solely on convenience and not critical vetting, as has been the case when Pearson's is misused, is antithetical to applied, scientific research. On the other hand, conveniently sweeping and overly restrictive pronouncements for or against their use even if perhaps especially if they are based on commonly held beliefs can make oversights in applied research more obscured and insidious, if not necessarily more misleading and damaging in the long run.
- Pearson's correlation matrix remains, at the very least, a foundational baseline. Given the current state of the literature, for most serious portfolio analyses, failing to include a rigorous examination of its potential impacts, or those of its unscaled fraternal twin, the variance-covariance matrix, would be ill advised, even if the focus is on questions not obviously directly related to them.

Wide-ranging Application:

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Using a geometric framework, we have developed a method defining the **sampling densities of** widely used correlation/dependence matrices – **Pearson's, Kendall's, Spearman's, Szekely's (generalized) distance correlation, the Tail Dependence Matrix, and more – under the most general conditions possible**: simply, a positive definite matrix.

- The only requirements for the Nonparametric Angles-based Correlations (NAbC) method are i. a well estimated or known/specified correlation matrix and ii. a well estimated or known/specified data generating mechanism. Positive definiteness automatically is enforced on the unit hyper-hemisphere.
- NAbC is robustly applied herein under challenging, real-world data conditions characterized by varying degrees of tail heaviness, asymmetry, non-stationarity, and serial correlation in the marginal distributions, as well as under complex copula functions and near-singular matrices.
- NAbC appears to be valid for application to ANY measure of dependence/concordance characterized by positive definiteness and scale invariance (unrestrictive and typical, if not required qualities).
- <u>Confidence Intervals, p-values, and Inverse (quantile) Function Consistently and Simultaneously</u>
 <u>Determined at the Cell– and Matrix–levels</u>:

Each unique matrix in the sampling density is associated, one-to-one, with a unique matrix of cdf's of all the angle densities associated with all the correlation cells. Given any (positive definite) correlation matrix, NAbC can provide the matrix of cdf's. NAbC also provides the inverse 'quantile function': when given a matrix of cdf's, it provides the unique, associated correlation matrix.

The sampling density determines confidence intervals, as well as p-values for hypothesis tests, statistical process control, and as a probabilistic distance metric with some advantages over norm-based metrics. Cls & p-values of the cells and the entire matrix are determined simultaneously.



- <u>Flexible</u>: NAbC can be applied to submatrices of the given correlation matrix, while holding all or most of the 'untouched' cells constant, and still automatically enforce positive definiteness.
- This is required for flexible scenario specification, and not only can be used in views-based models like Black-Litterman (1991) and its variants, but also can augment and improve those views on the correlation matrix by quantifying them PROBABILISTICALLY.
- <u>Robust</u>: NAbC also is more robust than eigen-structure approaches, remaining stable even as matrices approach singularity (where many approximations (e.g. Fisher Z-transform) breakdown). And real-world data conditions (eg varying degrees of tail heaviness, asymmetry, non-stationarity, and serial correlation in the marginal distributions) pose no implementation issues for NAbC.
- **<u>Scalable</u>**: Finally, the algorithm is **scalable**, executing with non-prohibitive runtimes for non-small matrices (e.g. 100x100).

Correlation Breakdowns are widely documented, arguably endemic characteristics of major financial markets, and their destructive potential to our attempts to estimate and forecast market behavior is difficult to overestimate. Modeling efforts in this area simply cannot be effective in any way (ie accurate, precise, or robust) without knowledge of, and the ability to implement and utilize, the true sampling distributions of the metrics measuring dependence structure under real world conditions. Yet to date, the range of application, and arguably effectiveness, of relevant methods in the extant literature remains limited.

What's more, proactively flagging, probabilistically monitoring, and potentially mitigating and avoiding correlation breakdowns requires all the characteristics of NAbC described above, which collectively are shared by no other approach. NAbC thus fills a crucially important hole in the literature for robust inference, scenario analytics, and stress testing of dependence structure in all aspects of financial portfolio analysis, including risk assessment, portfolio construction, efficient allocation, and all related forecasting efforts.

- <u>"All-cases" Analytic Angles Distribution</u>: Analogous to deriving the angles distribution (pdf, cdf, and analytic quantile function) for Pearson's under the Gaussian identity matrix presumption, as is done herein, deriving the "all cases" <u>analytic</u> solution, under general conditions, for Pearson's, Kendall's, and Spearman's matrices, would add great insight to this applied research and speed up implementation over our nonparametric (NAbC) solution. This open problem appears to be nontrivial, if as similarly challenging as deriving the same for the spectral distributions of these matrices (see Bandeira et al., 2017, and Li, Wang, and Li, 2021; and Li, Wang, and Wang, 2021 for more on the spectral distributions associated with Kendall's matrix; as well as Lindskog et al., 2003 for more on at least modestly surprising connections between Pearson's and Kendall's).
- <u>Comparisons Against / Conjoint Usage with, Competing Methods</u>: Comparing NAbC, under wideranging test conditions, to Hansen & Archakov (2021) and the Bayesian approaches of Lan et al. (2020) and Ghosh et al. (2020) likely would be a useful and insightful exercise. The four approaches differ in notable ways, yet have some strong similarities: the two Bayesian approaches both adopt the hyperspherical geometric framework that forms the foundation for NAbC. Perhaps using them in conjunction would be fruitful, as the limitations of each cited above may be the strengths of the other. This certainly appears to be the case for the evolutionary approach of Papenbrock et al. (2021).
- <u>Statistical Process Monitoring</u>: While NAbC's application to hypothesis testing is self-evident, it would be useful to see if the many statistical process control methods designed for monitoring correlation and covariance matrices could make use of the NAbC sampling density, or if the latter could be used to validate results of the methods proposed in papers like Adegoke et al. (2022), Ajadi et al. (2021), Bours & Steland (2020), Wang et al. (2019), Choi and Shin (2021), and others like those reviewed in Ebadi et al. (2021).



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• Empirical Results from NAbC – Notes (see APPENDIX 1 for Graphs/Tables):

- NOTE: All results show a comparison of Step 3. vs. Step 4. as validation of the latter EXCEPT Case A using the C3 quantile function (because no data is being simulated the method is analytical).
- Where noted results are for p=100x100, but otherwise, for illustrative purposes all are for p=5x5 (sample size n=p+1, and n=126 for 6 months of trading days; for p=100, n=p+1 and n=252 for a year of trading days). Kernel bandwidth $h = 0.9 \cdot \min(IQR/1.34, \hat{\sigma}) \cdot N^{-1/5}$ version of Silverman (1986) x 0.15. Results include distributions of the angles, eigenvalues, Euclidian/Frobenius norm, LNP, square-root-rule elliptical aggregated 'capital,'* and the differences between all the cells in the two empirical mean correlation matrices. A scatterplot and Pearson's correlation between LNP and the Euclidian/Frobenius norm also is presented. All tables of matrices are presented in lower triangle column format. The matrices include conversions from i. specified correlation matrices to cdf matrices**; ii. specified cdf matrices to unique correlation matrices; and iii. specified %CDF-shifts from the empirical mean matrix to correlation matrices (wherein each cell's %shift is the percentage of the cumulative density above or below the cdf of the mean angle). Finally, angle number is determined based on the following fill-order (the motivation for this is discussed in a later section):

Rightmost Triangle Fill Order

11					
12	7				
13	8	4			
14	9	5	2		
15	10	6	3	1	

** Note that the inverse relationship between angles cdf's and correlation values applies here: for the %CDF-shift tables, this has been reversed so that positive/negative CDF shifts correspond with increases/decreases in correlation values.

* Under multivariate ellipticity, extreme portfolio quantiles (aka Value-at-Risk, "VaR", aka "Capital") can be calculated asymptotically with the 'square root rule,' where portfolio VaR=Capital = $\sqrt{VRV^t}$ where V = vector of p VaRs (quantiles) and R = correlation matrix of dimension p (see Tao et al., 2019, and Frachot et al., 2001). While this does not hold under other distributions, its distribution under Step 3. still should be identical to that of Step 4., so they are presented for comparison purposes. All 99VaR's are set to \$50m herein: when different marginals are used, they are scaled so that 99VaR = \$50m.

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- Empirical Results from NAbC Example Cases:
- A. C3 v MVG Gaussian Identity Matrix: side-by-side comparison of use of fully analytic C3 quantile function (no data generation) vs. a multivariate gaussian data generating mechanism (n=6, n=126)
- B. MVG with all ρ = 0.2 (n=6, n=126)
- C. MVG with $\rho = 0.4$ for one cell (#5) v. $\rho = 0.4$ for one factor (row 4, cells 2,5,9) (n=6)
- D. MVT (df=3) with all ρ = 0.2 (n=6, n=126)
- E. MVTM MVT (df=3) with different marginals (2 LN(0,1), 3 Gamma(1,1)), all ρ = 0.2 (n=6, n=126)
- F. MVTVNS (ρ = block structure, df=3,4,5,6,7, skewness=1,0.6,0,-0.6,-1, autocorrelation= -0.25,0,0.25,0.5,0.75, nonstationarity=3 σ , σ /3, σ , n/3 each) (n=6, n=126) **1** -0.1
- G. MVTVNS (all *ρ* = 0.2, **p=100x100**, df=3 to 27.5 by 0.5, skewness=1 to -0.96 by 0.04, autocorrelation=0.68 to -0.3 by 0.02, nonstationarity=3σ, σ/3, σ, n/3 each) (n=101, n=252)

26)	1	-0.1	-0.1	0.2	0.2		
-	-0.1	1	-0.1	0.2	0.2		
Case F.: <i>R</i> =	-0.1	-0.1	1	0.2	0.2		
	0.2	0.2	0.2	1	0.5		
	0.2	0.2	0.2	0.5	1		

- H. Gumbel Copula (ϕ = 1.25), as upper tail dependence is relevant in this setting (n=6, n=126)
- NOTE: Cases of n ≤ p are not within the scope of this work, as NAbC requires positive definite, and thus, full rank matrices.



• <u>Empirical Results from NAbC – Summary</u>:

- All results are distributionally identical for Steps 3. v 4., indicating validity in NAbC's methodology.
- All spectral distributions for both 3. and 4. match those defined in the RMT literature (e.g. Marchenko-Pastur distribution (Marchenko & Pastur, 1967) under the identity matrix, and otherwise where spectral distributions are derived analytically (e.g. tests not included herein were performed using Lillo & Mantegna, 2005, and Livan et al., 2011 as examples). Also the effects of heavy-tailed distributions and serial correlation appear consistent with those of the RMT literature generally (see for the former Burda et al., 2004, Burda et al., 2006, Akemann et al., 2009; Abul-Magd et al., 2009, Bouchaud & Potters, 2015, Martin & Mahoney, 2018; and for the latter, see Burda et al., 2004, 2011).
- The relationship between LNP and the Euclidian/Frobenius norm is consistent with the latter's inability to distinguish between the RELATIVE distances from the estimated/given correlation matrix: they are very similar under the Gaussian identity matrix (especially under larger sample sizes), where relative and absolute distances are more similar, but when the original correlation matrix is asymmetric or contains more extreme values, the strength of their association diminishes somewhat, as expected.
- Note that the scale of the LNP distribution is not dependent on sample size, whereas that of the Euclidean/Frobenius and other norms do change with sample size. The scales of both are functions of the dimension of the matrix.
- On tests not presented herein, LNP remains numerically robust even under p=100x100, and readily provides rankings of 'distance' for all specified correlation and cdf matrices.



• For completeness, we include the pdf of the Marchenko-Pastur (1967) distribution below:

 $f(x) = \frac{1}{2\pi\sigma^2} \frac{\sqrt{(\lambda_+ - x)(\lambda_- - x)}}{\lambda x} \text{ for } \lambda_- \le x \le \lambda_+, \text{ and zero otherwise, when } 0 \le \lambda \le 1,$ and when $\lambda > 1$ we have an additional mass point of $\left(1 - \frac{1}{\lambda}\right)$ at x = 0

where $\lambda_{\pm} = \sigma^2 (1 \pm \sqrt{\lambda}), \lambda = \frac{p}{n}, p = \# \text{factors/rows/columns}, n = \text{sample size},$ and $\sigma = 1$ for correlation matrices

- This analytic result overlays the empirical spectral distributions presented in the appendices.
- Note that exceptions to convergence to this celebrated distribution do exist (see Li and Yao (2018), Hisakado and Kaneko (2023), and Maltsev and Malysheva (2024) for examples).



• <u>Empirical Results from NAbC – Summary</u>:

- <u>Case A</u>: shows distributionally identical results when comparing those generated from 3. vs. those based on 4. Also, NAbC applied to multivariate gaussian data yields distributionally identical results vs. probability inverse transform sampling using the analytic C3 quantile function derived herein, thus validating the latter empirically.
- <u>Case B</u>: shows identical results distributionally for 3. v 4., with the expected effects for sample size and for non-zero correlations on both the spectral and angles distributions.
- <u>Case C</u>: demonstrates how eigen-structure/decomposition approaches are too blunt a tool for analyzing correlation matrices. We start with the identity matrix and examine the spectral distributions of two cases: one where an entire factor/row (cells 2, 5, and 9) is assigned $\rho = 0.4$ vs. one where only a single correlation cell (#5) gets $\rho = 0.4$. Under small sample sizes (n=p+1) they look virtually the same. The angles distributions, however, show the story very clearly, for each of the right cells, and the untouched zero cells show exact matches with both the defined pdf (sin^k) and the C3 analytic derivation. As sample size increases, as expected, the spectral distributions can more readily be distinguished, but still do not provide the unambiguous story provided by the angles distributions. Note also how cell #1 (in row 5) is (unintentionally) affected by the changes of the factor corresponding to row 4 due to the rightmost cell-change rule described on pp. 47-49. NAbC is the first method able to explicitly control that kind of unintended correlation 'contamination.'
- <u>Case D</u>: shows the incremental effects of heavy tails, all else equal, compared to Case B. NAbC still matches empirical 'truth' perfectly, by every distributional criteria.
- <u>Case E</u>: shows the incremental effects of different marginals, all else equal, compared to Case D. NAbC still matches empirical 'truth' perfectly, by every distributional criteria.



Empirical Results from NAbC – Summary:

- <u>Case F</u>: shows a more 'real world' example (with a block correlation matrix) including heavy tails (varying by factor), skewness (varying by factor), autocorrelation (varying by factor), and nonstationarity. We see, again, identical results distributionally for 3. v 4. across all comparison criteria. Examination of both the spectral and angles distributions in this case shows complex structure, providing further evidence that deriving an 'all-cases' analytic density for either appears to be nontrivial.
- <u>Case G</u>: shows Case F. but for a non-toy matrix of p=100x100 and a constant correlation matrix with $\rho = 0.2$. Again, NAbC still matches empirical 'truth' perfectly, by every distributional criteria.
- <u>Case H</u>: shows a copula with upper tail dependence (Gumbel) and we see, again, identical results distributionally for 3. v 4. across all comparison criteria.
- These extensive results confirm empirically the assertions of Fernandez-Duran & Gregorio-Dominguez (2023), and Zhang & Songshan (2023): that the angles distributions preserve the entire dependence structure represented by the correlation/dependence measures themselves, thus making them the key to a single, unified approach to all inference and scenario/stress testing of such measures (as well as COMPARISONS across such measures).



• <u>Empirical Results from NAbC – Code</u>:

- NOTE: Complete SAS/IML (v9.4) code that generates the above results and many possible specifications for a wide range of input parameters, including input .csv files (e.g. the specified cdf matrices, correlation matrices, %CDF-shift matrices, the baseline correlation matrices, eigenvalues for defining baseline correlation matrices, etc.) will be made publicly available for download on GitHub and in the forthcoming book, "Beating the Correlation Breakdown: Robust Inference and Flexible Stress Testing and Scenarios for Financial Portfolios," Elements in Quantitative Finance series, Cambridge University Press, eds. Ricardo Rebonato, PhD.
- The SAS/IML code has extensive functionality. In addition to the optionality mentioned above, correlation matrices can be specified either as correlation matrices, or defined by eigenvalues (which are converted via Givens rotations a la Davies and Higham, 2000), matrix re-orderings for targeted submatrices can be specified as an input parameter (either as a list or in a .csv file), scales of axes on graphs and kernel bandwidths on the kernel-based graphs are input parameters, the number of angles graphs generated is a parameter, etc.





Elliptical Capital Sample Size n = 6

MVG Data





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RunID = Ains5x5MVG







RunID = Ainb5x5MVG











C3 (Analytic)

MVG Data





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Euclidian/Frobenius Norm Sample Size n = 126











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Angle Distributions Sample Size n = 6





C3 (Analytic)

Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 6











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Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 6







Angle Distributions Sample Size n = 126







Angle Distributions Sample Size n = 126







Angle Distributions Sample Size n = 126







Angle Distributions Sample Size n = 126







Angle Distributions Sample Size n = 126







Angle Distributions Sample Size n = 126






Angle Distributions Sample Size n = 126







Angle Distributions Sample Size n = 126













Angle Distributions Sample Size n = 126





Correlation Matrices to CDF Matrices (n=6)

Correlation Matrices

CDFs: C3 (Analytic)

CDFs: MVG Data

1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.	50 0.43	0.35	0.57	0.57	0.5	0.43	0.35	0.58	
0	0.1	0.2	-0.1	-0.1	0.	50 0.43	0.35	0.57	0.57	0.5	0.42	0.35	0.57	
0	0.1	0.2	-0.1	-0.1	0.	50 0.43	0.35	0.57	0.57	0.4	9 0.42	0.35	0.57	
0	0.1	0.2	-0.1	-0.1	0.	50 0.43	0.35	0.57	0.57	0.5	0.42	0.35	0.58	
1	1	1	1	1										
0	0.1	0.2	-0.1	0.2	0.	50 0.44	0.39	0.57	0.38	0.5	0 0.44	0.40	0.57	
0	0.1	0.2	-0.1	0.2	0.	50 0.44	0.39	0.57	0.38	0.4	9 0.44	0.39	0.57	
0	0.1	0.2	-0.1	0.2	0.	50 0.44	0.39	0.57	0.38	0.4	9 0.44	0.39	0.57	
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.	50 0.46	0.43	0.56	0.58	0.5	0.46	0.43	0.56	
0	0.1	0.2	-0.1	-0.1	0.	50 0.46	0.43	0.56	0.58	0.5	0.46	0.43	0.56	
1	1	1	1	1										
0	0.1	0.2	-0.1	0.3	0.	50 0.48	0.46	0.55	0.42	0.5	0.48	0.47	0.55	
1	1	1	1	1										



0.58 0.57 0.57 0.58

0.38 0.37 0.37

0.58 0.57

0.43

Correlation Matrices to CDF Matrices (n=126)

Correlation Matrices

CDFs: C3 (Analytic)

CDFs: MVG Data

1		

1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.50	0.13	0.01	0.87	0.87	0.51	0.14	0.01	0.87	0.87
0	0.1	0.2	-0.1	-0.1	0.50	0.13	0.01	0.87	0.87	0.50	0.13	0.01	0.86	0.86
0	0.1	0.2	-0.1	-0.1	0.50	0.13	0.01	0.87	0.87	0.51	0.13	0.01	0.87	0.87
0	0.1	0.2	-0.1	-0.1	0.50	0.13	0.01	0.87	0.87	0.50	0.13	0.01	0.87	0.87
1	1	1	1	1										
0	0.1	0.2	-0.1	0.2	0.50	0.16	0.03	0.89	0.02	0.50	0.16	0.03	0.90	0.02
0	0.1	0.2	-0.1	0.2	0.50	0.16	0.03	0.89	0.02	0.50	0.15	0.03	0.89	0.01
0	0.1	0.2	-0.1	0.2	0.50	0.16	0.03	0.89	0.02	0.49	0.16	0.03	0.89	0.02
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.50	0.18	0.06	0.92	0.96	0.50	0.18	0.06	0.91	0.95
0	0.1	0.2	-0.1	-0.1	0.50	0.18	0.06	0.92	0.96	0.51	0.19	0.06	0.92	0.96
1	1	1	1	1										
0	0.1	0.2	-0.1	0.3	0.50	0.20	0.08	0.94	0.00	0.50	0.19	0.08	0.94	0.00
1	1	1	1	1										



CDF Matrices to Correlation Matrices (n=6)

CDF Matrices

Correlations: C3 (Analytic)



					1	. 1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.00	0.13	-0.13	0.20	-0.61	0.00	0.13	-0.13	0.20	-0.61
0.5	0.4	0.6	0.35	0.9	0.00	0.13	-0.13	0.20	-0.61	0.00	0.13	-0.14	0.19	-0.60
0.5	0.4	0.6	0.35	0.9	0.00	0.13	-0.13	0.20	-0.61	-0.01	0.12	-0.14	0.20	-0.61
0.5	0.4	0.6	0.35	0.9	0.00	0.13	-0.13	0.20	-0.61	0.00	0.14	-0.13	0.21	-0.62
					1	. 1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.6	0.8	0.00	0.17	-0.14	-0.11	0.06	0.00	0.18	-0.14	-0.11	0.05
0.5	0.4	0.6	0.6	0.8	0.00	0.17	-0.14	-0.11	0.06	-0.01	0.16	-0.13	-0.11	0.06
0.5	0.4	0.6	0.6	0.8	0.00	0.17	-0.14	-0.11	0.06	-0.01	0.16	-0.14	-0.12	0.07
					1	. 1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.00	0.23	-0.15	0.25	0.71	0.00	0.24	-0.14	0.26	0.72
0.5	0.4	0.6	0.4	0.3	0.00	0.23	-0.15	0.25	0.71	0.00	0.23	-0.16	0.25	0.72
					1	. 1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.00	0.36	-0.20	-0.04	0.20	0.02	0.38	-0.17	-0.02	0.23
					1	. 1	1	1	1	1	1	1	1	1



CDF Matrices to Correlation Matrices (n=126)

CDF Matrices

Correlations: C3 (Analytic)



					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.00	0.02	-0.02	0.03	-0.11	0.00	0.02	-0.02	0.04	-0.11
0.5	0.4	0.6	0.35	0.9	0.00	0.02	-0.02	0.03	-0.11	0.00	0.02	-0.02	0.03	-0.12
0.5	0.4	0.6	0.35	0.9	0.00	0.02	-0.02	0.03	-0.11	0.00	0.02	-0.02	0.04	-0.11
0.5	0.4	0.6	0.35	0.9	0.00	0.02	-0.02	0.03	-0.11	0.00	0.02	-0.02	0.03	-0.11
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.6	0.8	0.00	0.02	-0.02	-0.02	-0.06	0.00	0.02	-0.02	-0.02	-0.06
0.5	0.4	0.6	0.6	0.8	0.00	0.02	-0.02	-0.02	-0.06	0.00	0.02	-0.02	-0.02	-0.06
0.5	0.4	0.6	0.6	0.8	0.00	0.02	-0.02	-0.02	-0.06	0.00	0.02	-0.02	-0.02	-0.06
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.00	0.02	-0.02	0.02	0.07	0.00	0.02	-0.02	0.03	0.06
0.5	0.4	0.6	0.4	0.3	0.00	0.02	-0.02	0.02	0.07	0.00	0.03	-0.02	0.03	0.07
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.00	0.02	-0.02	-0.01	-0.12	0.00	0.02	-0.02	-0.01	-0.13
					1	1	1	1	1	1	1	1	1	1



CDF %Shift Matrices to Correlation Matrices (n=6)

CDF %Shift Matrices

Correlations: C3 (Analytic)



Image						InMatPr	-2.231	-2.231	-6.931	-6.931	-12.799	-2.22	-2.241	-6.893	-6.970	-12.755
Image Image <t< td=""><td></td><td></td><td></td><td></td><td></td><td>FNorm</td><td>0.899</td><td>0.654</td><td>2.308</td><td>1.160</td><td>3.040</td><td>0.90</td><td>0.638</td><td>2.322</td><td>1.160</td><td>3.045</td></t<>						FNorm	0.899	0.654	2.308	1.160	3.040	0.90	0.638	2.322	1.160	3.045
Image Image Image Rnk_FNor 2 1 4 3 5 1 1 1 1 Image Image <						Rnk_InMat	1	1	3	3	5		1 2	3	4	5
Image						Rnk_FNorr	2	1	4	3	5		2 1	4	3	5
1 < 0 $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ $1 < 0$ <																
20 -20 50 -50 60 0.13 -0.13 0.35 -0.35 0.43 0.13 -0.13 0.35 -0.35 0.43 20 -20 50 -50 60 0.13 -0.13 0.35 -0.35 0.43 0.13 -0.13 0.05 -0.35 0.43 20 -20 50 -50 60 0.13 -0.13 0.35 -0.35 0.43 0.12 -0.14 0.35 -0.35 0.43 20 -20 50 -50 60 0.13 -0.13 0.35 -0.35 0.43 0.14 0.14 0.13 0.35 -0.35 20 -20 50 -50 60 0.13 -0.14 0.35 -0.35 0.43 0.44 0.14 0.13 0.55 -0.35 0.66 20 -20 50 -50 70 0.17 -0.14 0.48 -0.23 0.66 0.18 0.17 0.48 0.14 20 -20 50 -50 70 0.17 -0.14 0.48 -0.23 0.66 0.17 0.14 0.48 -0.24 0.17 0.14 0.48 0.14 0.16 0.17 0.14 <t< td=""><td></td><td></td><td></td><td></td><td></td><td></td><td>1</td><td>1</td><td>1</td><td>1</td><td>1</td><td></td><td>1 1</td><td>1</td><td>1</td><td>1</td></t<>							1	1	1	1	1		1 1	1	1	1
20-2050-50600.13-0.130.03-0.030.0430.13-0.140.035-0.0420-2050-50600.13-0.130.05-0.050.040.140.150.050.0420-2050-50600.13-0.130.05-0.050.040.14-0.140.05-0.050.0420-2050-50600.11	20	-20	50	-50	60		0.13	-0.13	0.35	-0.35	0.43	0.1	-0.13	0.35	-0.35	0.43
120 -20 50 -50 -50 60 0.13 -0.13 0.03 -0.35 0.43 0.12 -0.14 0.35 -0.35 0.43 120 -20 50 -50 -50 0.0 0.01 -0.14 0.01 -0.14 0.01 0.0	20	-20	50	-50	60		0.13	-0.13	0.35	-0.35	0.43	0.1	-0.14	0.35	-0.34	0.42
$ \begin{array}{c c c c c c c c c c c c c c c c c c c $	20	-20	50	-50	60		0.13	-0.13	0.35	-0.35	0.43	0.1	-0.14	0.35	-0.35	0.43
1 < 1 < 1 < 1 < 1 < 1 < 1 < 1 < 1 < 1 <	20	-20	50	-50	60		0.13	-0.13	0.35	-0.35	0.43	0.1	4 -0.13	0.35	-0.35	0.42
20-2050-50700.17-0.140.48-0.230.660.180.18-0.130.48-0.240.4720-20505050700.17-0.140.48-0.230.660.170.170.140.480.470.470.470.480.47 </td <td></td> <td></td> <td></td> <td></td> <td></td> <td></td> <td>1</td> <td>1</td> <td>1</td> <td>1</td> <td>1</td> <td></td> <td>1 1</td> <td>1</td> <td>1</td> <td>1</td>							1	1	1	1	1		1 1	1	1	1
20-2050-50700.17-0.140.48-0.230.660.17-0.130.48-0.240.4720-2050-5070-111 </td <td>20</td> <td>-20</td> <td>50</td> <td>-50</td> <td>70</td> <td></td> <td>0.17</td> <td>-0.14</td> <td>0.48</td> <td>-0.23</td> <td>0.66</td> <td>0.1</td> <td>-0.13</td> <td>0.48</td> <td>-0.24</td> <td>0.67</td>	20	-20	50	-50	70		0.17	-0.14	0.48	-0.23	0.66	0.1	-0.13	0.48	-0.24	0.67
$\begin{array}{c c c c c c c c c c c c c c c c c c c $	20	-20	50	-50	70		0.17	-0.14	0.48	-0.23	0.66	0.1	7 -0.13	0.48	-0.24	0.66
$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$	20	-20	50	-50	70		0.17	-0.14	0.48	-0.23	0.66	0.1	7 -0.14	0.47	-0.24	0.65
$\begin{array}{c c c c c c c c c c c c c c c c c c c $							1	1	1	1	1		1 1	1	1	1
20 -20 50 -50 80 0.23 -0.15 0.63 -0.09 0.89 0.24 -0.15 0.64 -0.10 0.99 100	20	-20	50	-50	80		0.23	-0.15	0.63	-0.10	0.89	0.2	4 -0.14	0.64	-0.10	0.90
Image: Normal state Image: Normal state<	20	-20	50	-50	80		0.23	-0.15	0.63	-0.10	0.89	0.2	4 -0.15	0.64	-0.10	0.90
20 -20 50 -50 90 0.36 -0.20 0.84 0.06 1.00 0.37 -0.18 0.84 0.07 1.00 1<							1	1	1	1	1		1 1	1	1	1
	20	-20	50	-50	90		0.36	-0.20	0.84	0.06	1.00	0.3	7 -0.18	0.84	0.07	1.00
							1	1	1	1	1		1 1	1	1	1



CDF %Shift Matrices to Correlation Matrices (n=126)

CDF %Shift Matrices

Correlations: C3 (Analytic)



					InMatPr	-2.231	-2.231	-6.931	-6.931	-12.799	-2.233	-2.230	-6.938	-6.925	-12.881
					FNorm	0.105	0.100	0.288	0.255	0.479	0.103	0.099	0.286	0.255	0.480
					Rnk_InMat	1	1	3	3	5	2	1	4	3	5
					Rnk_FNorr	2	1	4	3	5	2	1	4	3	5
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	60		0.02	-0.02	0.06	-0.06	0.08	0.02	-0.02	0.06	-0.06	0.08
20	-20	50	-50	60		0.02	-0.02	0.06	-0.06	0.08	0.02	-0.02	0.06	-0.06	0.08
20	-20	50	-50	60		0.02	-0.02	0.06	-0.06	0.08	0.02	-0.02	0.06	-0.06	0.08
20	-20	50	-50	60		0.02	-0.02	0.06	-0.06	0.08	0.02	-0.02	0.06	-0.06	0.08
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	70		0.02	-0.02	0.06	-0.06	0.10	0.02	-0.02	0.06	-0.06	0.10
20	-20	50	-50	70		0.02	-0.02	0.06	-0.06	0.10	0.02	-0.02	0.06	-0.06	0.10
20	-20	50	-50	70		0.02	-0.02	0.06	-0.06	0.10	0.02	-0.02	0.06	-0.06	0.10
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	80		0.02	-0.02	0.07	-0.05	0.13	0.02	-0.02	0.07	-0.05	0.13
20	-20	50	-50	80		0.02	-0.02	0.07	-0.05	0.13	0.03	-0.02	0.07	-0.05	0.13
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	90		0.02	-0.02	0.07	-0.05	0.17	0.02	-0.02	0.07	-0.05	0.17
						1	1	1	1	1	1	1	1	1	1



n = 6



n = 126









RunID = Bns5x5MVG



RunID = Bnb5x5MVG



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PRNG Seed = 12345 RunID = Bns5x5MVG

-25

-20

-15

Distribution of Log(Pr Matrix|Mean Matrix)

Angle-Sampled
 Data Generated

-10

-5

0.025

0.000

PRNG Seed = 12345 RunID = Bnb5x5MVG

-25

-20

-15

Distribution of Log(Pr Matrix|Mean Matrix)

- Angle-Sampled ----- Data Generated

-10

-5

0.025

0.000



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n = 6



n = 126





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RunID = Bnb5x5MVG



-25

Distribution of Frobenius Norm



n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





n = 6



n = 126





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n = 6



n = 126





n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





n = 6



n = 126





	Correlation Matrices to CDF Matrices	
Correlation Matrices	n=6	n=126

1	1	1	1	1										
0	0.3	0.4	0.1	0.1	0.67	0.44	0.35	0.60	0.60	0.99	0.12	0.01	0.88	0.88
0	0.3	0.4	0.1	0.1	0.67	0.44	0.35	0.59	0.59	0.99	0.12	0.01	0.87	0.87
0	0.3	0.4	0.1	0.1	0.66	0.44	0.35	0.59	0.59	0.99	0.12	0.01	0.88	0.88
0	0.3	0.4	0.1	0.1	0.67	0.44	0.35	0.60	0.60	0.99	0.12	0.01	0.88	0.88
1	1	1	1	1										
0	0.3	0.4	0.1	0.4	0.63	0.48	0.44	0.57	0.36	0.97	0.23	0.08	0.81	0.00
0	0.3	0.4	0.1	0.4	0.62	0.48	0.44	0.56	0.36	0.97	0.23	0.08	0.81	0.00
0	0.3	0.4	0.1	0.4	0.62	0.47	0.43	0.56	0.36	0.97	0.23	0.08	0.80	0.00
1	1	1	1	1										
0	0.3	0.4	0.1	0.1	0.59	0.50	0.48	0.55	0.63	0.94	0.31	0.18	0.75	0.99
0	0.3	0.4	0.1	0.1	0.58	0.49	0.48	0.55	0.62	0.95	0.31	0.19	0.75	0.99
1	1	1	1	1										
0	0.3	0.4	0.1	0.3	0.57	0.52	0.51	0.54	0.52	0.92	0.35	0.26	0.69	0.35
1	1	1	1	1										



CDF

ľ	Matrico	es		CDF Matri	ces to	Correl n=6	ation I	Matric	es	I	n=126	
T					1 1	1	1	1	1	1	1	
ŀ	0.6	0.35	0.9	0.2	.3 0.35	0.10	0.41	-0.44	0.20	0.22	0.18	(
ŀ	0.6	0.35	0.9	0.2	.2 0.35	0.09	0.41	-0.44	0.20	0.22	0.18	(
ŀ	0.6	0.35	0.9	0.2	0.34	0.08	0.40	-0.44	0.20	0.22	0.18	(
ŀ	0.6	0.35	0.9	0.2	2 0.35	0.09	0.41	-0.45	0.20	0.22	0.18	(

						1 1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.	.3 0.35	0.10	0.41	-0.44	0.2	0 0.22	0.18	0.24	0.09
0.5	0.4	0.6	0.35	0.9	0.	.2 0.35	0.09	0.41	-0.44	0.2	0 0.22	0.18	0.23	0.09
0.5	0.4	0.6	0.35	0.9	0.	0.34	0.08	0.40	-0.44	0.2	0 0.22	0.18	0.24	0.09
0.5	0.4	0.6	0.35	0.9	0.	.2 0.35	0.09	0.41	-0.45	0.2	0 0.22	0.18	0.23	0.09
						1 1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.6	0.8	0.	.4 0.42	0.05	0.20	-0.07	0.2	0 0.23	0.17	0.19	0.10
0.5	0.4	0.6	0.6	0.8	0.	.42	0.04	0.19	-0.07	0.2	0 0.23	0.17	0.19	0.10
0.5	0.4	0.6	0.6	0.8	0.	0.41	0.04	0.19	-0.07	0.2	0 0.23	0.17	0.19	0.10
						1 1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.	.5 0.51	-0.01	0.48	0.67	0.2	0 0.23	0.17	0.23	0.20
0.5	0.4	0.6	0.4	0.3	0.	.4 0.51	-0.02	0.47	0.67	0.2	0 0.24	0.17	0.23	0.20
						1 1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.	0.65	-0.08	0.32	0.01	0.2	0 0.24	0.16	0.20	0.03
						1 1	1	1	1		1 1	1	1	1



					CDF 9	%Shift	Matrio	ces to	Corre	lation	Matri	ces				
	CDF ^o	%Shif	t Matri	ces				n=6					n	=126		
						\rightarrow										
					InMatPr	-2.373	-2.092	-7.511	-6.385	-14.597		-2.261	-2.202	-7.050	-6.814	-13.154
					FNorm	1.011	0.876	2.199	1.748	2.684		0.128	0.125	0.347	0.334	0.548
					Rnk_InMat	2	1	4	3	5		2	1	4	3	5
					Rnk_FNorr	2	1	4	3	5		2	1	4	3	5
						1	1		1	1		1	1	1	1	1
20	20	50	50	<u> </u>		1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	60		0.31	0.06	0.51	-0.16	0.58		0.22	0.18	0.26	0.14	0.27
20	-20	50	-50	60		0.31	0.06	0.51	-0.16	0.58		0.22	0.18	0.26	0.14	0.27
20	-20	50	-50	60)	0.31	0.05	0.51	-0.16	0.58		0.22	0.18	0.26	0.14	0.27
20	-20	50	-50	60		0.31	0.05	0.51	-0.17	0.57		0.22	0.18	0.26	0.14	0.27
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	70		0.38	0.01	0.66	-0.22	0.80		0.23	0.17	0.28	0.13	0.31
20	-20	50	-50	70		0.38	0.01	0.66	-0.22	0.79		0.23	0.17	0.28	0.13	0.31
20	-20	50	-50	70		0.38	0.00	0.65	-0.23	0.79		0.23	0.17	0.28	0.12	0.31
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	80		0.47	-0.05	0.80	-0.25	0.95		0.23	0.17	0.29	0.11	0.35
20	-20	50	-50	80)	0.46	-0.06	0.80	-0.25	0.95		0.23	0.17	0.29	0.11	0.36
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	90		0.59	-0.16	0.93	-0.26	1.00		0.24	0.16	0.31	0.10	0.41
						1	1	1	1	1		1	1	1	1	1









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RunID = CnsO5x5MVG



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	Correlation Matrices to CDF Matrices													
	Corre	lation	Matric	es:		one f	actor	= 0.4		(n = 6)	one	e cell =	0.4	
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.50	0.43	0.35	0.58	0.58	0.50	0.43	0.35	0.58	0.58
0	0.1	0.2	-0.1	-0.1	0.50	0.42	0.35	0.57	0.57	0.50	0.42	0.35	0.57	0.57
0	0.1	0.2	-0.1	-0.1	0.81	0.76	0.69	0.86	0.86	0.50	0.42	0.35	0.57	0.57
0	0.1	0.2	-0.1	-0.1	0.50	0.43	0.35	0.58	0.58	0.50	0.42	0.35	0.58	0.58
1	1	1	1	1										
0	0.1	0.2	-0.1	0.2	0.50	0.44	0.40	0.57	0.38	0.50	0.44	0.40	0.57	0.38
0	0.1	0.2	-0.1	0.2	0.80	0.76	0.73	0.84	0.71	0.78	0.74	0.70	0.83	0.68
0	0.1	0.2	-0.1	0.2	0.49	0.44	0.39	0.57	0.37	0.49	0.44	0.39	0.57	0.37
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.80	0.77	0.74	0.84	0.84	0.50	0.46	0.43	0.56	0.58
0	0.1	0.2	-0.1	-0.1	0.50	0.46	0.43	0.56	0.57	0.50	0.46	0.43	0.56	0.57
1	1	1	1	1										
0	0.1	0.2	-0.1	0.3	0.79	0.77	0.75	0.82	0.73	0.50	0.48	0.47	0.55	0.43
1	1	1	1	1										



CDF Watrices to Correlation Matrices														
	CDF	Matric	es			one f	actor	= 0.4		(n = 6)	one	cell =	0.4	
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.00	0.13	-0.13	0.20	-0.61	0.00	0.13	-0.13	0.20	-0.61
0.5	0.4	0.6	0.35	0.9	0.00	0.13	-0.14	0.19	-0.60	0.00	0.13	-0.14	0.19	-0.60
0.5	0.4	0.6	0.35	0.9	0.44	0.54	0.32	0.59	-0.22	0.00	0.13	-0.14	0.20	-0.61
0.5	0.4	0.6	0.35	0.9	0.00	0.14	-0.13	0.20	-0.62	0.00	0.14	-0.13	0.21	-0.62
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.6	0.8	0.00	0.18	-0.14	-0.11	0.05	0.00	0.18	-0.14	-0.11	0.05
0.5	0.4	0.6	0.6	0.8	0.45	0.58	0.31	0.41	0.14	0.46	0.58	0.33	0.35	0.34
0.5	0.4	0.6	0.6	0.8	-0.01	0.16	-0.14	-0.12	0.06	-0.01	0.16	-0.14	-0.12	0.07
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.45	0.61	0.27	0.57	0.67	0.00	0.28	-0.21	0.18	0.61
0.5	0.4	0.6	0.4	0.3	0.00	0.24	-0.15	0.26	0.72	0.00	0.23	-0.16	0.25	0.72
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.46	0.63	0.27	0.51	0.01	0.01	0.38	-0.24	-0.09	0.04
					1	1	1	1	1	1	1	1	1	1



CDF %Shift Matrices to Correlation Matrices

one factor = 0.4

(n = 6) one cell = 0.4

					InMatPr	-2.434	-2.040	-7.821	-6.220	-14.443	-2.259	-2.205	-7.051	-6.830	-13.15
					FNorm	0.806	0.698	1.905	1.512	2.478	0.936	0.698	2.273	1.229	2.93
					Rnk_InMat	2	1	4	3	5	2	1	4	3	
					Rnk_FNorr	2	1	4	3	5	2	1	4	3	
						1	1	1	1	1	1	1	1	1	
20	-20	50	-50	60		0.13	-0.13	0.35	-0.35	0.43	0.13	-0.13	0.35	-0.35	0.4
20	-20	50	-50	60		0.13	-0.14	0.35	-0.34	0.42	0.13	-0.14	0.35	-0.34	0.4
20	-20	50	-50	60		0.49	0.25	0.65	0.05	0.70	0.13	-0.14	0.34	-0.35	0.4
20	-20	50	-50	60		0.13	-0.13	0.34	-0.35	0.42	0.14	-0.13	0.35	-0.35	0.4
						1	1	1	1	1	1	1	1	1	
20	-20	50	-50	70		0.18	-0.13	0.48	-0.24	0.67	0.18	-0.13	0.48	-0.24	0.6
20	-20	50	-50	70		0.53	0.23	0.73	0.01	0.82	0.52	0.26	0.73	0.12	0.8
20	-20	50	-50	70		0.17	-0.14	0.46	-0.24	0.65	0.17	-0.14	0.47	-0.24	0.6
						1	1	1	1	1	1	1	1	1	
20	-20	50	-50	80		0.58	0.19	0.80	-0.04	0.89	0.28	-0.20	0.67	-0.28	0.8
20	-20	50	-50	80		0.24	-0.15	0.64	-0.09	0.89	0.24	-0.15	0.64	-0.10	0.9
						1	1	1	1	1	1	1	1	1	
20	-20	50	-50	90)	0.62	0.10	0.82	-0.16	0.89	0.38	-0.24	0.82	-0.10	0.9
						1	1	1	1	1	1	1	1	1	



n = 6



n = 126





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RunID = Dns5x5MVT



RunID = Dnb5x5MVT







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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





n = 6



n = 126





n = 6



n = 126





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n = 6



n = 126





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Correlation Matrices to CDF Matrices											
Correlation Matrices	n=6	n=126									

1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.64	0.58	0.52	0.69	0.69	0.86	0.72	0.51	0.94	0.94
0	0.1	0.2	-0.1	-0.1	0.64	0.58	0.52	0.69	0.69	0.87	0.72	0.50	0.94	0.94
0	0.1	0.2	-0.1	-0.1	0.64	0.58	0.52	0.69	0.69	0.87	0.72	0.51	0.94	0.94
0	0.1	0.2	-0.1	-0.1	0.64	0.58	0.52	0.69	0.69	0.87	0.72	0.51	0.94	0.94
1	1	1	1	1										
0	0.1	0.2	-0.1	0.2	0.61	0.56	0.52	0.66	0.51	0.83	0.68	0.51	0.94	0.45
0	0.1	0.2	-0.1	0.2	0.61	0.57	0.53	0.67	0.51	0.83	0.68	0.51	0.93	0.45
0	0.1	0.2	-0.1	0.2	0.61	0.56	0.52	0.66	0.51	0.83	0.68	0.50	0.93	0.44
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.58	0.55	0.52	0.64	0.65	0.80	0.65	0.50	0.93	0.95
0	0.1	0.2	-0.1	-0.1	0.59	0.55	0.52	0.64	0.65	0.81	0.65	0.51	0.94	0.95
1	1	1	1	1										
0	0.1	0.2	-0.1	0.3	0.56	0.54	0.52	0.60	0.48	0.78	0.62	0.50	0.94	0.22
1	1	1	1	1										



	Natric	es		CDF Matric	r	n=126							
				1	. 1	1	1	1	1	1	1		
0.4	0.6	0.35	0.9	0.23	0.38	0.06	0.46	-0.57	0.20	0.25	0.16		
0.4	0.6	0.35	0.9	0.24	0.40	0.06	0.47	-0.56	0.20	0.24	0.16		
0.4	0.6	0 35	0.9	0.23	0.38	0.07	0 46	-0 57	0.21	0.25	0.16		

					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.23	0.38	0.06	0.46	-0.57	0.20	0.25	0.16	0.27	-0.04
0.5	0.4	0.6	0.35	0.9	0.24	0.40	0.06	0.47	-0.56	0.20	0.24	0.16	0.27	-0.04
0.5	0.4	0.6	0.35	0.9	0.23	0.38	0.07	0.46	-0.57	0.21	0.25	0.16	0.27	-0.03
0.5	0.4	0.6	0.35	0.9	0.23	0.38	0.06	0.45	-0.58	0.20	0.25	0.16	0.27	-0.03
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.6	0.8	0.25	0.48	0.02	0.22	0.04	0.21	0.26	0.15	0.19	0.02
0.5	0.4	0.6	0.6	0.8	0.25	0.47	0.03	0.23	0.05	0.20	0.26	0.15	0.19	0.03
0.5	0.4	0.6	0.6	0.8	0.25	0.47	0.02	0.22	0.05	0.20	0.26	0.15	0.19	0.02
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.27	0.57	-0.03	0.53	0.76	0.20	0.27	0.14	0.25	0.23
0.5	0.4	0.6	0.4	0.3	0.27	0.56	-0.03	0.53	0.77	0.21	0.27	0.14	0.26	0.23
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.29	0.69	-0.14	0.36	0.27	0.20	0.28	0.13	0.21	-0.10
					1	1	1	1	1	1	1	1	1	1



CDF %Shift Matrices to Correlation Matrices																		
	CDF %Shift Matrices						n=6						n=126					
					InMatPr	-2.395	-2.071	-7.602	-6.304	-14.832		-2.297	-2.167	-7.196	-6.675	-13.439		
					FNorm	1.175	1.010	2.478	1.800	2.910		0.253	0.237	0.679	0.625	1.058		
					Rnk_InMat	2	1	4	3	5		2	1	4	3	5		
					Rnk_FNorr	2	1	4	3	5		2	1	4	3	5		
						1	1	1	1	1		1	1	1	1	1		
20	-20	50	-50	60		0.34	0.01	0.58	-0.26	0.65		0.24	0.15	0.31	0.08	0.34		
20	-20	50	-50	60		0.35	0.01	0.59	-0.25	0.66		0.24	0.15	0.31	0.08	0.34		
20	-20	50	-50	60		0.34	0.02	0.58	-0.25	0.65		0.24	0.15	0.31	0.08	0.34		
20	-20	50	-50	60		0.34	0.01	0.56	-0.26	0.65		0.24	0.16	0.31	0.08	0.34		
						1	1	1	1	1		1	1	1	1	1		
20	-20	50	-50	70		0.42	-0.03	0.73	-0.24	0.86		0.25	0.14	0.35	0.06	0.42		
20	-20	50	-50	70		0.42	-0.02	0.73	-0.24	0.86		0.25	0.14	0.35	0.06	0.42		
20	-20	50	-50	70		0.41	-0.03	0.72	-0.24	0.85		0.25	0.14	0.35	0.05	0.42		
						1	1	1	1	1		1	1	1	1	1		
20	-20	50	-50	80		0.50	-0.08	0.85	-0.19	0.97		0.26	0.13	0.37	0.04	0.49		
20	-20	50	-50	80		0.49	-0.09	0.85	-0.21	0.97		0.26	0.14	0.38	0.04	0.50		
						1	1	1	1	1		1	1	1	1	1		
20	-20	50	-50	90		0.62	-0.20	0.95	-0.12	1.00		0.27	0.13	0.40	0.02	0.59		
						1	1	1	1	1		1	1	1	1	1		

0/ Chiff Matriaga to Correlation M



n = 6



n = 126









RunID = Es5x5MVTM



RunID = Eb5x5MVTM



n = 6



LNP

n = 126



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n = 6

n = 126





















n = 6



n = 126





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Angle Distributions



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Angle Distributions



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n = 6



n = 126





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n = 6



Angle Distributions

n = 126



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n = 6







n = 6



n = 126





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n = 6



n = 126





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n = 6







n = 6



n = 126





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n = 6



Angle Distributions

n = 126



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	Corre	elation	Matri	ces to	CDF I	Matric	es				
Correlation Matrices	n=6										
		\Rightarrow									

1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.59	0.53	0.47	0.65	0.65	0.926	0.753	0.528	0.995	0.995
0	0.1	0.2	-0.1	-0.1	0.60	0.54	0.48	0.66	0.66	0.940	0.781	0.531	0.995	0.995
0	0.1	0.2	-0.1	-0.1	0.60	0.54	0.48	0.66	0.66	0.940	0.773	0.536	0.994	0.994
0	0.1	0.2	-0.1	-0.1	0.60	0.54	0.48	0.66	0.66	0.939	0.774	0.537	0.996	0.996
1	1	1	1	1										
0	0.1	0.2	-0.1	0.2	0.58	0.53	0.50	0.64	0.48	0.882	0.702	0.515	0.983	0.448
0	0.1	0.2	-0.1	0.2	0.59	0.54	0.50	0.65	0.48	0.886	0.706	0.512	0.984	0.445
0	0.1	0.2	-0.1	0.2	0.58	0.53	0.49	0.64	0.48	0.881	0.702	0.513	0.981	0.447
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.58	0.54	0.51	0.63	0.64	0.875	0.690	0.511	0.989	0.994
0	0.1	0.2	-0.1	-0.1	0.57	0.54	0.51	0.63	0.64	0.881	0.694	0.522	0.987	0.995
1	1	1	1	1										
0	0.1	0.2	-0.1	0.3	0.55	0.53	0.52	0.59	0.48	0.838	0.654	0.510	0.987	0.176
1	1	1	1	1										



CDF Matrices to Correlation Matrices CDF Matrices n=6 n														
					\rightarrow									
					1	1	1	1	1		1	1	1	
0.4	0.6	0.35	0.9		0.143	0.322	-0.023	0.413	-0.476		0.21	0.26	0.17	

					1	1	1	1	1		1 1	. 1	1	1
0.5	0.4	0.6	0.35	0.9	0.143	0.322	-0.023	0.413	-0.476	0.	0.26	0.17	0.29	0.02
0.5	0.4	0.6	0.35	0.9	0.167	0.357	0.001	0.439	-0.481	0.	0.25	0.17	0.28	0.03
0.5	0.4	0.6	0.35	0.9	0.166	0.331	-0.002	0.420	-0.483	0.	0.26	0.17	0.28	0.03
0.5	0.4	0.6	0.35	0.9	0.162	0.332	-0.006	0.416	-0.483	0.	0.26	0.17	0.28	0.03
					1	1	1	1	1		1 1	. 1	1	1
0.5	0.4	0.6	0.6	0.8	0.176	0.422	-0.031	0.156	-0.073	0.	0.26	0.16	0.20	0.05
0.5	0.4	0.6	0.6	0.8	0.184	0.422	-0.021	0.156	-0.073	0.	0.26	0.16	0.20	0.05
0.5	0.4	0.6	0.6	0.8	0.169	0.414	-0.035	0.142	-0.065	0.	0.26	0.16	0.20	0.05
					1	1	1	1	1		1 1	. 1	1	1
0.5	0.4	0.6	0.4	0.3	0.204	0.511	-0.055	0.486	0.715	0.	0.27	0.16	0.26	0.22
0.5	0.4	0.6	0.4	0.3	0.202	0.513	-0.060	0.488	0.725	0.	0.27 0.27	0.16	0.26	0.22
					1	1	1	1	1		1 1	. 1	1	1
0.5	0.4	0.6	0.55	0.95	0.247	0.660	-0.162	0.297	0.129	0.	0.28	0.15	0.22	-0.02
					1	1	1	1	1		1 1	. 1	1	1



	CDF %Shift Matrices to Correlation Matrices														
	CDF	%Shift	Matri	ces				n=6				r	n=126		
					InMatPr	-2.251	-2.213	-7.009	-6.857	-13.451	-2.12	0 -2.345	-6.493	-7.395	-11.780
					FNorm	1.187	1.020	2.581	1.765	3.018	0.23	0.234	0.631	0.581	0.963
					Rnk_InMat	2	1	4	3	5		1 2	3	4	5
					Rnk_FNorr	2	1	4	3	5		1 2	4	3	5
						1	1	1	1	1		1 1	1	1	1
20	-20	50	-50	60		0.3391	-0.0058	0.6174	-0.2381	0.7001	0.2	9 0.19	0.37	0.11	0.41
20	-20	50	-50	60		0.3618	0.0062	0.6119	-0.2357	0.6939	0.2	5 0.18	0.33	0.12	0.36
20	-20	50	-50	60		0.3313	-0.0007	0.6003	-0.2301	0.6825	0.2	5 0.18	0.33	0.11	0.36
20	-20	50	-50	60		0.3350	-0.0048	0.5921	-0.2456	0.6747	0.2	6 0.18	0.33	0.12	0.36
						1	1	1	1	1		1 1	1	1	1
20	-20	50	-50	70		0.4210	-0.0356	0.7581	-0.2325	0.8854	0.2	8 0.17	0.37	0.09	0.44
20	-20	50	-50	70		0.4189	-0.0311	0.7609	-0.2329	0.8859	0.2	8 0.17	0.37	0.09	0.44
20	-20	50	-50	70		0.4131	-0.0385	0.7476	-0.2218	0.8791	0.2	8 0.17	0.37	0.09	0.44
						1	1	1	1	1		1 1	1	1	1
20	-20	50	-50	80		0.4969	-0.0757	0.8594	-0.1942	0.9754	0.2	8 0.16	0.38	0.08	0.48
20	-20	50	-50	80		0.4903	-0.0825	0.8642	-0.2021	0.9764	0.2	8 0.17	0.38	0.08	0.48
						1	1	1	1	1		1 1	1	1	1
20	-20	50	-50	90		0.6249	-0.2052	0.9549	-0.1897	0.9996	0.2	9 0.16	0.40	0.06	0.55
						1	1	1	1	1		1 1	1	1	1



n = 6



n = 126





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RunID = Fb5MMMblk



LNP

n = 6







n = 126

RunID = Fb5MMMblk



n = 6



n = 126





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Angles (epan-SILVR09-15)--Log(Pr(M)|Mean Matrix) v FNorm [r=-0.740]: DGM=MVTVNS, Mat=5x5, #Obs=126, #Sims = 10000



RunID = Fb5MMMblk

RunID = Fs5MMMblk



n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6



n = 126





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n = 6









n = 6



n = 126





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n = 6







n = 6



n = 126





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n = 6







n = 6



n = 126





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n = 6







Correlation Mat	rices	n=6										
			\Rightarrow									
1 1	1	1										

1	1	1	1	1										
0.1	0.2	0.3	0.4	0.3	0.53	0.49	0.45	0.40	0.45	0.4376	0.2705	0.1422	0.0672	0.1422
0.1	0.2	0.3	0.4	0.3	0.39	0.34	0.30	0.26	0.30	0.1493	0.0518	0.0116	0.0014	0.0116
0.1	0.2	0.3	0.4	0.3	0.38	0.35	0.31	0.27	0.31	0.2696	0.1437	0.0601	0.0177	0.0601
0.1	0.2	0.3	0.4	0.3	0.36	0.32	0.29	0.25	0.29	0.2125	0.099	0.0353	0.0077	0.0353
1	1	1	1	1										
0.1	0.2	0.3	0.4	0.4	0.46	0.44	0.41	0.39	0.37	0.206	0.0938	0.042	0.0184	0.0068
0.1	0.2	0.3	0.4	0.4	0.55	0.52	0.50	0.48	0.46	0.5734	0.4045	0.2671	0.169	0.0922
0.1	0.2	0.3	0.4	0.4	0.52	0.50	0.47	0.46	0.44	0.4607	0.2918	0.1656	0.0941	0.0426
1	1	1	1	1										
0.1	0.2	0.3	0.4	0.1	0.67	0.65	0.64	0.63	0.72	0.9824	0.953	0.9189	0.8742	0.9996
0.1	0.2	0.3	0.4	0.1	0.63	0.61	0.60	0.59	0.69	0.9535	0.8954	0.8262	0.7569	0.9996
1	1	1	1	1										
0.1	0.2	0.3	0.4	0.2	0.68	0.67	0.67	0.66	0.70	0.9978	0.9939	0.9902	0.986	0.9995
1	1	1	1	1										



1

1

1

0.56

0.49

0.81

1

0.75

0.70

0.94

1

1

0.5

0.5

0.5

0.5

0.5

0.5

0.5

0.5

0.5

0.5

0.4

0.4

0.4

0.6

0.6

0.6

0.4

0.4

0.55

0.3

0.3

0.95

CDF I	Matric	es		CDF Matrice	es to (orreia n=6	n=6 n=126									
• • • •																
				1	1	1	1	1	1	1	1	1				
0.4	0.6	0.35	0.9	0.18	0.40	-0.07	0.52	-0.80	0.07	0.12	0.01	0.15				
0.4	0.6	0.35	0.9	-0.17	0.07	-0.40	0.19	-0.87	-0.09	-0.05	-0.14	-0.02				
0.4	0.6	0.35	0.9	-0.18	0.06	-0.41	0.20	-0.88	-0.05	0.02	-0.11	0.05				
0.4	0.6	0.35	0.9	-0.25	-0.01	-0.47	0.12	-0.89	-0.09	-0.02	-0.14	0.01				
				1	1	1	1	1	1	1	1	1				
0.4	0.6	0.6	0.8	-0.04	0.27	-0.24	-0.15	0.48	-0.05	-0.01	-0.08	-0.08				
0.4	0.6	0.6	0.8	0.19	0.47	-0.01	0.06	0.53	0.12	0.17	0.08	0.08				
0.4	0.6	0.6	0.8	0.11	0.40	-0.09	-0.05	0.53	0.07	0.11	0.03	0.03				

1

1

1

0.43

0.37

0.58

1

1

1

0.72

0.69

0.81

1

1

1

0.98

0.98

0.93

1

0.37

0.31

0.55

1

1

1

1

1

0.40

0.34

0.59

1

1

1

0.35

0.29

0.52



1

-0.20

-0.32

-0.41

-0.43

-0.10

0.06

0.02

0.47

0.50

1 0.51

1

1

1

1

1

0.39

0.34

0.55

1

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		CDF %Shift Matrices to Correlation Matrices														
	CDF	%Shift	t Matri	ces				n=6					r	=126		
						\rightarrow										
					InMatPr	-2.369	-2.101	-7.529	-6.449	-14.780		-2.292	-2.175	-7.204	-6.731	-12.865
					FNorm	1.297	0.935	2.820	1.632	3.266		0.218	0.190	0.593	0.500	0.847
					Rnk_InMat	2	1	4	3	5		2	1	4	3	5
					Rnk_FNorr	2	1	4	3	5		2	1	4	3	5
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	60		0.364	-0.129	0.670	-0.494	0.749		0.13	0.02	0.22	-0.07	0.25
20	-20	50	-50	60		0.118	-0.355	0.459	-0.644	0.572		-0.05	-0.14	0.03	-0.22	0.06
20	-20	50	-50	60		0.129	-0.357	0.502	-0.656	0.616		0.00	-0.13	0.10	-0.24	0.14
20	-20	50	-50	60		0.081	-0.390	0.458	-0.685	0.594		-0.04	-0.16	0.07	-0.27	0.10
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	70		0.283	-0.234	0.723	-0.119	0.864		-0.01	-0.09	0.07	-0.14	0.13
20	-20	50	-50	70		0.396	-0.125	0.804	-0.052	0.913		0.14	0.05	0.24	-0.01	0.31
20	-20	50	-50	70		0.349	-0.177	0.766	-0.048	0.900		0.09	0.00	0.18	-0.04	0.25
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	80		0.655	0.236	0.934	0.498	0.987		0.40	0.35	0.46	0.33	0.54
20	-20	50	-50	80		0.617	0.215	0.922	0.522	0.988		0.35	0.30	0.41	0.28	0.48
						1	1	1	1	1		1	1	1	1	1
20	-20	50	-50	90		0.880	0.307	0.992	0.551	0.999		0.60	0.53	0.67	0.49	0.77
						1	1	1	1	1		1	1	1	1	1





n = 252

Data v Angles (epan-SILVR09-15) - Elliptical Risk Agg / Capital: DGM=MVTVNS, Mat=100x100, #Obs=252, #Sims = 10000











RunID = Gs5MXMp2



RunID = Gb5MXMp2



LNP

n = 101



0.003 0.002 0.001 0.000 -6000 -5500 -5000 -4500 Distribution of Log(Pr Matrix|Mean Matrix) Angle-Sampled ----- Data Generated PRNG Seed = 12345

RunID = Gb5MXMp2





n = 101

n = 252





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RunID = Gs5MXMp2





n = 101







Angle Distributions


n = 101



n = 252





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Angle Distributions



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n = 101



n = 252







Angle Distributions





Angle Distributions



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Angle Distributions



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n = 101





n = 252













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Angle Distributions



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Angle Distributions



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n = 6



n = 126





500

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RunID = Hns5x5GUM



RunID = Hnb5x5GUM





n = 126

Data v Angles (epan-SILVR09-15) - Log(Pr Matrix|Mean Matrix): DGM=gumcop, Mat=5x5, #Obs=126, #Sims = 10000







RunID = Hnb5x5GUM





n = 6

n = 126





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RunID = Hnb5x5GUM



n = 6

RunID = Hns5x5GUM



-10

n = 6



n = 126





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n = 6



Angle Distributions

n = 126



n = 6



Angle Distributions

n = 126



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n = 6



Angle Distributions

n = 126





Angle Distributions

n = 126



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n = 6



Angle Distributions

n = 126



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n = 6



Angle Distributions

n = 126



n = 6



n = 126





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n = 6



n = 126





n = 6



n = 126





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n = 6



n = 126





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	Correlation Matrices to CDF Matrices										
Correlation Matrices	n=6										
		\Rightarrow									
1 1 1	1										

1	1	1	1	1										
0	0.1	0.2	0.55	0.6	0.69	0.64	0.58	0.35	0.32	0.9980	0.9693	0.8607	0.1698	0.1173
0	0.1	0.2	0.55	0.6	0.70	0.63	0.57	0.35	0.32	0.9979	0.9706	0.8614	0.1766	0.1199
0	0.1	0.2	0.55	0.6	0.70	0.63	0.57	0.35	0.32	0.9985	0.9728	0.8620	0.1744	0.1194
0	0.1	0.2	0.55	0.6	0.69	0.63	0.56	0.35	0.32	0.9988	0.9715	0.8625	0.1735	0.1206
1	1	1	1	1										
0	0.1	0.2	0.55	0.3	0.63	0.57	0.53	0.42	0.68	0.9897	0.9207	0.7681	0.2606	0.9997
0	0.1	0.2	0.55	0.3	0.63	0.57	0.53	0.42	0.68	0.9887	0.9177	0.7701	0.2656	0.9998
0	0.1	0.2	0.55	0.3	0.62	0.57	0.53	0.42	0.67	0.9895	0.9184	0.7621	0.2647	0.9997
1	1	1	1	1										
0	0.1	0.2	0.55	0.6	0.60	0.56	0.53	0.47	0.42	0.9721	0.8485	0.6819	0.3200	0.1070
0	0.1	0.2	0.55	0.6	0.59	0.55	0.52	0.47	0.41	0.9699	0.8450	0.6800	0.3138	0.1069
1	1	1	1	1										
0	0.1	0.2	0.55	0.4	0.57	0.54	0.53	0.50	0.59	0.9447	0.7838	0.6334	0.3550	0.9959
1	1	1	1	1										



0.5

0.5

0.5

0.5

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0.5

0.5

0.5

0.5

0.4

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0.4

0.4

0.4

0.6

0.6

0.6

0.6

0.6

0.6

0.6

0.4

0.4

0.55

0.8

0.8

0.3

0.3

0.95

0.29

0.29

0.32

0.30

0.35

1

1

1

0.53

0.52

0.61

0.60

0.73

1

1

1

				CDF Matrice	es to (Correl	ation I	Matric	es			
CDF Matrices							n=126					
				1	1	1	1	1	1	1	1	1
0.4	0.6	0.35	0.9	0.32	0.48	0.16	0.55	-0.36	0.37	0.42	0.33	0.44
0.4	0.6	0.35	0.9	0.32	0.47	0.16	0.55	-0.36	0.37	0.42	0.33	0.44
0.4	0.6	0.35	0.9	0.32	0.48	0.15	0.55	-0.36	0.37	0.42	0.33	0.44
0.4	0.6	0.35	0.9	0.31	0.47	0.15	0.55	-0.36	0.37	0.42	0.33	0.44
				1	1	1	1	1	1	1	1	1
0.4	0.6	0.6	0.8	0.29	0.52	0.07	0.34	-0.14	0.36	0.42	0.31	0.38

0.07

0.06

0.03

0.00

-0.09

1

1

1

0.34

0.33

0.59

0.58

0.45

1

1

1

-0.14

-0.14

0.68

0.68

-0.01

1

1

1



1

0.38

0.38

0.42

0.41

0.39

1

1

1

0.42

0.42

0.42

0.42

0.43

1

1

1

0.31

0.31

0.30

0.30

0.29

1

1

1

0.37

0.36

0.36

0.36

0.35

1

1

1

1

0.17

0.17

0.18

0.17

0.18

0.18

0.18

0.31

0.31

0.12

1

1

1

1

0/Chiff Matriaga to Correlatio

					CDF 7	/oSnint	watri	ces to	Corre	ation	watrices					
CDF %Shift Matrices								n=6			n=126					
						\rightarrow										
					InMatPr	-2.279	-2.185	-7.124	-6.748	-13.519	-2.008	-2.460	-6.069	-7.882	-10.321	
					FNorm	1.103	1.096	2.303	2.172	2.661	0.234	0.263	0.649	0.649	0.950	
					Rnk_InMat	2	1	4	3	5	1	2	3	4	5	
					Rnk_FNorr	2	1	4	3	5	1	2	3	4	5	
						1	1	1	1	1	1	1	1	1	1	
20	-20	50	-50	60		0.46	0.13	0.69	-0.10	0.76	0.43	0.34	0.51	0.27	0.54	
20	-20	50	-50	60		0.45	0.13	0.69	-0.10	0.76	0.43	0.34	0.51	0.27	0.54	
20	-20	50	-50	60		0.45	0.13	0.68	-0.11	0.76	0.43	0.34	0.51	0.26	0.54	
20	-20	50	-50	60		0.45	0.12	0.68	-0.11	0.76	0.43	0.34	0.51	0.27	0.54	
						1	1	1	1	1	1	1	1	1	1	
20	-20	50	-50	70		0.51	0.07	0.81	-0.20	0.91	0.44	0.33	0.53	0.24	0.59	
20	-20	50	-50	70		0.52	0.07	0.81	-0.19	0.91	0.44	0.33	0.53	0.24	0.59	
20	-20	50	-50	70		0.51	0.05	0.81	-0.21	0.91	0.44	0.33	0.53	0.24	0.59	
						1	1	1	1	1	1	1	1	1	1	
20	-20	50	-50	80		0.59	0.01	0.90	-0.26	0.98	0.45	0.32	0.55	0.22	0.64	
20	-20	50	-50	80		0.59	0.00	0.90	-0.27	0.98	0.45	0.32	0.55	0.22	0.64	
						1	1	1	1	1	1	1	1	1	1	
20	-20	50	-50	90		0.70	-0.14	0.97	-0.34	1.00	0.45	0.31	0.57	0.20	0.70	
						1	1	1	1	1	1	1	1	1	1	



XII. APPENDIX 2: Empirical Results of NAbC, Targeted Scenario Matrix

- Empirical Results from NAbC Example Cases, Targeted Scenario versions:
- A. TS: MVG Gaussian Identity Matrix: Targeted Scenario with only first 5 (of 10) cells targeted (n=6, n=126)
- F. TS: MVTVNS: Targeted Scenario with only first 5 (of 10) cells targeted(ρ = block structure, df=3,4,5,6,7, skewness=1,0.6,0,-0.6,-1, autocorrelation=-0.25,0,0.25,0.5,0.75,nonstationarity=3 σ , $\sigma/3$, σ , n/3 each) (n=6, n=126)1 -0.1 -0.1 0.2 0.2

	1	-0.1	-0.1	0.2	0.2
	-0.1	1	-0.1	0.2	0.2
Case F.TS: <i>R</i> =	-0.1	-0.1	1	0.2	0.2
	0.2	0.2	0.2	1	0.5
	0.2	0.2	0.2	0.5	1

- <u>Empirical Results from NAbC Summary</u>:
- <u>Case A.TS</u>: shows the effects of imposing targeted scenarios wherein only specified cells (cells 1-5) are allowed to vary. Note that the tabular results show that the values of targeted cells are essentially identical to those of Case A., as they should be, while the non-targeted cells (cells 6-10) all remain untouched (see the spikes in these angles' distributions, which are a constant value): they are the cdf's/correlation values associated with the mean empirical correlation matrix (which will be the same as the specified one under elliptical data); yet all the simulated correlation matrices that result from converting the angles matrices to Cholesky factorizations to correlation matrices all are positive definite, as required, because we remain on the hyper hemisphere by sampling angles.
- <u>Case F.TS</u>: shows the same results as Case A.TS under more real world data conditions. The relative effects of imposing a targeted scenario compared to Case F., the unconstrained case, are the same as those when Case A.TS is compared to Case A.







RunID = ATib5x5MVG



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Difference Between Mean Empirical Matrices – Both Targeted Scenario n = 6 n = 126



RunID = ATis5x5MVG






XII. APPENDIX 2: Empirical Results of NAbC, Targeted – Case A.TS







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Euclidian/Frobenius Norm – Targeted Scenario v Unconstrained n = 6 n = 126





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Correlation Matrices to CDF Matrices – Both Targeted Scenario Correlation Matrices n=6 n=126

1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.50	0.50	0.50	0.50	0.50	0.503	0.503	0.503	0.503	0.503
0	0.1	0.2	-0.1	-0.1	0.50	0.50	0.50	0.50	0.50	0.499	0.499	0.499	0.499	0.499
0	0.1	0.2	-0.1	-0.1	0.50	0.50	0.50	0.50	0.50	0.501	0.501	0.501	0.501	0.501
0	0.1	0.2	-0.1	-0.1	0.50	0.50	0.50	0.50	0.50	0.501	0.501	0.501	0.501	0.501
1	1	1	1	1										
0	0.1	0.2	-0.1	0.2	0.50	0.44	0.40	0.57	0.38	0.496	0.156	0.032	0.896	0.015
0	0.1	0.2	-0.1	0.2	0.49	0.44	0.39	0.57	0.37	0.500	0.151	0.029	0.892	0.015
0	0.1	0.2	-0.1	0.2	0.50	0.50	0.50	0.50	0.50	0.494	0.494	0.494	0.494	0.494
1	1	1	1	1										
0	0.1	0.2	-0.1	-0.1	0.50	0.46	0.43	0.56	0.58	0.502	0.176	0.055	0.915	0.953
0	0.1	0.2	-0.1	-0.1	0.50	0.46	0.43	0.56	0.57	0.508	0.186	0.056	0.920	0.956
1	1	1	1	1										
0	0.1	0.2	-0.1	0.3	0.50	0.48	0.47	0.55	0.43	0.496	0.193	0.084	0.939	0.003
1	1	1	1	1										



CDF Matrices to Correlation Matrices – Both Targeted Scenario CDF Matrices n=6 n=126

1 1 1 1 1 1 1 1 1 1 0.00129 0.00129 0.00138 0.00138 0.00138 0.00138 0.5 0.4 0.6 0.35 0.9 0.00129 0.00129 0.00129 0.00138 -0.00062 -0.00062 0.00002 0.5 0.4 0.6 0.35 0.9 -0.00062 -0.00062 -0.00062 0.00002 0.00002 0.00002 0.00002 0.00156 0.00156 0.00156 0.5 0.35 0.9 -0.00438 -0.00438 -0.00438 0.00156 0.4 0.6 -0.00438 -0.00438 0.00156 0.5 0.4 0.6 0.35 0.9 -0.00175 -0.00175 -0.00175 -0.00175 -0.00175 0.00023 0.00023 0.00023 0.00023 0.00023 1 1 1 1 1 1 1 1 1 1 0.00032 0.16523 -0.15568 -0.49259 0.02137 -0.02326 -0.02326 -0.07492 0.5 0.4 0.6 0.6 0.8 -0.15568 -0.001130.5 -0.02142 0.4 0.6 0.6 0.8 -0.010120.14603 -0.15493 -0.15493-0.49518-0.000040.02195 -0.02142 -0.07445 -0.00083 0.5 0.4 0.6 0.6 0.8 -0.00599 -0.00599 -0.00599 -0.00599 -0.00599 -0.00083 -0.00083 -0.00083 -0.00083 1 1 1 1 1 1 1 1 1 1 0.55471 -0.02178 0.02401 0.5 0.4 0.6 0.4 0.3 0.00004 0.22686 -0.16786 0.22690 0.00026 0.02398 0.05183 0.5 0.4 0.6 0.4 0.3 -0.00484 0.19668 -0.20375 0.19891 0.36323 0.00173 0.02512 -0.02207 0.02515 0.04911 1 1 1 1 1 1 1 1 1 1 -0.02471 0.5 0.4 0.6 0.55 0.95 0.01793 0.34785 -0.21883 -0.08378 -0.56057 -0.001200.02173 -0.01286 -0.14962 1 1 1 1 1 1 1 1 1 1



CDF %Shift Matrices to Correlation Matrices – Both Targeted Scenario CDF %Shift Matrices n=6 n=126

					InMatPr	-2.222	-2.241	-6.893	-6.970	-12.755	-2.233	-2.230	-6.938	-6.925	-12.881
					FNorm	0.717	0.584	1.674	1.156	2.241	0.073	0.072	0.195	0.187	0.378
					Rnk_InMat	1	2	3	4	5	2	1	4	3	5
					Rnk_FNorr	2	1	4	3	5	2	1	4	3	5
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	60		0.0013	0.0013	0.0013	0.0013	0.0013	0.00138	0.00138	0.00138	0.00138	0.00138
20	-20	50	-50	60		-0.0006	-0.0006	-0.0006	-0.0006	-0.0006	0.00002	0.00002	0.00002	0.00002	0.00002
20	-20	50	-50	60		-0.0044	-0.0044	-0.0044	-0.0044	-0.0044	0.00156	0.00156	0.00156	0.00156	0.00156
20	-20	50	-50	60		-0.0017	-0.0017	-0.0017	-0.0017	-0.0017	0.00023	0.00023	0.00023	0.00023	0.00023
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	70		0.1698	-0.1509	0.4105	-0.4035	0.5916	0.02242	-0.02225	0.06101	-0.05946	0.09345
20	-20	50	-50	70		0.1517	-0.1495	0.4049	-0.4066	0.5885	0.02188	-0.02150	0.05874	-0.05964	0.09178
20	-20	50	-50	70		-0.0060	-0.0060	-0.0060	-0.0060	-0.0060	-0.00083	-0.00083	-0.00083	-0.00083	-0.00083
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	80		0.2325	-0.1663	0.5953	-0.2490	0.8728	0.02348	-0.02251	0.06302	-0.05724	0.12099
20	-20	50	-50	80		0.2028	-0.1972	0.4656	-0.4548	0.6567	0.02469	-0.02254	0.06181	-0.06089	0.11648
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	90		0.3405	-0.2326	0.7159	-0.2490	0.8010	0.02112	-0.02547	0.06286	-0.06024	0.15955
						1	1	1	1	1	1	1	1	1	1









Difference Between Mean Empirical Matrices – Both Targeted Scenario n = 126n = 6



Data v Angles-targ (epan-SILVR09-15) - Diff btwn Emp Mean Matrices: DGM=MVTVNS, Mat=5x5, #Obs=126, #Sims = 10000 Mean Dif 80000 60000 Density 40000 20000 0 00002 -0.00002 0.00000 0.00004 Distribution of Difference between Empirical Mean Correlation Matrices PRNG Seed = 12345 RunID = FTb5Mblk





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XII. APPENDIX 2: Empirical Results of NAbC, Targeted – Case F.TS







RunID = FTb5Mblk



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Angles-targ (epan-SILVR09-15)--Log(Pr(M)|Mean Matrix) v FNorm [r=-0.897]: DGM=MVTVNS, Mat=5x5, #Obs=126, #Sims = 10000



RunID = FTb5Mblk





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Correlation Matrices to CDF Matrices – Both Targeted Scenario Correlation Matrices n=6 n=126

111															
00.10.2-0.1-0.10.530.530.530.530.530.48740.48740.48740.48740.48740.48740.48740.487400.10.2-0.1-0.10.470.470.470.470.470.470.50240.5	1	1	1	1	1										
00.10.2-0.1-0.10.470.470.470.470.470.50240.50360.50360.50360.5040.50240.50360.50360.50360.50360.50360.50360.5040.50240.50360.5040.50240.50360.50360.50360.5040.50360.5070.50770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.56770.567	0	0.1	0.2	-0.1	-0.1	0.53	0.53	0.53	0.53	0.53	0.4874	0.4874	0.4874	0.4874	0.4874
00.10.2-0.1-0.10.470.470.470.470.470.53260.5570	0	0.1	0.2	-0.1	-0.1	0.47	0.47	0.47	0.47	0.47	0.5024	0.5024	0.5024	0.5024	0.5024
1000.10.10.10.40.460.460.460.460.52630.5670.567	0	0.1	0.2	-0.1	-0.1	0.47	0.47	0.47	0.47	0.47	0.5326	0.5326	0.5326	0.5326	0.5326
111	0	0.1	0.2	-0.1	-0.1	0.46	0.46	0.46	0.46	0.46	0.5263	0.5263	0.5263	0.5263	0.5263
00.10.2-0.10.20.40.440.440.430.370.2660.09380.0420.01840.006800.10.20.10.20.50.550.550.480.460.5730.4050.4050.2670.567 <td>1</td> <td>1</td> <td>1</td> <td>1</td> <td>1</td> <td></td>	1	1	1	1	1										
00.10.2-0.10.20.00.550.550.4040.40450.26710.1690.092200.10.20.00.00.00.00.00.00.5570.567<	0	0.1	0.2	-0.1	0.2	0.46	0.44	0.41	0.39	0.37	0.206	0.0938	0.042	0.0184	0.0068
00.10.2-0.10.20.50.550.550.550.550.5670.56770.5	0	0.1	0.2	-0.1	0.2	0.55	0.52	0.50	0.48	0.46	0.5734	0.4045	0.2671	0.169	0.0922
111	0	0.1	0.2	-0.1	0.2	0.55	0.55	0.55	0.55	0.55	0.5677	0.5677	0.5677	0.5677	0.5677
0 0.1 0.2 -0.1 -0.1 0.67 0.65 0.64 0.63 0.72 0.9824 0.953 0.9189 0.8742 0.9996 0 0.1 0.2 -0.1 -0.1 0.63 0.61 0.60 0.59 0.69 0.953 0.8954 0.8262 0.7569 0.9996 1	1	1	1	1	1										
0 0.1 0.2 -0.1 -0.1 0.63 0.61 0.60 0.59 0.69 0.9955 0.8954 0.8262 0.7569 0.9966 1 1 1 1 1 1 1 1 1 1 0.68 0.67 0.66 0.67 0.68 0.9958 0.8954 0.8262 0.7569 0.9966 0 0.1 1 1 1 0.68 0.67 0.66 0.760 0.9978 0.9998 0.9902 0.996 0.9995 1 1 1 1 1 1 0.66 0.67 0.66 0.70 0.9978 0.9993 0.9902 0.995	0	0.1	0.2	-0.1	-0.1	0.67	0.65	0.64	0.63	0.72	0.9824	0.953	0.9189	0.8742	0.9996
1 1	0	0.1	0.2	-0.1	-0.1	0.63	0.61	0.60	0.59	0.69	0.9535	0.8954	0.8262	0.7569	0.9996
0 0.1 0.2 -0.1 0.3 0.68 0.67 0.66 0.70 0.9978 0.9939 0.9902 0.986 0.9995 1	1	1	1	1	1										
1 1 1 1 1	0	0.1	0.2	-0.1	0.3	0.68	0.67	0.67	0.66	0.70	0.9978	0.9939	0.9902	0.986	0.9995
	1	1	1	1	1										



CDF Matrices to Correlation Matrices – Both Targeted Scenario n=6 n=126

CDF Matrices

				1	1	1	1	1	1	1	1	1	1
0.4	0.6	0.35	0.9	0.10	0.10	0.10	0.10	0.10	0.07	0.07	0.07	0.07	0.07
0.4	0.6	0.35	0.9	-0.11	-0.11	-0.11	-0.11	-0.11	-0.09	-0.09	-0.09	-0.09	-0.09
0.4	0.6	0.35	0.9	-0.11	-0.11	-0.11	-0.11	-0.11	-0.07	-0.07	-0.07	-0.07	-0.07
0.4	0.6	0.35	0.9	-0.15	-0.15	-0.15	-0.15	-0.15	-0.10	-0.10	-0.10	-0.10	-0.10
				1	1	1	1	1	1	1	1	1	1
0.4	0.6	0.6	0.8	-0.03	0.25	-0.30	-0.30	-0.74	-0.05	-0.01	-0.09	-0.09	-0.18
0.4	0.6	0.6	0.8	0.22	0.47	-0.06	-0.06	-0.63	0.12	0.16	0.07	0.07	-0.04
0.4	0.6	0.6	0.8	0.01	0.01	0.01	0.01	0.01	0.04	0.04	0.04	0.04	0.04
				1	1	1	1	1	1	1	1	1	1
0.4	0.6	0.4	0.3	0.55	0.75	0.32	0.72	0.91	0.37	0.40	0.34	0.40	0.44
0.4	0.6	0.4	0.3	0.49	0.67	0.19	0.65	0.54	0.32	0.35	0.28	0.35	0.37
				1	1	1	1	1	1	1	1	1	1
0.4	0.6	0.55	0.95	0.79	0.83	0.49	0.81	0.30	0.55	0.59	0.51	0.55	0.39
				1	1	1	1	1	1	1	1	1	1
	0.4 0.4 0.4 0.4 0.4 0.4 0.4 0.4 0.4 0.4	0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6 0.4 0.6	0.4 0.6 0.35 0.4 0.6 0.35 0.4 0.6 0.35 0.4 0.6 0.35 0.4 0.6 0.35 0.4 0.6 0.35 0.4 0.6 0.35 0.4 0.6 0.6 0.4 0.6 0.6 0.4 0.6 0.6 0.4 0.6 0.6 0.4 0.6 0.4 0.4 0.6 0.4 0.4 0.6 0.4 0.4 0.6 0.4 0.4 0.6 0.4 0.4 0.6 0.55	Image: Normal system Image: No	Image: Marking	Image: Marking Series Image: Marking Series	Image: Marking Series Image: Marking Series	1000000000000000000000000000000000000	1000000000000000000000000000000000000	$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$	Image: series of the series	Image: series of the series	Image: series of the series



CDF %Shift Matrices to Correlation Matrices – Both Targeted Scenario CDF %Shift Matrices n=6 n=126

-6.731	-7.204	-2.175	-2.292		-14.780	-6.449	-7.529	-2.101	-2.369	InMatPr					
0.325	0.318	0.117	0.122		1.884	1.913	1.620	0.927	0.889	FNorm					
3	4	1	2		5	3	4	1	2	Rnk_InMat					
4	3	1	2		4	5	3	2	1	Rnk_FNorr					
1	1	1	1		1	1	1	1	1						
0.07	0.07	0.07	0.07		0.10	0.10	0.10	0.10	0.10		60	-50	50	-20	20
-0.09	-0.09	-0.09	-0.09		-0.11	-0.11	-0.11	-0.11	-0.11		60	-50	50	-20	20
-0.07	-0.07	-0.07	-0.07		-0.11	-0.11	-0.11	-0.11	-0.11		60	-50	50	-20	20
-0.10	-0.10	-0.10	-0.10		-0.15	-0.15	-0.15	-0.15	-0.15		60	-50	50	-20	20
1	1	1	1		1	1	1	1	1						
-0.16	0.06	-0.09	-0.01		0.78	-0.66	0.61	-0.31	0.25		70	-50	50	-20	20
-0.03	0.21	0.04	0.14		0.84	-0.58	0.71	-0.19	0.36		70	-50	50	-20	20
0.04	0.04	0.04	0.04		0.01	0.01	0.01	0.01	0.01		70	-50	50	-20	20
1	1	1	1		1	1	1	1	1						
0.29	0.46	0.34	0.41		0.98	0.14	0.92	0.14	0.65		80	-50	50	-20	20
0.23	0.41	0.29	0.35		0.60	-0.32	0.67	0.02	0.57		80	-50	50	-20	20
1	1	1	1		1	1	1	1	1						
0.46	0.65	0.53	0.60		0.53	-0.21	0.70	0.16	0.81		90	-50	50	-20	20
1	1	1	1		1	1	1	1	1						
	-6.731 0.325 3 4 1 0.07 -0.09 -0.07 -0.10 1 1 -0.16 -0.03 0.04 1 0.29 0.23 1 0.46 1	-7.204 -6.731 0.318 0.325 4 3 3 4 1 1 0.07 0.07 -0.09 -0.09 -0.07 -0.07 -0.10 -0.10 1 1 0.06 -0.16 0.21 -0.03 0.04 0.04 1 1 0.46 0.29 0.41 0.23 1 1 0.65 0.46 1 1	-2.175 -7.204 -6.731 0.117 0.318 0.325 1 4 3 1 3 4 1 3 4 1 3 4 1 1 1 0.07 0.07 0.07 -0.09 -0.09 -0.09 -0.07 -0.07 -0.07 -0.10 -0.10 -0.10 1 1 1 -0.09 0.06 -0.16 0.01 -0.10 -0.03 0.02 0.04 0.04 0.04 0.21 -0.03 0.04 0.24 0.29 0.29 0.41 0.23 1 1 1 0.53 0.65 0.46 1 1 1	-2.292 -2.175 -7.204 -6.731 0.122 0.117 0.318 0.325 2 1 4 3 2 1 4 3 2 1 3 4 1 1 1 1 1 1 1 1 0.07 0.07 0.07 0.07 -0.09 -0.09 -0.09 -0.09 -0.07 -0.07 -0.07 -0.07 -0.01 -0.10 -0.10 -0.10 1 1 1 1 -0.01 -0.09 0.06 -0.16 0.14 0.04 0.21 -0.03 0.04 0.04 0.04 0.04 0.14 0.34 0.46 0.29 0.35 0.29 0.41 0.23 1 1 1 1 0.60 0.53 0.65 0.46 1 1 1 1	$\begin{array}{ c c c c c c c c } -2.292 & -2.175 & -7.204 & -6.731 \\ \hline 0.122 & 0.117 & 0.318 & 0.325 \\ \hline 0.122 & 0.117 & 0.318 & 0.325 \\ \hline 2 & 1 & 4 & 3 \\ \hline 2 & 1 & 4 & 3 \\ \hline 2 & 1 & 3 & 44 \\ \hline 2 & 1 & 3 & 44 \\ \hline 2 & 1 & 1 & 1 & 1 \\ \hline 1 & 1 & 1 & 1 \\ \hline 1 & 1 & 1 & 1 \\ \hline 1 & 0.07 & 0.07 & 0.07 \\ \hline 0.07 & 0.07 & 0.07 & 0.07 \\ \hline 0.09 & -0.09 & -0.09 \\ \hline 0.00 & -0.09 & -0.09 \\ \hline 0.00 & -0.07 & -0.07 & -0.07 \\ \hline 0.01 & -0.01 & -0.10 & -0.10 \\ \hline 0.01 & -0.01 & -0.10 & -0.10 \\ \hline 1 & 1 & 1 & 1 \\ \hline 1 & 1 & 1 & 1 \\ \hline 0.01 & 0.04 & 0.04 & 0.04 \\ \hline 0.04 & 0.04 & 0.04 & 0.04 \\ \hline 0.04 & 0.04 & 0.04 & 0.04 \\ \hline 1 & 1 & 1 & 1 \\ \hline 0.01 & 0.35 & 0.29 & 0.41 & 0.23 \\ \hline 0.04 & 0.60 & 0.53 & 0.65 & 0.46 \\ \hline 1 & 1 & 1 & 1 & 1 \\ \hline \end{array}$	-14.780 -2.292 -2.175 -7.204 -6.731 1.884 0.122 0.117 0.318 0.325 5 2 1 4 3 4 2 1 3 4 1 1 1 3 4 1 1 1 3 4 1 1 1 1 1 0.10 0.07 0.07 0.07 0.07 0.11 0.07 0.07 0.07 0.07 -0.11 -0.09 -0.09 -0.09 -0.09 -0.15 -0.10 -0.10 -0.10 -0.10 1 1 1 1 1 0.78 -0.01 -0.09 0.06 -0.16 0.84 0.14 0.04 0.04 0.04 1 1 1 1 1 0.98 0.41 0.34 0.46 0.29 0.60 0.35 0.29 0.41 0.23 1 1 1 1 1 0.53 0.60 0.53 0.65 0.46	-6.449-14.780-2.292-2.175-7.204-6.7311.9131.8840.1220.1170.3180.325352143542134111134000.00.070.070.070.100.100.070.070.070.070.11-0.11-0.09-0.09-0.09-0.09-0.11-0.15-0.10-0.10-0.10-0.100.05-0.15-0.10-0.07-0.07-0.07-0.15-0.15-0.01-0.090.06-0.16-0.580.840.140.040.040.041111110.140.980.410.340.460.29-0.320.600.350.290.411-0.210.530.600.530.650.46111111	-7.529 -6.449 -14.780 -2.292 -2.175 -7.204 -6.731 1.620 1.913 1.884 0.122 0.117 0.318 0.325 4 3 5 2 1 4 3 3 5 4 2 1 3 4 1 1 1 2 1 3 4 1 1 1 1 1 1 1 1 1 1 1 1 1 1 0.10 0.10 0.00 0.07 0.07 0.07 0.11 0.10 0.10 0.07 0.07 0.07 0.11 -0.11 -0.11 -0.09 -0.09 -0.09 -0.11 -0.11 -0.15 -0.10 -0.07 -0.07 -0.15 -0.15 -0.15 -0.10 -0.10 -0.10 1 1 1 1 1 1 0.61 -0.66 0.78 -0.01 -0.09 0.06 0.71 -0.58 0.84 0.14 0.04 0.04 0.01 0.01 0.04 0.04 0.04 0.21 0.02 0.14 0.98 0.41 0.34 0.46 0.72 0.53 0.60 0.53 0.65 0.46 1 1 1 1 1 1 1	-2.101 -7.529 -6.449 -14.780 -2.292 -2.175 -7.204 -6.731 0.927 1.620 1.913 1.884 0.122 0.117 0.318 0.325 1 4 3 5 2 1 4 3 2 3 5 4 2 1 4 3 1 1 1 1 1 1 1 1 1 1 1 1 1 1 1 0.10 0.10 0.10 0.07 0.07 0.07 0.11 0.10 0.10 0.00 0.07 0.07 0.11 -0.11 -0.11 -0.09 -0.09 -0.09 -0.11 -0.11 -0.11 -0.07 -0.07 -0.07 -0.15 -0.15 -0.15 -0.15 -0.10 -0.10 -0.10 -0.11 -0.11 -0.11 -0.11 -0.07 -0.07 -0.07 -0.15 -0.15 -0.15 -0.15 -0.10 -0.10 -0.10 -0.15 -0.15 -0.15 -0.01 -0.09 0.06 -0.16 -0.19 0.71 -0.58 0.84 0.14 0.04 0.04 0.04 1 1 1 1 1 1 1 1 1 0.01 0.01 0.01 0.01 0.04 0.04 0.04 0.04 0.14 0.92 0.14 0.98 0.41	-2.369 -2.101 -7.529 -6.449 -14.780 -2.292 -2.175 -7.204 -6.731 0.889 0.927 1.620 1.913 1.884 0.122 0.117 0.318 0.325 2 1 4 3 5 2 1 4 3 1 2 3 5 4 2 1 4 3 1 2 3 5 4 2 1 4 3 1 0.10 0.10 0.10 0.10 0.07 0.07 0.07 0.07 0.11 -0.11 -0.11 -0.11 -0.11 -0.09 -0.09 -0.09 -0.09 -0.11 -0.11 -0.11 -0.11 -0.17 -0.07 -0.07 -0.07 -0.15 -0.15 -0.15 -0.15 -0.10 -0.10 -0.10 -0.10 1 1 1 1 1 1 1 1 1 1 0.25 -0.31 0.61 -0.66 0.78 -0.01 -0.09 0.06 -0.16 0.36 -0.19 0.71 -0.58 0.84 0.14 0.04 0.04 0.04 0.15 0.16 0.14 0.98 0.41 0.34 0.46 0.29 </td <td>InMatPr -2.369 -2.101 -7.529 -6.449 -14.780 -2.292 -2.175 -7.204 -6.731 FNorm 0.889 0.927 1.620 1.913 1.884 0.122 0.117 0.318 0.325 Rnk_INMI 2 1 4 3 5 2 1 4 3 Rnk_FNort 1 2 3 5 4 2 1 4 3 Rnk_FNort 1 1 1 1 1 1 3 4 1 1 1 1 1 1 1 3 4 1</td> <td>$\begin{array}{ c c c c c c c c c c c c c c c c c c c$</td> <td>$\begin{array}{ c c c c c c c c c c c c c c c c c c c$</td> <td>$\begin{array}{ c c c c c c c c c c c c c c c c c c c$</td> <td>$\begin{array}{ c c c c c c c c c c c c c c c c c c c$</td>	InMatPr -2.369 -2.101 -7.529 -6.449 -14.780 -2.292 -2.175 -7.204 -6.731 FNorm 0.889 0.927 1.620 1.913 1.884 0.122 0.117 0.318 0.325 Rnk_INMI 2 1 4 3 5 2 1 4 3 Rnk_FNort 1 2 3 5 4 2 1 4 3 Rnk_FNort 1 1 1 1 1 1 3 4 1 1 1 1 1 1 1 3 4 1	$\begin{array}{ c c c c c c c c c c c c c c c c c c c$	$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$	$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$	$ \begin{array}{ c c c c c c c c c c c c c c c c c c c$


- Empirical Results from NAbC Example Cases (APPENDIX 3):
- **F. KT and F.SP**: MVTVNS (*ρ* = block structure, df=3,4,5,6,7, skewness=1,0.6,0,-0.6,-1, autocorrelation= -0.25,0,0.25,0.5,0.75, nonstationarity=3σ, σ/3, σ, n/3 each) (n=12, n=126)

Cases F.KT and F.SP: *R* (*Pearson's*) =

1	-0.1	-0.1	0.2	0.2
-0.1	1	-0.1	0.2	0.2
-0.1	-0.1	1	0.2	0.2
0.2	0.2	0.2	1	0.5
0.2	0.2	0.2	0.5	1

- Empirical Results from NAbC Summary (see APPENDIX 3 for Graphs/Tables):
- <u>Cases F.KT and F.SP</u>: These results are based on a 'real world' example (Case F.) where marginal distributions vary by tail heaviness, skewness, autocorrelation, and degree of nonstationarity. As with Pearson's matrix in Case F., we again see identical results distributionally for 3. (direct data simulations) v 4. (kernel-based angle perturbation via NAbC) across all comparison criteria: eigenvalue distributions, correlation cell distributions, angle distributions, capital (VaR) distributions, and norm distributions (only Frobenius/Euclidean is shown, but Taxicab & Chebyshev also are equal).
- Note the increased precision in the angles distributions of Kendall's over Spearman's (and Spearman's over Pearson's (from Case F.)).
- Also note that numerical considerations become relevant for small samples (eg matrix dimension p = n+1) when calculating Kendall's and Spearman's matrices: to avoid a non-trivial percentage of numerically non-positive definite results (out of N=10k simulations), when p=5, n must equal 12 rather than 6 (or p+1), as in Case F. for Pearson's. So for consistency, Case F.PR is presented for n = 12.





n = 126



n = 12





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PRNG Seed = 12345 RunID = Fs5 KT

 $RunID = Fb5_KT$



LNP



n = 12



n = 126

RunID = Fb5 KT



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n = 12



n = 126





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n = 12





RunID = Fb5_KT



n = 12



n = 126





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n = 12



n = 126





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n = 12



n = 126





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n = 12



n = 126





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Angle Distributions







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n = 12





n = 12



n = 126





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n = 12



n = 126





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Angle Distributions







n = 12







n = 12



n = 126





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Correlation Matrices to CDF Matrices

	Mean τ, N=10k			N=10k Correlation Matrices					new τ CDFs						new τ CDFs					
ρ	n=12	n=126	(new	Pearso	on's ρ i	nto Co	pula)				n=12				r	n=126				
1	1	1	1	1	1	1	1													
-0.1	0.09	0.10	0	0.1	0.2	0.2	0.2		0.65	0.53	0.36	0.36	0.36	0.8349	0.4482	0.1462	0.1462	0.1462		
-0.1	-0.07	-0.05	0	0.1	0.2	-0.1	0.2		0.43	0.30	0.17	0.55	0.17	0.3404	0.0623	0.0037	0.7232	0.0037		
0.2	-0.05	-0.01	0	0.1	0.2	0.1	-0.1		0.46	0.36	0.22	0.36	0.57	0.5636	0.2969	0.0769	0.2969	0.7499		
0.2	-0.06	-0.01	0	0.1	0.2	0.1	0.1		0.45	0.34	0.21	0.34	0.34	0.5595	0.2770	0.0608	0.2770	0.2770		
1	1	1	1	1	1	1	1													
-0.1	-0.03	0.00	0	0.1	0.2	-0.1	0.2		0.48	0.35	0.26	0.60	0.26	0.5323	0.1528	0.0270	0.8594	0.0270		
0.2	0.07	0.11	0	0.1	0.2	0.1	-0.1		0.62	0.49	0.38	0.51	0.73	0.8816	0.6462	0.3549	0.6769	0.9659		
0.2	0.05	0.10	0	0.1	0.2	0.1	0.2		0.60	0.47	0.37	0.48	0.34	0.8577	0.5881	0.2875	0.6224	0.2226		
1	1	1	1	1	1	1	1													
0.2	0.26	0.28	0	0.1	0.2	0.1	0.3		0.82	0.73	0.66	0.69	0.37	0.9993	0.9893	0.9428	0.9701	0.1654		
0.2	0.22	0.25	0	0.1	0.2	0.2	0		0.77	0.67	0.60	0.50	0.82	0.9975	0.9660	0.8641	0.5818	0.9996		
1	1	1	1	1	1	1	1													
0.5	0.38	0.39	0	0.1	0.2	0.2	0.3		0.85	0.79	0.74	0.70	0.44	0.9980	0.9891	0.9669	0.9298	0.2876		
1	1	1	1	1	1	1	1													



CDF Matrices

CDF Matrices to $\, \mathcal{T} \,$ Correlation Matrices

n=12

					1	1	1	1	1	1	1	1	1	1	
0.5	0.4	0.6	0.35	0.9	0.12	0.17	0.03	0.21	-0.24	0.09	0.11	0.06	0.12	-0.03	
0.5	0.4	0.6	0.35	0.9	-0.06	0.02	-0.12	0.05	-0.43	-0.04	-0.01	-0.07	0.00	-0.18	
0.5	0.4	0.6	0.35	0.9	-0.03	0.05	-0.13	0.10	-0.46	0.03	0.06	-0.02	0.08	-0.27	
0.5	0.4	0.6	0.35	0.9	-0.02	0.03	-0.12	0.10	-0.46	0.02	0.06	-0.02	0.07	-0.25	
					1	1	1	1	1	1	1	1	1	1	
0.5	0.4	0.6	0.6	0.8	-0.02	0.06	-0.09	-0.07	-0.12	0.00	0.03	-0.02	-0.01	-0.06	
0.5	0.4	0.6	0.6	0.8	0.08	0.16	0.01	0.03	-0.02	0.13	0.16	0.10	0.11	0.04	
0.5	0.4	0.6	0.6	0.8	0.07	0.15	-0.01	0.02	-0.04	0.12	0.14	0.09	0.10	0.03	
					1	1	1	1	1	1	1	1	1	1	
0.5	0.4	0.6	0.4	0.3	0.26	0.33	0.20	0.33	0.53	0.27	0.29	0.25	0.29	0.34	
0.5	0.4	0.6	0.4	0.3	0.22	0.30	0.15	0.29	0.51	0.24	0.26	0.21	0.26	0.31	
					1	1	1	1	1	1	1	1	1	1	
0.5	0.4	0.6	0.55	0.95	0.37	0.47	0.29	0.37	0.21	0.38	0.41	0.34	0.37	0.27	
					1	1	1	1	1	1	1	1	1	1	



n=126

	CDF %Shift Matrices to \mathcal{L} Correlation Matrices													
	CDF	%Shift	: Matri	ces		I	n=12				n	126		
				InMatPr	-2.288	-2.176	-7.162	-6.714	-13.056	-2.320	-2.151	-7.342	-6.656	-12.973
				FNorm	0.363	0.326	0.980	0.758	1.493	0.142	0.128	0.362	0.354	0.542
				Rnk_InMatPr	2	1	4	3	5	2	1	4	3	5
				Rnk_FNorm	2	1	4	3	5	2	1	4	3	5
					1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	60	0.17	0.03	0.30	-0.10	0.33	0.12	0.07	0.16	0.03	0.18
20	-20	50	-50	60	-0.01	-0.12	0.13	-0.24	0.18	-0.02	-0.07	0.02	-0.11	0.04
20	-20	50	-50	60	0.04	-0.15	0.16	-0.28	0.22	0.04	-0.05	0.10	-0.13	0.12
20	-20	50	-50	60	0.03	-0.16	0.15	-0.28	0.22	0.03	-0.05	0.09	-0.12	0.12
					1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	70	0.05	-0.09	0.20	-0.18	0.31	0.02	-0.02	0.06	-0.06	0.09
20	-20	50	-50	70	0.15	0.00	0.30	-0.07	0.41	0.14	0.08	0.20	0.04	0.24
20	-20	50	-50	70	0.14	-0.02	0.28	-0.10	0.40	0.13	0.07	0.18	0.03	0.22
					1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	80	0.33	0.20	0.47	0.15	0.63	0.30	0.26	0.34	0.23	0.39
20	-20	50	-50	80	0.30	0.16	0.44	0.11	0.61	0.27	0.22	0.31	0.19	0.36
					1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	90	0.47	0.30	0.64	0.20	0.84	0.42	0.35	0.48	0.30	0.57
					1	1	1	1	1	1	1	1	1	1





n = 126



n = 12



 $RunID = Fb5_SP$



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 $RunID = Fb5_SP$













n = 12

n = 126





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n = 126







n = 12

n = 126





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n = 12







n = 12

n = 126





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n = 12

n = 126





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n = 12







n = 12

n = 126





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n = 12







n = 12







n = 12







n = 12

n = 126





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Correlation Matrices to CDF Matrices

	Mean p	o-sp, N=10k	c Co	orrelat	ion M	latrice	S		new	ρ-sp C	DFs		new ρ-sp CDFs					
ρ	n=12	n=126	(new P	earso	n'sρi	nto Co	pula)			n=12	=12			I	า=126			
								L/										
1	1	1	1	1	1	1	1											
-0.1	0.12	0.14	0	0.1	0.2	0.2	0.2	0.63	0.53	0.43	0.43	0.43	0.8356	0.5656	0.2836	0.2836	0.2836	
-0.1	-0.09	-0.07	0	0.1	0.2	-0.1	0.2	0.41	0.31	0.22	0.51	0.22	0.3327	0.1091	0.0183	0.6181	0.0183	
0.2	-0.07	-0.03	0	0.1	0.2	0.1	-0.1	0.44	0.36	0.27	0.36	0.53	0.5446	0.3542	0.1663	0.3542	0.6929	
0.2	-0.08	-0.02	0	0.1	0.2	0.1	0.1	0.43	0.34	0.26	0.34	0.34	0.5422	0.3382	0.1515	0.3382	0.3382	
1	1	1	1	1	1	1	1											
-0.1	-0.03	-0.01	0	0.1	0.2	-0.1	0.2	0.48	0.39	0.32	0.56	0.32	0.5407	0.2374	0.0846	0.7969	0.0846	
0.2	0.09	0.14	0	0.1	0.2	0.1	-0.1	0.62	0.54	0.45	0.54	0.70	0.9050	0.7377	0.5262	0.7545	0.9694	
0.2	0.07	0.12	0	0.1	0.2	0.1	0.2	0.60	0.51	0.44	0.52	0.42	0.8733	0.6839	0.4640	0.7054	0.4069	
1	1	1	1	1	1	1	1											
0.2	0.33	0.38	0	0.1	0.2	0.1	0.3	0.80	0.74	0.70	0.71	0.51	0.9993	0.9940	0.9750	0.9849	0.5818	
0.2	0.29	0.33	0	0.1	0.2	0.2	0	0.75	0.69	0.64	0.58	0.79	0.9960	0.9755	0.9324	0.8061	0.9991	
1	1	1	1	1	1	1	1											
0.5	0.47	0.52	0	0.1	0.2	0.2	0.3	0.82	0.78	0.75	0.73	0.57	0.9957	0.9850	0.9700	0.9471	0.6041	
1	1	1	1	1	1	1	1											



CDF Matrices

CDF Matrices to P_{sp} Correlation Matrices n=12

					1	1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.13	0.23	0.03	0.28	-0.35	0.1	2 0.16	0.09	0.17	-0.04
0.5	0.4	0.6	0.35	0.9	-0.09	0.01	-0.19	0.05	-0.56	-0.0	6 -0.02	-0.09	-0.01	-0.26
0.5	0.4	0.6	0.35	0.9	-0.07	0.05	-0.18	0.11	-0.63	0.0	3 0.08	-0.03	0.10	-0.39
0.5	0.4	0.6	0.35	0.9	-0.07	0.03	-0.18	0.09	-0.64	0.0	2 0.07	-0.03	0.09	-0.37
					1	1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.6	0.8	-0.03	0.08	-0.12	-0.10	-0.07	0.0	0 0.04	-0.02	-0.02	-0.07
0.5	0.4	0.6	0.6	0.8	0.11	0.23	0.02	0.05	0.08	0.1	8 0.22	0.14	0.16	0.07
0.5	0.4	0.6	0.6	0.8	0.09	0.20	-0.01	0.02	0.05	0.1	6 0.20	0.12	0.14	0.05
					1	1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.36	0.46	0.28	0.45	0.72	0.3	6 0.39	0.34	0.39	0.47
0.5	0.4	0.6	0.4	0.3	0.31	0.42	0.22	0.41	0.71	0.3	2 0.35	0.29	0.35	0.44
					1	1	1	1	1		1 1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.53	0.66	0.40	0.52	0.44	0.5	0 0.55	0.45	0.50	0.40
					1	1	1	1	1		1 1	1	1	1



n=126

					CDF %S	Shift Ma	atrice	sto /		prrelatio	n Matrices				
	CDF	%Shif	t Matri	ces			r		n=126						
					InMatPr	-2.327	-2.137	-7.322	-6.563	-14.243	-2.350	-2.122	-7.470	-6.545	-13.366
					FNorm	0.507	0.441	1.362	0.997	1.917	0.200	0.176	0.507	0.484	0.742
					Rnk_InMatPr	2	1	4	3	5	2	1	4	3	5
					Rnk_FNorm	2	1	4	3	5	2	1	4	3	5
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	60)	0.23	0.03	0.39	-0.13	0.45	0.17	0.10	0.23	0.05	0.26
20	-20	50	-50	60)	0.01	-0.19	0.17	-0.35	0.23	-0.03	-0.10	0.03	-0.16	0.05
20	-20	50	-50	60)	0.03	-0.19	0.21	-0.37	0.28	0.04	-0.08	0.13	-0.20	0.16
20	-20	50	-50	60)	0.02	-0.19	0.20	-0.38	0.28	0.04	-0.08	0.12	-0.19	0.15
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	70)	0.08	-0.13	0.29	-0.22	0.43	0.03	-0.03	0.09	-0.08	0.13
20	-20	50	-50	70)	0.20	-0.01	0.41	-0.10	0.55	0.19	0.11	0.27	0.05	0.32
20	-20	50	-50	70)	0.18	-0.03	0.39	-0.13	0.54	0.17	0.09	0.24	0.04	0.30
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	80)	0.44	0.26	0.63	0.21	0.80	0.40	0.35	0.45	0.32	0.52
20	-20	50	-50	80)	0.40	0.21	0.60	0.17	0.79	0.36	0.30	0.42	0.27	0.49
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	90)	0.62	0.35	0.82	0.23	0.96	0.56	0.47	0.64	0.40	0.75
						1	1	1	1	1	1	1	1	1	1





n = 12

n = 126





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LNP









n = 12

n = 126





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n = 12









n = 12







n = 12







n = 12







n = 12







n = 12

n = 126





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n = 12







n = 12

n = 126





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n = 12







n = 12

n = 126





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Angle Distributions





n = 126







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Correlation Matrices to CDF Matrices

	Mean p	o, N=10k	10k Correlation Matrices							nev	νρCD	Fs	new ρ CDFs					
ρ	n=12	n=126	(new	Pearso	on's ρ into Copula)						n=12				n	=126		
									ľ									
1	1	1		1 1	1	1	1											
-0.1	0.141	0.151		0 0.1	0.2	0.2	0.2		0.63	0.55	0.48	0.48	0.48	0.81	0.60	0.36	0.36	0.36
-0.1	-0.099	-0.072		0 0.1	0.2	-0.1	0.2		0.41	0.34	0.27	0.48	0.27	0.34	0.15	0.05	0.58	0.05
0.2	-0.070	-0.001		0 0.1	0.2	0.1	-0.1		0.44	0.37	0.31	0.37	0.52	0.56	0.39	0.22	0.39	0.70
0.2	-0.079	-0.004		0 0.1	0.2	0.1	0.1		0.43	0.36	0.30	0.36	0.36	0.55	0.38	0.21	0.38	0.38
1	1	1		1 1	1	1	1											
-0.1	-0.035	-0.006		0 0.1	0.2	-0.1	0.2		0.48	0.42	0.37	0.54	0.37	0.54	0.30	0.15	0.73	0.15
0.2	0.105	0.172		0 0.1	0.2	0.1	-0.1		0.61	0.55	0.49	0.56	0.66	0.89	0.76	0.60	0.78	0.96
0.2	0.080	0.149		0 0.1	0.2	0.1	0.2		0.58	0.52	0.47	0.53	0.46	0.86	0.70	0.53	0.72	0.49
1	1	1		1 1	1	1	1											
0.2	0.368	0.422		0 0.1	0.2	0.1	0.3		0.78	0.74	0.70	0.72	0.58	1.00	0.99	0.97	0.98	0.71
0.2	0.320	0.372		0 0.1	0.2	0.2	0		0.73	0.69	0.66	0.61	0.76	0.99	0.96	0.92	0.83	1.00
1	1	1		1 1	1	1	1											
0.5	0.508	0.560		0 0.1	0.2	0.2	0.3		0.79	0.76	0.74	0.72	0.62	0.99	0.97	0.95	0.93	0.67
1	1	1		1 1	1	1	1											



CDF Matrices to $ ho$	Correlation Matrices	
n=	12	n=126

CDF Matrices

					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.35	0.9	0.17	0.31	0.04	0.37	-0.48	0.14	0.18	0.10	0.20	-0.07
0.5	0.4	0.6	0.35	0.9	-0.12	0.01	-0.26	0.08	-0.67	-0.07	-0.02	-0.11	0.00	-0.29
0.5	0.4	0.6	0.35	0.9	-0.08	0.06	-0.22	0.14	-0.69	0.04	0.10	-0.02	0.12	-0.37
0.5	0.4	0.6	0.35	0.9	-0.09	0.04	-0.23	0.12	-0.70	0.03	0.09	-0.03	0.12	-0.35
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.6	0.8	-0.04	0.12	-0.17	-0.12	0.01	0.00	0.05	-0.03	-0.02	-0.09
0.5	0.4	0.6	0.6	0.8	0.14	0.30	0.01	0.07	0.14	0.21	0.26	0.16	0.19	0.09
0.5	0.4	0.6	0.6	0.8	0.10	0.26	-0.04	0.02	0.12	0.18	0.23	0.13	0.16	0.06
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.4	0.3	0.45	0.58	0.35	0.57	0.85	0.42	0.45	0.38	0.45	0.54
0.5	0.4	0.6	0.4	0.3	0.39	0.53	0.29	0.53	0.84	0.37	0.41	0.33	0.40	0.50
					1	1	1	1	1	1	1	1	1	1
0.5	0.4	0.6	0.55	0.95	0.65	0.79	0.49	0.65	0.60	0.56	0.62	0.50	0.56	0.41
					1	1	1	1	1	1	1	1	1	1



	CDF %Shift Matrices to $ ho$ Correlation Matrices														
	CDF	%Shif	t Matri	ices			r	າ=12໌		n=126					
				InN	MatPr	-2.354	-2.113	-7.448	-6.482	-14.624	-2.363	-2.105	-7.491	-6.452	-13.825
				FN	lorm	0.709	0.586	1.788	1.225	2.337	0.232	0.206	0.598	0.552	0.868
				Rn	k_InMatPr	2	1	4	3	5	2	1	4	3	5
				Rn	k_FNorm	2	1	4	3	5	2	1	4	3	5
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	60		0.28	0.01	0 49	-0 21	0.55	0.19	0 11	0.26	0.04	0.29
20	-20	50	-50	60		0.20	-0.24	0.45	-0.44	0.33	-0.03	-0.11	0.04	-0.18	0.23
20	-20	50	-50	60		0.03	-0.21	0.29	-0.43	0.38	0.03	-0.06	0.16	-0.18	0.20
20	-20	50	-50	60		0.05	-0.22	0.29	-0.45	0.38	0.06	-0.06	0.16	-0.18	0.19
-						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	70		0.12	-0.16	0.41	-0.25	0.59	0.04	-0.04	0.11	-0.10	0.18
20	-20	50	-50	70		0.27	-0.02	0.54	-0.13	0.70	0.23	0.13	0.32	0.07	0.39
20	-20	50	-50	70		0.25	-0.05	0.53	-0.16	0.68	0.21	0.11	0.30	0.04	0.36
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	80		0.53	0.27	0.76	0.26	0.91	0.46	0.39	0.52	0.34	0.60
20	-20	50	-50	80		0.49	0.23	0.74	0.22	0.90	0.41	0.33	0.48	0.29	0.57
						1	1	1	1	1	1	1	1	1	1
20	-20	50	-50	90		0.72	0.34	0.91	0.25	0.99	0.62	0.50	0.70	0.42	0.82
						1	1	1	1	1	1	1	1	1	1

